
Correlation Robust Influence Maximization: Supplementary Material

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Broader Impact

The aim of this work is to address the possible pitfalls to the independence assumption in a social network, as used in the study of influence maximization. As discussed previously, how an idea, product, or piece of news makes its way through a network could very well be impacted by natural social biases, thus connecting parts of a social network in ways that could have been unforeseen. The methodology presented thus attempts to make this possibility a consideration during the selection of seed set, and hence find “influential” members to a network regardless of whatever underlying correlations may exist. This potentially can reduce the impact of biases that the independence assumption may cause.

The Correlation Robust Influence Function f^{corr}

Theorem 1 *Let $G = (V, E)$ be a directed graph, $\mathcal{S} \subseteq V$ a seed set, and $\mathbf{p} \in [0, 1]^E$ a vector of edge likelihoods. Then $\min_{\theta \in \Theta} \mathbf{E}_{\tilde{\mathbf{c}} \sim \theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})]$ is the value to the following polynomial sized linear program.*

$$\begin{aligned} \min_{\theta \in \Theta} \mathbf{E}_{\tilde{\mathbf{c}} \sim \theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})] &= \min_{\boldsymbol{\pi} \in \mathbb{R}^V} \sum_{i \in V \setminus \mathcal{S}} \pi_i \\ \text{s.t.} \quad &\pi_i = 1 \text{ for } i \in \mathcal{S}, \\ &\pi_i - \pi_j \leq 1 - p_{ij} \text{ for } (i, j) \in E, \\ &0 \leq \pi_i \leq 1 \text{ for } i \in V \end{aligned} \tag{3}$$

Proof: According to [1], if we let M assume a large value (anything at least $|V \setminus \mathcal{S}|$), then $\min_{\theta \in \Theta(p)} \mathbf{E}_{\tilde{\mathbf{c}} \sim \theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})]$ can be formulated as the following linear program:

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$$\begin{aligned}
& \min_{\pi \in \mathbb{R}^{V \cup \{s,t\}}, \lambda} \sum_{i,j \in E} 0 \cdot \lambda_{ij}^0 + M \cdot \lambda_{ij}^1 + \sum_{i \in \mathcal{S}} 0 \cdot \lambda_{si}^0 + M \cdot \lambda_{si}^1 + \sum_{i \in N \setminus \mathcal{S}} 0 \cdot \lambda_{it}^0 + \lambda_{it}^1 \\
& \text{subject to} \quad \pi_i - \pi_j \leq \lambda_{ij}^0 \quad \forall (i,j) \in E \\
& \quad \pi_i - \pi_j - (1 - p_{ij}) \leq \lambda_{ij}^1 \quad \forall (i,j) \in E \\
& \quad \pi_s - \pi_i \leq \lambda_{si}^0 \quad \forall i \in \mathcal{S} \\
& \quad \pi_s - \pi_i \leq \lambda_{si}^1 \quad \forall i \in \mathcal{S} \\
& \quad \pi_i - \pi_t \leq \lambda_{it}^0 \quad \forall i \in V \setminus \mathcal{S} \\
& \quad \pi_i - \pi_t \leq \lambda_{it}^1 \quad \forall i \in V \setminus \mathcal{S} \\
& \quad 0 \leq \pi \leq 1; \quad \forall i \in V \\
& \quad \lambda \geq 0 \\
& \quad \pi_s = 1, \pi_t = 0
\end{aligned}$$

Upon inspection, the program reduces to the desired program. \square

Corollary 1 (Correlation Robust Influence Likelihood) *For an arbitrary seed set \mathcal{S} and vector of edge likelihoods $\mathbf{p} \in [0, 1]^E$, let π^* solve (3). Then for each $i \in V \setminus \mathcal{S}$,*

$$\pi_i^* = \left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+,$$

$$\mathbb{P}_{\theta^*}(\text{Node } i \text{ is reachable from } \mathcal{S} \text{ in } G(\tilde{\mathbf{c}})) = \pi_i^* \quad \forall \theta^* \in \arg \min_{\theta \in \Theta} \mathbb{E}_{\tilde{\mathbf{c}} \sim \theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})].$$

*In light of this, we define the **correlation robust influence likelihood of i** as π_i^* . In particular, π_i^* is no greater than the IC model's likelihood that i is influenced, that is, $\pi_i^* \leq \mathbb{P}_{\theta_{ic}}(\text{Node } i \text{ is reachable from } \mathcal{S})$.*

Proof: We begin by establishing the equality

$$\pi_i^* = \left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+.$$

Let i be such that $\Gamma(\mathcal{S}, i) \neq \emptyset$. Consider any path $\gamma \in \Gamma(\mathcal{S}, i)$ and let $\gamma = (i_0 \rightarrow i_1 \rightarrow i_2 \rightarrow \dots \rightarrow i_l = i)$, where $i_0 \in \mathcal{S}$. Since π^* is feasible to (3), we must have,

$$\begin{aligned}
\pi_{i_0}^* - \pi_{i_1}^* &\leq 1 - p_{i_0, i_1} \\
\pi_{i_1}^* - \pi_{i_2}^* &\leq 1 - p_{i_1, i_2} \\
&\vdots \\
\pi_{i_{l-1}}^* - \pi_i^* &\leq 1 - p_{i_{l-1}, i_l}
\end{aligned}$$

Summation of these inequalities gives $\pi_{i_0}^* - \pi_i^* \leq \sum_{l=1}^l (1 - p_{i_{l-1}, i_l})$. Since $i_0 \in \mathcal{S}$, it follows that $\pi_{i_0}^* = 1$, so that $\pi_i^* \geq L(\gamma)$. Hence, $\pi_i^* = \left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+$.

On the other hand, observe that if $\Gamma(\mathcal{S}, i) = \emptyset$, then the decision variable π_i has no lower bound other than 0. Further, $\left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+ = 0$, in such a case, as desired.

We next establish the remaining equality

$$\mathbb{P}_{\theta^*}(\text{Node } i \text{ is reachable from } \mathcal{S}) = \pi_i^* \quad \forall \theta^* \in \arg \min_{\theta \in \Theta} \mathbb{E}_{\tilde{\mathbf{c}} \sim \theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})].$$

Taking note of

$$\begin{aligned}
\min_{\theta \in \Theta} \mathbb{E}_{\theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})] &= \min_{\theta \in \Theta} \sum_{i \in V \setminus \mathcal{S}} \mathbb{P}_{\theta}(\text{Node } i \text{ is reachable from } \mathcal{S} \text{ in } G(\tilde{\mathbf{c}})) \\
&\geq \sum_{i \in V \setminus \mathcal{S}} \min_{\theta \in \Theta} \mathbb{P}_{\theta}(\text{Node } i \text{ is reachable from } \mathcal{S} \text{ in } G(\tilde{\mathbf{c}})) \geq \sum_{i \in V \setminus \mathcal{S}} \left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+ \\
&= \sum_{i \in V \setminus \mathcal{S}} \pi_i^* = \min_{\theta \in \Theta} \mathbb{E}_{\theta} [Z(\tilde{\mathbf{c}}, \mathcal{S})],
\end{aligned}$$

and that for any $i \in V \setminus \mathcal{S}$, it holds that

$$\min_{\theta \in \Theta} \mathbb{P}_{\theta}(\text{Node } i \text{ is reachable from } \mathcal{S} \text{ in } G(\tilde{\mathbf{c}})) \geq \left[\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) \right]^+,$$

so we arrive at the desired conclusion. \square

Corollary 2 (Path existence under Correlation Robustness) *Let \mathcal{S} be an arbitrary seed set, and let $\theta^* \in \Theta$ be any solution to $\min_{\theta \in \Theta} \mathbb{E}_{\theta}[R(\tilde{\mathbf{c}}, \mathcal{S})]$. Let $\bar{\Gamma}(\mathcal{S}, i) := \arg \max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma)$, and π^* is any optimal solution to (3). If $i \notin \mathcal{S}$ and $\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) > 0$, then*

$$\mathbb{P}_{\theta^*}(\cup_{\gamma \in \bar{\Gamma}(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) = \pi_i^* = \mathbb{P}_{\theta^*}(\cap_{\gamma \in \bar{\Gamma}(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]),$$

In addition, if $\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) > 0$, then for any path $\gamma \in \bar{\Gamma}(\mathcal{S}, i)$, at most one of the arcs in γ is ever missing in the random graph $G(\tilde{\mathbf{c}}) \sim \theta^$, almost surely.*

Proof: If θ^* solves $\min_{\theta \in \Theta} \mathbb{E}_{\theta}[R(\tilde{\mathbf{c}}, \mathcal{S})]$, $i \notin \mathcal{S}$, and $\max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) > 0$, then for any $\gamma^* \in \bar{\Gamma}(\mathcal{S}, i)$,

$$\begin{aligned} \max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma) &= \mathbb{P}_{\theta^*}(\text{Node } i \text{ is reachable from } \mathcal{S} \text{ in } G(\tilde{\mathbf{c}})) = \mathbb{P}_{\theta^*}(\cup_{\gamma \in \Gamma(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) \\ &\geq \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*]) \stackrel{(5)}{\geq} L(\gamma^*) = \max_{\gamma \in \Gamma(\mathcal{S}, i)} L(\gamma). \end{aligned}$$

So we conclude that

$$\mathbb{P}_{\theta^*}(\cup_{\gamma \in \bar{\Gamma}(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) = \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*]),$$

which implies

$$\begin{aligned} \mathbb{P}_{\theta^*}(\cup_{\gamma \in \bar{\Gamma}(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) &\geq \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*]) \\ &= \mathbb{P}_{\theta^*}(\cup_{\gamma \in \Gamma(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]), \end{aligned}$$

as desired.

For the remaining equality in the statement, we note that if

$$\mathbb{P}_{\theta^*}(\cap_{\gamma \in \bar{\Gamma}(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) < \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*]),$$

then $\mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*] \setminus [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma']) > 0$ for some $\gamma' \in \bar{\Gamma}(\mathcal{S}, i)$, which means

$$\begin{aligned} \mathbb{P}_{\theta^*}(\cup_{\gamma \in \Gamma(\mathcal{S}, i)} [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma]) &\geq \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma']) \\ &\quad + \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*] \setminus [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma']) \\ &> \mathbb{P}_{\theta^*}([G(\tilde{\mathbf{c}}) \text{ contains path } \gamma^*] \setminus [G(\tilde{\mathbf{c}}) \text{ contains path } \gamma']), \end{aligned}$$

a contradiction.

As for the last statement, if $\gamma \in \bar{\Gamma}(\mathcal{S}, i)$, we observe that under the joint distribution θ^* it cannot be the case that - with positive probability - more than one arc is missing from $G(\tilde{\mathbf{c}})$, else (5) would be a strict inequality, contradicting the fact that Corollary 1 implies that it should be an equality. \square

Corollary 3 *Given an arbitrary seed set \mathcal{S} and vector of edge likelihoods $\mathbf{p} \in [0, 1]^E$, let π^* denote the optimal solution to (3). Let $\tilde{q} \sim \text{Unif}[0, 1]$, $V(\tilde{q}) := \{i : \tilde{q} < \pi_i^*\}$,*

$$E(\tilde{q}) := \{(k, j) : \pi_k^* > \pi_j^*, \tilde{q} \notin [\pi_k^* - 1 + p_{kj}, \pi_k^*]\} \cup \{(k, j) : \pi_k^* \leq \pi_j^*, \tilde{q} \in (0, p_{kj}]\},$$

and $\mathbf{c}(\tilde{q}) \in \{0, 1\}^E$ be such that $c(\tilde{q})_{ij} = 1$ iff $(i, j) \in E(\tilde{q})$. Then $\mathbf{c}(\tilde{q}) \sim \theta^$ for some θ^* solving Equation (1). In particular, $V(\tilde{q})$ is the set of all nodes reachable from \mathcal{S} in the graph $G(\tilde{q}) = (V, E(\tilde{q}))$, so that $\mathbb{E}_{\tilde{q}}[|V(\tilde{q})|] = \min_{\theta \in \Theta} \mathbb{E}_{\tilde{\mathbf{c}} \sim \theta}[R(\tilde{\mathbf{c}}, \mathcal{S})] = |\mathcal{S}| + \mathbb{E}_{\tilde{q}}[Z(\mathbf{c}(\tilde{q}), \mathcal{S})]$.*

Proof: Consider the max-flow problem of $Z(c, \mathcal{S})$ for arbitrary $c \in \{0, 1\}^E$. Then the two collections $\{s\} \cup \mathcal{S} \cup \{j : x_{jt}^* = 1, j \in V \setminus \mathcal{S}\}$ and $\{t\} \cup \{j : x_{jt}^* = 0, j \in V \setminus \mathcal{S}\}$ form a minimum s - t cut. In particular, $\{j : x_{jt}^* = 1, j \in V \setminus \mathcal{S}\}$ is precisely the set of nodes outside of \mathcal{S} that are reached, and j is reached if and only if the edge (j, t) runs across this minimum cut.

With π^* an optimal solution to (3), we may characterize a $\theta^* \in \Theta$ consistent with \mathbf{p} that solves $\min_{\theta \in \Theta} \mathbf{E}[Z(\tilde{c}, \mathcal{S})] = \min_{\theta \in \Theta} \mathbf{E}[R(\tilde{c}, \mathcal{S})] - |\mathcal{S}|$. This characterization will be defined on the probability space $((0, 1], \mathcal{B}, \lambda)$, and for the sake of notation, in the following we'll let F_{ij} denote the cdf for edge (i, j) that is live with probability \mathbf{p}_{ij} . For all $(i, j) \in E$, if $\pi_i^* > \pi_j^*$, define for all $q \in (0, 1]$,

$$\tilde{c}_{ij}(q) := \begin{cases} F_{ij}^{-1}(q - \pi_j^*); & \pi_j^* < q \leq \pi_i^* \\ F_{ij}^{-1}(1 - p_{ij} + q); & 0 < q \leq \pi_i^* - (1 - p_{ij}) \\ F_{ij}^{-1}(1 - p_{ij} - \pi_j^* + q); & \pi_i^* - (1 - p_{ij}) < q \leq \pi_j^* \\ F_{ij}^{-1}(q); & \pi_i^* < q \leq 1, \end{cases}$$

otherwise if $\pi_i^* \leq \pi_j^*$ define $\tilde{c}_{ij}(q) := F_{ij}^{-1}(1 - q)$. Finally, for all $(i, j) \notin E$ but are auxillary arcs with s or t as an endpoint, we can let $\tilde{c}_{ij}(q) := +\infty$ if $i = s$, else $\tilde{c}_{ij}(q) := 1$ for the case that $j = t$. As well, we define

$$\tilde{\chi}_{ij}(q) := \begin{cases} 1; & \pi_i^* > \pi_j^*, q \in [\pi_j^*, \pi_i^*] \\ 0; & \text{otherwise.} \end{cases}$$

The resulting random vector \tilde{c} has as its distribution a solution to $\min_{\theta \in \Theta} \mathbf{E}[Z(\tilde{c}, \mathcal{S})]$. This follows after adopting the arguments in Theorem 3.1 of [1]. It is not hard to see that with $\tilde{q} \sim \text{Unif}(0, 1]$, $E(\tilde{q})$ as defined in the statement is precisely $\{(k, j) : \tilde{c}_{kj}(\tilde{q}) = 1\}$. Furthermore, according to Theorem 3.1 of [1], $\tilde{\chi}_{jt}(\tilde{q})$ is 1 if and only if (j, t) runs across the minimum cut - equivalently, when j is reached. And since $\pi_i^* = 0$ always, we arrive at the characterization of $V(\tilde{q})$. \square

Correlation Robustness: Maximization and Robust Ratios

Theorem 2 *The problem of computing $\max_{\mathcal{S}: |\mathcal{S}| \leq k} f^{\text{corr}}(\mathcal{S})$, given a graph $G = (V, E)$, a vector of edge likelihoods $\mathbf{p} \in [0, 1]^E$, and an integer number k , is NP-Hard. In particular, we have the following exact formulation as a mixed-integer program.*

$$\begin{aligned} \max_{\mathcal{S}: |\mathcal{S}| \leq k} f_p^{\text{corr}}(\mathcal{S}) &= \max \sum_{(i,j) \in E} z_{ij}(p_{ij} - 1) + \sum_{i \in V} w_i \\ &1 - y_i - \sum_{j: (j,i) \in E} z_{ji} + \sum_{j: (i,j) \in E} z_{ij} \geq 0 \quad \forall i \in V \\ &w_i \geq |V|x_i + y_i - |V| \quad \forall i \in V \\ &w_i \leq \min(|V|x_i, y_i) \quad \forall i \in V \\ &\sum_{i \in V} x_i = k \\ &y_i \geq 0 \quad w_i \geq 0 \quad \forall i \in V \\ &z_{ij} \geq 0 \quad \forall (i, j) \in E \\ &x_i \in \{0, 1\} \quad \forall i \in V \end{aligned}$$

Proof: We prove the hardness of computing $\max_{\mathcal{S}: |\mathcal{S}| \leq k} f^{\text{corr}}(\mathcal{S})$ through a reduction from the set cover problem. The proof is along the lines of the proof of hardness of the independent cascade model in [2]. In the set cover problem, there is a universe of elements $\Omega = \{1, \dots, n\}$, a collection of subsets $J_1, \dots, J_m \subseteq \Omega$ (whose union gives Ω), and an integer k . The decision version of the set cover problem is to check if there exists a collection of k subsets, whose union gives Ω . We will now reduce an instance of set cover problem to (2). For this, consider a bipartite graph with a total of $m + n$ vertices corresponding to the m subsets and the n elements of Ω . This bipartite graph contains an edge between a subset node i and an element node j if $j \in J_i$. Fix $p_{ij} = 1$ for all edges (i, j) in this graph. Then there exist k subsets whose union is Ω if and only if the optimal value to (2) is $k + n$.

Next we will derive the MILP formulation. Using Theorem 1, we have,

$$\begin{aligned}
\max_{\mathcal{S}:|\mathcal{S}|\leq k} f_p^{corr}(\mathcal{S}) &= \max_{x \in \mathbb{R}^V} \min_{\pi \in \mathbb{R}^V} \sum_i \pi_i \\
\text{subject to} & \quad x_i \leq \pi_i \quad \forall i \in V \\
& \quad \pi_i - \pi_j \leq 1 - p_{ij} \quad \forall (i, j) \in E \\
& \quad 0 \leq \pi_i \leq 1 \quad \forall i \in V \\
& \quad \sum_{i \in V} x_i = k \\
& \quad x_i \in \{0, 1\} \quad \forall i \in V
\end{aligned}$$

The dual of the inner minimization problem is,

$$\max_{z \geq 0, y \geq 0, w \geq 0} \sum_{(i,j) \in E} z_{ij}(p_{ij} - 1) + \sum_{i \in V} x_i y_i : 1 - y_i - \sum_{j:(j,i) \in E} z_{ji} + \sum_{j:(i,j) \in E} z_{ij} \geq 0 \quad \forall i \in V$$

Further we linearize the product terms $w_i = x_i y_i$. Summing up the inequality over all i gives us, $\sum_{i \in V} (1 - y_i - \sum_{j:(j,i) \in E} z_{ji} + \sum_{j:(i,j) \in E} z_{ij}) \geq 0$. The terms involving z cancel out and we are left with $\sum_{i \in V} y_i \leq V$ and since $y_i \geq 0$ for all i , we get an upper bound $y_i \leq V$.

Using the bounds $0 \leq x_i \leq 1$ and $0 \leq y_i \leq |V|$, the McCormick inequalities introduced in [3] for w_i give us,

$$|V|x_i + y_i - |V| \leq w_i \leq \min(|V|x_i, y_i) \quad \forall i \in V$$

To see that these inequalities are sufficient to capture $w_i = x_i y_i$, when $x_i \in \{0, 1\}$, first let $x_i = 0$. Then the inequalities give us $y_i - |V| \leq w_i \leq \min(0, y_i)$ and along with the fact that $w_i \geq 0$, we get $w_i = 0$. Now Let $x_i = 1$. Then the inequalities give us $y_i \leq w_i \leq \min(|V|, y_i) = y_i$. Therefore we get $w_i = y_i$ and hence these inequalities are tight. \square

Theorem 3 *The correlation robust influence function $f^{corr} : 2^V \rightarrow \mathbb{R}_+$ is a monotone, submodular function.*

Proof: Since $f^{corr}(\mathcal{S}) = |\mathcal{S}| + \min_{\theta \in \Theta} \mathbb{E}_\theta[Z(\tilde{c}, \mathcal{S})]$, submodularity of $g(\mathcal{S}) := \min_{\theta \in \Theta} \mathbb{E}_{\tilde{c} \sim \theta} Z[\tilde{c}, \mathcal{S}]$ implies submodularity of f^{corr} . If two seed sets S and T with $S \subset T$ and vertex $v \notin T$ are given, then by (7),

$$\begin{aligned}
g(S+v) - g(S) &= \sum_{i \notin (S \cup v)} \max \left(\left[\max_{\gamma \in \Gamma(S, i)} L(\gamma) \right]^+, \left[\max_{\gamma \in \Gamma(\{v\}, i)} L(\gamma) \right]^+ \right) \\
&\quad - \left[\sum_{i \notin (S+v)} \left[\max_{\gamma \in \Gamma(S, i)} L(\gamma) \right]^+ + \left[\max_{\gamma \in \Gamma(S, v)} L(\gamma) \right]^+ \right] \\
&= \sum_{i \notin (S+v)} \left[\left[\max_{\gamma \in \Gamma(\{v\}, i)} L(\gamma) \right]^+ - \left[\max_{\gamma \in \Gamma(S, i)} L(\gamma) \right]^+ \right]^+ - \left[\max_{\gamma \in \Gamma(S, v)} L(\gamma) \right]^+ \quad (1) \\
&\geq \sum_{i \notin (T+v)} \left[\left[\max_{\gamma \in \Gamma(\{v\}, i)} L(\gamma) \right]^+ - \left[\max_{\gamma \in \Gamma(T, i)} L(\gamma) \right]^+ \right]^+ - \left[\max_{\gamma \in \Gamma(T, v)} L(\gamma) \right]^+ \\
&= g(T+v) - g(T),
\end{aligned}$$

as desired. As for monotonicity, simply observe that by (1),

$$f^{corr}(S+v) - f^{corr}(S) = g(S+v) - g(S) + 1 \geq 1 - \left[\max_{\gamma \in \Gamma(S, v)} L(\gamma) \right]^+ \geq 0.$$

\square

Corollary 4 *Let \mathcal{S}_{corr}^g denote the seed set generated upon termination of the greedy algorithm for maximization of f^{corr} . Then*

$$f^{corr}(\mathcal{S}_{corr}^g) \geq (1 - 1/e) \max_{|\mathcal{S}| \leq k} f^{corr}(\mathcal{S})$$

Proof: By Theorem 3 and known approximation guarantees for submodular optimization [4] we get the result. \square

Computations for Example 2, POC study

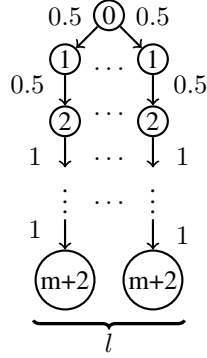


Figure 1: Example 2 for POC study

We consider the tree in Figure 1 with a root node, containing l children. There are a total of l paths from the root to all the leaf nodes, starting from the root node. Each path contains $m + 2$ nodes (apart from the root). The labels on the nodes indicate the “type” of each node. Between nodes of type 0 and 1 as well as between type 1 and type 2 nodes, the activation probability = 0.5. For all other edges, activation probability is 1. The total number of nodes in the graph is $n = l(m + 2) + 1$. Suppose we are interested in choosing a single seed node, so $k = 1$.

Independent cascade model: We first compute the values of $f^{ic}(\cdot)$ for each type of node.

Type 2: For such nodes, $f^{ic}(\{2\}) = m + 1$. Also it can be verified that nodes of type 2 reach more than nodes of type 3, 4, \dots $m + 2$.

Type 1: There is one random edge which, if active, will enable $m + 1$ nodes to be reached. However if this edge is inactive, none of the nodes are reached. Therefore, $f^{ic}(\{1\}) = \frac{m+1}{2} + 1$.

Type 0 (root): Here we are l sub-trees (each corresponding to a path graph) in which the nodes could be potentially reached. Let the number of nodes reached in each of the sub-trees be denoted by the random variables $\tilde{X}_1, \dots, \tilde{X}_l$. The object of our interest is $\mathbb{E}_{\theta_{ic}} \left[\sum_{i=1}^l \tilde{X}_i \right] + 1$. \tilde{X}_i takes values $m + 2, 1$ and 0 with probabilities $0.25, 0.25$ and 0.5 respectively. and therefore $\mathbb{E}[\tilde{X}_i] = (m + 3)/4$. Therefore the overall reachability $f^{ic}(\{0\}) = 1 + l(m + 3)/4$.

Clearly the choice to be made is between the root node and any node of type 2 (as node 2 is always better than node 1 (assuming $m \geq 1$). The root node is preferred when $l(m + 3)/4 \geq m$ which occurs when $l \geq \frac{4m}{m+3}$.

Worst case analysis: We perform a similar analysis on the values of $f^{corr}(\cdot)$ too. For any type 2 node, we have $f^{corr}(\{2\}) = m + 1$. When $\mathcal{S} = \{1\}$, $f^{corr}(\{1\}) = 1 + \frac{m+1}{2}$ as an optimal solution to the LP that computes $f^{corr}(\{1\})$ is $\pi_2^* = \pi_3^* = \dots = \pi_{m+2}^* = 0.5$ from Corollary 1.

Type 0 (root): In each sub-tree of the root node, our LP solution gives $\pi_1^* = 0.5, \pi_2^* = \pi_3^* = \dots = \pi_{m+2}^* = 0$. Therefore $f^{corr}(\{0\}) = 1 + l/2$.

Between type 0 and type 2 nodes, type 0 is selected whenever $l > 2m$ and a type 2 node can be selected otherwise.

Suppose $\frac{4m}{m+3} \leq l \leq 2m$. Then if $k = 1$, \mathcal{S}_{corr} is any one of the type 2 nodes while $\mathcal{S}_{ic} = \{0\}$. Then the price of correlations is $\frac{(l/2)+1}{m+1}$. If $l = \frac{4m}{m+3}$, then $POC = \frac{2m+3}{(m+1)(m+3)}$ which tends to zero as $m \rightarrow \infty$.

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