# Supplementary Materials of "A Universal Approximation Theorem of Deep Neural Networks for Expressing Probability Distributions" 

Yulong Lu
Department of Mathematics and Statistics
University of Massachusetts Amherst
Amherst, MA 01003
lu@math.umass.edu

## Jianfeng Lu

Mathematics Department
Duke University
Durham, NC 27708
jianfeng@math.duke.edu

## A Proof of Proposition 3.1

Proof. The proof follows from some previous results by [1] and [4]. In fact, in the one dimensional case, according to [1, Theorem 3.2], we know that if $\pi$ satisfies that

$$
\begin{equation*}
J_{1}(\pi)=\int_{-\infty}^{\infty} \sqrt{F(x)(1-F(x))} d x<\infty \tag{A.1}
\end{equation*}
$$

where $F$ is the cumulative distribution function of $\pi$, then for every $n \geq 1$,

$$
\begin{equation*}
\mathbf{E} \mathcal{W}_{1}\left(P_{n}, \pi\right) \leq \frac{J_{1}(\pi)}{\sqrt{n}} \tag{A.2}
\end{equation*}
$$

The condition A.1 is fulfilled if $\pi$ has finite third moment since

$$
J_{1}(\pi) \leq \int_{0}^{\infty} \sqrt{\mathbf{P}(|X| \geq x)} d x \leq 1+\int_{1}^{\infty} \frac{\sqrt{\mathbf{E}|X|^{3}}}{x^{\frac{3}{2}}} d x=1+2 \sqrt{M_{3}}
$$

In the case that $d \geq 2$, it follows from that [4, Theorem 3.1] if $M_{3}=\mathbb{E}_{X \sim \pi}|X|^{3} d \pi<\infty$, then there exists a constant $c>0$ independent of $d$ such that

$$
\mathbf{E} \mathcal{W}_{1}\left(P_{n}, \pi\right) \leq c M_{3}^{1 / 3} \cdot \begin{cases}\frac{\log n}{\sqrt{n}} & \text { if } d=2  \tag{A.3}\\ \frac{1}{n^{1 / d}} & \text { if } d \geq 3\end{cases}
$$

## B Proof of Proposition 3.2

Proof. Thanks to [7, Proposition 3.1], one has that

$$
\operatorname{MMD}\left(P_{n}, \pi\right)=\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}}
$$

Let us define $\varphi\left(X_{1}, X_{2}, \cdots, X_{n}\right):=\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}}$. Then by definition $\varphi\left(X_{1}, X_{2}, \cdots, X_{n}\right)$ satisfies that for any $i \in\{1, \cdots, n\}$,

$$
\begin{aligned}
& \left|\varphi\left(X_{1}, \cdots, X_{i-1}, X_{i}, \cdots, X_{n}\right)-\varphi\left(X_{1}, \cdots, X_{i-1}, X_{i}^{\prime}, \cdots, X_{n}\right)\right| \\
& \leq \frac{2}{N} \sup _{x}\|k(\cdot, x)\|_{\mathcal{H}_{k}} \\
& \leq \frac{2 \sqrt{K_{0}}}{N}, \forall X_{i}, X_{i}^{\prime} \in \mathbb{R}^{d}
\end{aligned}
$$

where we have used that $\|k(\cdot, x)\|_{\mathcal{H}_{k}}=\sup _{x} \sqrt{k(x, x)} \leq \sqrt{K_{0}}$ by assumption. It follows from above and the McDiarmid's inequality that for every $\tau>0$, with probability $1-e^{-\tau}$,

$$
\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}} \leq \mathbf{E}\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}}+\sqrt{\frac{2 \sqrt{K_{0}} \tau}{n}} .
$$

In addition, we have by the standard symmetrization argument that

$$
\mathbf{E}\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}} \leq 2 \mathbf{E E}_{\varepsilon}\left\|\frac{1}{n} \sum_{i=1}^{n} \varepsilon_{i} k\left(\cdot, X_{i}\right)\right\|_{\mathcal{H}_{k}},
$$

where $\left\{\varepsilon_{i}\right\}_{i=1}^{n}$ are i.i.d. Radmacher variables and $\mathbf{E}_{\varepsilon}$ represents the conditional expectation w.r.t $\left\{\varepsilon_{i}\right\}_{i=1}^{n}$ given $\left\{X_{i}\right\}_{i=1}^{n}$. To bound the right hand side above, we can apply McDiarmid's inequality again to obtain that with probability at least $1-e^{-\tau}$,

$$
\begin{aligned}
\mathbf{E E}_{\varepsilon}\left\|\frac{1}{n} \sum_{i=1}^{n} \varepsilon_{i} k\left(\cdot, X_{i}\right)\right\|_{\mathcal{H}_{k}} & \leq \mathbf{E}_{\varepsilon}\left\|\frac{1}{n} \sum_{i=1}^{n} \varepsilon_{i} k\left(\cdot, X_{i}\right)\right\|_{\mathcal{H}_{k}}+\sqrt{\frac{2 \sqrt{K_{0}} \tau}{n}} \\
& \leq\left(\mathbf{E}_{\varepsilon}\left\|\frac{1}{n} \sum_{i=1}^{n} \varepsilon_{i} k\left(\cdot, X_{i}\right)\right\|_{\mathcal{H}_{k}}^{2}\right)^{1 / 2}+\sqrt{\frac{2 \sqrt{K_{0}} \tau}{n}} \\
& \leq \sqrt{\frac{\sqrt{K_{0}}}{n}}+\sqrt{\frac{2 \sqrt{K_{0}} \tau}{n}}
\end{aligned}
$$

where we have used Jensen's inequality for expectation in the second inequality and the independence of $\varepsilon_{i}$ and the definition of $K_{0}$ in the last inequality. Combining the estimates above yields that with probability at least $1-2 e^{-\tau}$,

$$
\operatorname{MMD}\left(P_{n}, \pi\right)=\left\|\int_{\mathbb{R}^{d}} k(\cdot, x) d\left(P_{n}-\pi\right)(x)\right\|_{\mathcal{H}_{k}} \leq 2 \sqrt{\frac{\sqrt{K_{0}}}{n}}+3 \sqrt{\frac{2 \sqrt{K_{0}} \tau}{n}}
$$

## C Proof of Proposition 3.3

Thanks to [6, Theorem 3.6], $\operatorname{KSD}\left(P_{n}, \pi\right)$ is evaluated explicitly as

$$
\begin{equation*}
\operatorname{KSD}\left(P_{n}, \pi\right)=\sqrt{\mathbf{E}_{x, y \sim P_{n}}\left[u_{\pi}(x, y)\right]}=\sqrt{\frac{1}{n^{2}} \sum_{i, j=1}^{n} u_{\pi}\left(X_{i}, X_{j}\right)} \tag{C.1}
\end{equation*}
$$

where $u_{\pi}$ is a new kernel defined by

$$
\begin{aligned}
& u_{\pi}(x, y)=s_{\pi}(x)^{T} k(x, y) s_{\pi}(y)+s_{\pi}(x)^{T} \nabla_{y} k(x, y) \\
& +s_{\pi}(y)^{T} \nabla_{x} k(x, y)+\operatorname{Tr}\left(\nabla_{x} \nabla_{y} k(x, y)\right)
\end{aligned}
$$

with $s_{\pi}(x)=\nabla \log \pi(x)$. Moreover, according to [6, Proposition 3.3], if $k$ satisfies Assumption K1, then $\operatorname{KSD}\left(P_{n}, \pi\right)$ is non-negative.
Our proof of Proposition 3.3 relies on the fact that $\operatorname{KSD}^{2}\left(P_{n}, \pi\right)$ can be viewed as a von Mises' statistics ( $V$-statistics) and an important Bernstein type inequality due to [2] for the distribution of $V$-statistics, which gives a concentration bound of $\operatorname{KSD}^{2}\left(P_{n}, \pi\right)$ around its mean (which is zero). We recall this inequality in the theorem below, which is a restatement of [2, Theorem 1] for second order degenerate $V$-statistics.

## C. 1 Bernstein type inequality for von Mises' statistics

Let $X_{1}, \cdots, X_{n}, \cdots$ be a sequence of i.i.d. random variables on $\mathbb{R}^{d}$. For a kernel $h(x, y): \mathbb{R}^{d} \times$ $\mathbb{R}^{d} \rightarrow \mathbb{R}$, we call

$$
\begin{equation*}
V_{n}=\sum_{i, j=1}^{n} h\left(X_{i}, X_{j}\right) \tag{C.2}
\end{equation*}
$$

a von-Mises' statistic of order 2 with kernel $h$. We say that the kernel $h$ is degenerate if the following holds:

$$
\begin{equation*}
\mathbf{E}\left[h\left(X_{1}, X_{2}\right) \mid X_{1}\right]=\mathbf{E}\left[h\left(X_{1}, X_{2}\right) \mid X_{2}\right]=0 \tag{C.3}
\end{equation*}
$$

Theorem C. 1 ([2, Theorem 1]). Consider the $V$-statistic $M_{n}$ defined by (C.2) with a degenerate kernel h. Assume the kernel satisfies that

$$
\begin{equation*}
|h(x, y)| \leq g(x) \cdot g(y) \tag{C.4}
\end{equation*}
$$

for all $x, y \in \mathbb{R}^{d}$ with a function $g: \mathbb{R}^{d} \rightarrow \mathbb{R}$ satisfying for $\xi, J>0$,

$$
\begin{equation*}
\mathbf{E}\left[g\left(X_{1}\right)^{k}\right] \leq \xi^{2} J^{k-2} k!/ 2 \tag{C.5}
\end{equation*}
$$

for all $k=2,3, \cdots$. Then there exist some generic constants $C_{1}, C_{2}>0$ independent of $k, h, l, \xi$ such that for any $t \geq 0$ that

$$
\begin{equation*}
\mathbf{P}\left(\left|V_{n}\right| \geq n^{2} t\right) \leq C_{1} \exp \left(-\frac{C_{2} n t}{\xi^{2}+J t^{1 / 2}}\right) \tag{C.6}
\end{equation*}
$$

Remark C.1. As noted in [2] Remark 1], the inequality (C.6) is to some extent optimal. Moreover, a straightforward calculation shows that inequality (C.6) implies that for any $\delta \in(0,1)$,

$$
\begin{equation*}
\mathbf{P}\left(\frac{1}{n^{2}}\left|V_{n}\right| \leq \frac{\mathscr{V}}{n}\right) \geq 1-\delta, \tag{C.7}
\end{equation*}
$$

where

$$
\mathscr{V}=\left(\frac{J \log \left(\frac{C_{1}}{\delta}\right)}{C_{2}}+\sqrt{\frac{\log \left(\frac{C_{1}}{\delta}\right)}{C_{2}}} \xi\right)^{2} .
$$

## C. 2 Moment bound of sub-Gaussian random vectors

Let us first recall a useful concentration result on sub-Gaussian random vectors.
Theorem C. 2 ([3], Theorem 2.1]). Let $X \in \mathbb{R}^{d}$ be a sub-Gaussian random vector with parameters $m \in \mathbb{R}^{d}$ and $v>0$. Then for any $t>0$,

$$
\begin{equation*}
\mathbf{P}(|X-m| \geq v \sqrt{d+2 \sqrt{d t}+2 t}) \leq e^{-t} \tag{C.8}
\end{equation*}
$$

Moreover, for any $0 \leq \eta<\frac{1}{2 v^{2}}$,

$$
\begin{equation*}
\mathbf{E} \exp \left(\eta|X-m|^{2}\right) \leq \exp \left(v^{2} d \eta+\frac{v^{4} d \eta^{2}}{1-2 v^{2} \eta}\right) \tag{C.9}
\end{equation*}
$$

As a direct consequence of Theorem C.2, we have the following useful moment bound for subGaussian random vectors.
Proposition C.1. Let $X \in \mathbb{R}^{d}$ be a sub-Gaussian random vector with parameters $m \in \mathbb{R}^{d}$ and $v>0$. Then for any $k \geq 2$,

$$
\begin{equation*}
\mathbf{E}|X-m|^{k} \leq k\left((2 v \sqrt{d})^{k}+\frac{1}{2}\left(\frac{4 v}{\sqrt{2}}\right)^{k} k^{k / 2}\right) \tag{C.10}
\end{equation*}
$$

Proof. From the concentration bound (C.8) and the simple fact that

$$
d+2 \sqrt{d t}+2 t=2\left(\sqrt{t}+\frac{\sqrt{d}}{2}\right)^{2}+\frac{d}{2} \leq 4\left(\sqrt{t}+\frac{\sqrt{d}}{2}\right)^{2}
$$

one can obtain that

$$
\mathbf{P}\left(|X-m| \geq 2 v\left(\sqrt{t}+\frac{\sqrt{d}}{2}\right)\right) \leq \mathbf{P}(|X-m| \geq v \sqrt{d+2 \sqrt{d t}+2 t}) \leq e^{-t}
$$

Therefore, for any $s \geq v \sqrt{d}$, we obtain from above with $s=2 v(\sqrt{t}+\sqrt{d} / 2)$ that

$$
\begin{equation*}
\mathbf{P}(|X-m| \geq s) \leq e^{-\left(\frac{s}{2 v}-\frac{\sqrt{d}}{2}\right)^{2}} \tag{C.11}
\end{equation*}
$$

As a result, for any $k \geq 2$,

$$
\begin{align*}
\mathbf{E}|X-m|^{k} & =\int_{0}^{\infty} \mathbf{P}\left(|X-m|^{k} \geq s\right) d s \\
& =\int_{0}^{\infty} \mathbf{P}\left(|X-m|^{k} \geq s^{k}\right) k s^{k-1} d s  \tag{C.12}\\
& =\int_{0}^{2 v \sqrt{d}} \mathbf{P}(|X-m| \geq s) k s^{k-1} d s+\int_{2 v \sqrt{d}}^{\infty} \mathbf{P}(|X-m| \geq s) k s^{k-1} d s \\
& =I_{1}+I_{2}
\end{align*}
$$

where we have used the change of variable $s \mapsto s^{k}$ in the second line above. It is clear that the first term

$$
I_{1} \leq k(2 v \sqrt{d})^{k}
$$

For $I_{2}$, one first notices that if $s \geq 2 v \sqrt{d}$, then $s /(2 v)-\sqrt{d} / 2 \geq s /(4 v)$. Hence it follows from (C.8) that

$$
\begin{aligned}
I_{2} & \leq \int_{2 v \sqrt{d}}^{\infty} e^{-\left(\frac{s}{4 v}\right)^{2}} k s^{k-1} d s \\
& =\frac{k(4 v)^{k}}{2} \int_{\frac{d}{4}}^{\infty} e^{-t} t^{\frac{k-2}{2}} d t \\
& \leq \frac{k(4 v)^{k}}{2} \Gamma\left(\frac{k}{2}\right)
\end{aligned}
$$

The last two estimates imply that

$$
\begin{aligned}
\mathbf{E}|X-m|^{k} & \leq k(2 v \sqrt{d})^{k-1}+\frac{k(4 v)^{k}}{2} \Gamma\left(\frac{k}{2}\right) \\
& \leq k\left((2 v \sqrt{d})^{k}+\frac{1}{2}\left(\frac{4 v}{\sqrt{2}}\right)^{k} k^{k / 2}\right)
\end{aligned}
$$

where the second inequality above follows from $\Gamma\left(\frac{k}{2}\right) \leq(k / 2)^{k / 2}$ for $k \geq 2$.

## C. 3 Proof of Proposition 3.3

Our goal is to invoke Theorem C. 1 to obtain a concentration inequality for KSD. Recall that $\operatorname{KSD}\left(P_{n}, \pi\right)$ is defined by

$$
\operatorname{KSD}^{2}\left(P_{n}, \pi\right)=\frac{1}{n^{2}} \sum_{i, j=1}^{n} u_{\pi}\left(X_{i}, X_{j}\right)
$$

with the kernel

$$
u_{\pi}(x, y)=s_{\pi}(x)^{T} k(x, y) s_{\pi}(y)+s_{q}(x)^{T} \nabla_{y} k(x, y)+s_{q}(y)^{T} \nabla_{x} k(x, y)+\operatorname{Tr}\left(\nabla_{x} \nabla_{y} k(x, y)\right)
$$

Let us first verify that the new kernel $u_{\pi}$ satisfies the assumption of Theorem C. 1 In fact, since $s_{\pi}(x)=\nabla \log (\pi(x))$, one obtains from integration by part that

$$
\begin{aligned}
& \mathbf{E}\left[u_{\pi}\left(X_{1}, X_{2}\right) \mid X_{1}=x\right]=\int_{\mathbb{R}^{d}} u_{\pi}(x, y) d \pi(y) \\
& =\int_{\mathbb{R}^{d}} s_{\pi}(x) k(x, y) s_{\pi}(y)+s_{q}(x)^{T} \nabla_{y} k(x, y) \\
& \quad+s_{\pi}(y)^{T} \nabla_{x} k(x, y)+\operatorname{Tr}\left(\nabla_{x} \nabla_{y} k(x, y)\right) d \pi(y) \\
& =\int_{\mathbb{R}^{d}} k(x, y) s_{\pi}(x)^{T} \nabla_{y} \pi(y) d y-\int_{\mathbb{R}^{d}} \nabla_{y} \cdot\left(s_{\pi}(x) \pi(y)\right) d y \\
& \quad+\int_{\mathbb{R}^{d}} \nabla_{y} \pi(y)^{T} \nabla_{x} k(x, y) d y-\int_{\mathbb{R}^{d}} \nabla_{y} \pi(y)^{T} \nabla_{x} k(x, y) d y \\
& =0
\end{aligned}
$$

Similarly, one has

$$
\mathbf{E}\left[u_{\pi}\left(X_{1}, X_{2}\right) \mid X_{2}=y\right]=0
$$

This shows that $u_{\pi}$ satisfies the condition of degeneracy (C.3).
Next, we show that $u_{\pi}$ satisfies the bound (C.4) with a function $g$ satisfying the moment condition (C.5). In fact, by Assumption K3 on the kernel $k$ and Assumption 1 on the target density $\pi$,

$$
\begin{aligned}
\left|u_{\pi}(x, y)\right| & \leq L^{2} K_{1}(1+|x|) \cdot(1+|y|)+L K_{1}(1+|x|+1+|y|)+K_{1}(1+d) \\
& \leq K_{1}(L+1)^{2}(\sqrt{d}+1+|x|) \cdot(\sqrt{d}+1+|y|) \\
& =: g(x) \cdot g(y)
\end{aligned}
$$

where $g(x)=\sqrt{K_{1}}(L+1)(\sqrt{d}+1+|x|)$ and the constant $L$ is defined in 2.6. To verify $g$ satisfies (C.5), we write

$$
\begin{align*}
& \mathbf{E}_{X \sim \pi}\left[g(X)^{k}\right]=\left(\sqrt{K_{1}}(L+1)\right)^{k} \mathbf{E}_{X \sim \pi}\left[(\sqrt{d}+1+|X|)^{k}\right] \\
& \leq\left(\sqrt{K_{1}}(L+1)\right)^{k} \mathbf{E}_{X \sim \pi}\left[(\sqrt{d}+1+|m|+|X-m|)^{k}\right] \\
& =\left(\sqrt{K_{1}}(L+1)\right)^{k}\left((\sqrt{d}+1+|m|)^{k}+\sum_{j=1}^{k}\binom{k}{j}(\sqrt{d}+1+|m|)^{k-j} \mathbf{E}_{X \sim \pi}|X-m|^{j}\right) \tag{C.13}
\end{align*}
$$

Thanks to PropositionC.1, we have for any $j \geq 1$,

$$
\begin{align*}
\mathbf{E}_{X \sim \pi}|X-m|^{j} & \leq\left(\mathbf{E}_{X \sim \pi}|X-m|^{2 j}\right)^{1 / 2} \\
& \leq(2 j)^{1 / 2} \cdot\left((2 v \sqrt{d})^{2 j}+\frac{1}{2}\left(\frac{4 v}{\sqrt{2}}\right)^{2 j} \cdot(2 j)^{j}\right)^{1 / 2} \\
& \leq(2 j)^{1 / 2} \cdot(2 \max (4 v, 1) \cdot \sqrt{d} \cdot \sqrt{j})^{j}  \tag{C.14}\\
& \leq\left(2 e^{1 / e} \max (4 v, 1) \cdot \sqrt{d} \cdot \sqrt{j}\right)^{j}
\end{align*}
$$

where we have used the simple fact that $(2 j)^{1 /(2 j)} \leq e^{1 / e}$ for any $j \geq 1$ in the last inequality. Plugging (C.14) into (C.13) yields that

$$
\begin{align*}
& \mathbf{E}_{X \sim \pi}\left[g(X)^{k}\right] \leq 2\left(\sqrt{K_{1}}(L+1)\right)^{k}\left(\sqrt{d}+1+|m|+2 e^{1 / e} \max (4 v, 1) \sqrt{d} \cdot \sqrt{k}\right)^{k}  \tag{C.15}\\
& =2\left(\sqrt{K_{1}}(L+1)\right)^{k} \exp \left(k \log \left(\sqrt{d}+1+|m|+2 e^{1 / e} \max (4 v, 1) \sqrt{d} \cdot \sqrt{k}\right)\right) .
\end{align*}
$$

Using the fact that $\log (a+b)-\log (a)=\log (1+b / a) \leq b / a$ for all $a, b \geq 1$, one has

$$
\begin{align*}
& \exp \left(k \log \left(\sqrt{d}+1+|m|+2 e^{1 / e} \max (4 v, 1) \sqrt{d} \cdot \sqrt{k}\right)\right) \\
& \leq \exp (k \log (\underbrace{2 e^{1 / e} \max (4 v, 1) \sqrt{d}}_{=: A} \cdot \sqrt{k})) \cdot \exp (\sqrt{k} \cdot \underbrace{\frac{\sqrt{d}+1+|m|}{2 e^{1 / e} \max (4 v, 1) \sqrt{d}}}_{=: B}) \tag{C.16}
\end{align*}
$$

Since by assumption $|m| \leq m^{*} \sqrt{d}$ and $d \geq 1$, we have

$$
B \leq \frac{2+m^{*}}{2 e^{1 / e} \max (4 v, 1)}=: \tilde{B}
$$

As a consequence of above and the fact that $k!\geq\left(\frac{k}{3}\right)^{k}$ for any $k \in \mathbb{N}_{+}$,

$$
\begin{align*}
\exp (k \log & \left.\left(\sqrt{d}+1+|m|+2 e^{1 / e} \max (4 v, 1) \sqrt{d} \cdot \sqrt{k}\right)\right) \\
& \leq \exp (k \log k / 2) \cdot \exp (k(\log A+\tilde{B})) \\
& =\left(\frac{k}{3}\right)^{k / 2} \cdot\left(\sqrt{3} A \exp \left(\tilde{B}+\frac{1}{2}\right)\right)^{k}  \tag{C.17}\\
& \leq k!\cdot\left(\sqrt{3} A \exp \left(\tilde{B}+\frac{1}{2}\right)\right)^{k} .
\end{align*}
$$

Combining this with C.15) implies that the moment bound assumption (C.5) holds with the constants

$$
J=\sqrt{3 K_{1}}(L+1) A \exp \left(\tilde{B}+\frac{1}{2}\right) \text { and } \xi=2 J
$$

Therefore it follows from the definition of $\operatorname{KSD}\left(P_{n}, \pi\right)$ in (C.1) and the concentration bound (C.7) implied by TheoremC. 1 that with at least probability $1-\delta$,

$$
\operatorname{KSD}\left(P_{n}, \pi\right) \leq \frac{C}{\sqrt{n}}
$$

with the constant

$$
C=J\left(\frac{\log \left(\frac{C_{1}}{\delta}\right)}{C_{2}}+2 \sqrt{\frac{\log \left(\frac{C_{1}}{\delta}\right)}{C_{2}}}\right)
$$

Since by definition the constant $A=\mathcal{O}(\sqrt{d})$ for large $d$, we have that the constant $C=\mathcal{O}(\sqrt{d})$. This completes the proof.

## D Summarizing Propositions 3.1 - 3.3

The theorem below summarizes the Propositions $3.1-3.3$ above, serving as one of the ingredients for proving Theorem 2.1
Theorem D.1. Let $\pi$ be a probability measure on $\mathbb{R}^{d}$ and let $P_{n}=\frac{1}{n} \sum_{i=1}^{n} \delta_{X_{i}}$ be the empirical measure associated to the i.i.d. samples $\left\{X_{i}\right\}_{i=1}^{n}$ drawn from $\pi$. Then we have the following:

1. If $\pi$ satisfies $M_{3}=\mathbf{E}_{X \sim \pi}|X|^{3}<\infty$, then there exists a realization of empirical measure $P_{n}$ such that

$$
\mathcal{W}_{1}\left(P_{n}, \pi\right) \leq C \cdot \begin{cases}n^{-1 / 2}, & d=1 \\ n^{-1 / 2} \log n, & d=2 \\ n^{-1 / d}, & d \geq 3\end{cases}
$$

where the constant $C$ depends only on $M_{3}$.
2. If $k$ satisfies Assumption $K 2$ with constant $K_{0}$, then there exists a realization of empirical measure $P_{n}$ such that

$$
\operatorname{MMD}\left(P_{n}, \pi\right) \leq \frac{C}{\sqrt{n}}
$$

where the constant $C$ depending only on $K_{0}$.
3. If $\pi$ satisfies Assumption 1 and 2 and $k$ satisfies Assumption $K 3$ with constant $K_{1}$, then there exists a realization of empirical measure $P_{n}$ such that

$$
K S D\left(P_{n}, \pi\right) \leq C \sqrt{\frac{d}{n}}
$$

where the constant $C$ depends only on $L, K_{1}, m^{*}, v$.

## E Semi-discrete optimal transport with quadratic cost

## E. 1 Structure theorem of optimal transport map

We recall the structure theorem of optimal transport map between $\mu$ and $\nu$ under the assumption that $\mu$ does not give mass to null sets.
Theorem E. 1 ([8, Theorem 2.9 and Theorem 2.12]). Let $\mu$ and $\nu$ be two probability measures on $\mathbb{R}^{d}$ with finite second moments. Assume that $\mu$ is absolutely continuous with respect to the Lebesgue measure. Consider the functionals $\mathcal{K}$ and $\mathcal{J}$ defined in Monge's problem (4.1) and dual Kantorovich problem (4.2) with $c=\frac{1}{2}|x-y|^{2}$. Then
(i) there exists a unique solution $\pi$ to Kantorovich's problem, which is given by $\pi(d x d y)=(\mathbf{I d} \times$ $T)_{\#} \mu$ where $T(x)=\nabla \bar{\varphi}(x) \mu$-a.e. $x$ for some convex function $\bar{\varphi}: \mathbb{R}^{d} \rightarrow \mathbb{R}$. In another word, $T(x)=\nabla \bar{\varphi}(x)$ is the unique solution to Monge's problem.
(ii) there exists an optimal pair $\left(\varphi(x), \varphi^{c}(y)\right)$ or $\left(\psi^{c}(x), \psi(y)\right)$ solving the dual Kantorovich's problem, i.e. $\sup _{(\varphi, \psi) \in \Phi_{c}} \mathcal{J}(\varphi, \psi)=\mathcal{J}\left(\varphi, \varphi^{c}\right)=\mathcal{J}\left(\psi^{c}, \psi\right)$;
(iii) the function $\bar{\varphi}(x)$ can be chosen as $\bar{\varphi}(x)=\frac{1}{2}|x|^{2}-\varphi(x)\left(\right.$ or $\left.\bar{\varphi}(x)=\frac{1}{2}|x|^{2}-\psi^{c}(x)\right)$ where $\left(\varphi(x), \varphi^{c}(y)\right)\left(\right.$ or $\left.\left(\psi^{c}(x), \psi(y)\right)\right)$ is an optimal pair which maximizes $\mathcal{J}$ within the set $\Phi_{c}$.

## E. 2 Proof of Theorem 4.2

Recall that the dual Kantorovich problem in the semi-discrete case reduces to maximizing the following functional

$$
\begin{equation*}
\mathcal{F}(\psi)=\int \inf _{j}\left(\frac{1}{2}\left|x-y_{j}\right|^{2}-\psi_{j}\right) \rho(x) d x+\sum_{j=1}^{n} \psi_{j} \nu_{j} \tag{E.1}
\end{equation*}
$$

Proof of Theorem4.2 relies on two useful lemmas on the functional $\mathcal{F}$. The first lemma below shows that the functional $\mathcal{F}$ is concave, whose proof adapts that of [5, Theorem 2] for semi-discrete optimal transport with the quadratic cost.
Lemma E.1. Let $\rho$ be a probability density on $\mathbb{R}^{d}$. Let $\left\{y_{j}\right\}_{j=1}^{n} \subset \mathbb{R}^{d}$ and let $\left\{\nu_{j}\right\}_{j=1}^{n} \subset[0,1]$ be such that $\sum_{j=1}^{n} \nu_{j}=1$. Then the functional $\mathcal{F}$ be defined by (E.1) is concave.

Proof. Let $\mathcal{A}: \mathbb{R}^{d} \rightarrow\{1,2, \cdots, n\}$ be an assignment function which assigns a point $x \in \mathbb{R}^{d}$ to the index $j$ of some point $y_{j}$. Let us also define the function

$$
\widetilde{\mathcal{F}}(\mathcal{A}, \psi)=\int\left(\frac{1}{2}\left|x-y_{\mathcal{A}(x)}\right|^{2}-\psi_{\mathcal{A}(x)}\right) \rho(x) d x+\sum_{j=1}^{n} \psi_{j} \nu_{j}
$$

Then by definition $\mathcal{F}(\psi)=\inf _{\mathcal{A}} \widetilde{\mathcal{F}}(\mathcal{A}, \psi)$. Denote $\mathcal{A}^{-1}(j)=\left\{x \in \mathbb{R}^{d} \mid \mathcal{A}(x)=j\right\}$. Then

$$
\begin{aligned}
\widetilde{\mathcal{F}}(\mathcal{A}, \psi) & =\sum_{j=1}^{n}\left[\int_{\mathcal{A}^{-1}(j)}\left(\frac{1}{2}\left|x-y_{j}\right|^{2}-\psi_{j}\right) \rho(x) d x+\psi_{j} \nu_{j}\right] \\
& =\sum_{j=1}^{n} \int_{\mathcal{A}^{-1}(j)} \frac{1}{2}\left|x-y_{j}\right|^{2} \rho(x) d x+\sum_{j=1}^{n} \psi_{j}\left(\nu_{j}-\int_{\mathcal{A}^{-1}(j)} \rho(x) d x\right) .
\end{aligned}
$$

Since the function $\widetilde{\mathcal{F}}(\mathcal{A}, \psi)$ is affine in $\psi$ for every $\mathcal{A}$, it follows that $\mathcal{F}(\psi)=\inf _{\mathcal{A}} \widetilde{\mathcal{F}}(\mathcal{A}, \psi)$ is concave.

The next lemma computes the gradient of the concave function $\mathcal{F}$; see [5, Section 7.4] for the corresponding result with general transportation cost.
Lemma E.2. Let $\rho$ be a probability density on $\mathbb{R}^{d}$. Let $\left\{y_{j}\right\}_{j=1}^{n} \subset \mathbb{R}^{d}$ and let $\left\{\nu_{j}\right\}_{j=1}^{n} \subset[0,1]$ be such that $\sum_{j=1}^{n} \nu_{j}=1$. Denote by $P_{j}(\psi)$ the power diagram associated to $\psi$ and $y_{j}$. Then

$$
\begin{equation*}
\partial_{\psi_{i}} \mathcal{F}(\psi)=\nu_{i}-\mu\left(P_{i}(\psi)\right)=\nu_{i}-\int_{P_{i}(\psi)} \rho(x) d x \tag{E.2}
\end{equation*}
$$

Proof. By the definition of $\mathcal{F}$ in (E.1), we rewrite $\mathcal{F}$ as

$$
\begin{aligned}
\mathcal{F}(\psi) & =\int \frac{1}{2}|x|^{2} \rho(d x)+\int \inf _{j}\left\{-x \cdot y_{j}+\frac{1}{2}\left|y_{j}\right|^{2}-\psi_{j}\right\} \rho(x) d x+\sum_{j=1}^{n} \psi_{j} \nu_{j} \\
& =\int \frac{1}{2}|x|^{2} \rho(d x)-\int \sup _{j}\left\{x \cdot y_{j}+\psi_{j}-\frac{1}{2}\left|y_{j}\right|^{2}\right\} \rho(x) d x+\sum_{j=1}^{n} \psi_{j} \nu_{j}
\end{aligned}
$$

To prove (E.2), it suffices to prove that

$$
\begin{equation*}
\partial_{\psi_{i}}\left(\int \sup _{j}\left\{x \cdot y_{j}+\psi_{j}-\frac{1}{2}\left|y_{j}\right|^{2}\right\} \rho(x) d x\right)=\int_{P_{i}(\psi)} \rho(x) d x \tag{E.3}
\end{equation*}
$$

Note that the partial derivative on the left side of above makes sense since $g(x, \psi):=\sup _{j}\{x$. $\left.y_{j}+\psi_{j}-\frac{1}{2}\left|y_{j}\right|^{2}\right\}$ is convex with respect to $(x, \psi)$ on $\mathbb{R}^{d} \times \mathbb{R}^{d}$ so that the resulting integral against the measure $\rho$ is also convex (and hence Lipschitz) in $\psi$. To see E.3), since $g(x, \psi)$ is convex and piecewise linear in $\psi$ for any fixed x , it is easy to observe that

$$
\partial_{\psi_{i}} g(x, \psi)=\delta_{i j} \text { if } x \in\left\{\left.x \in \mathbb{R}^{d}\left|x \cdot y_{j}+\psi_{j}-\frac{1}{2}\right| y_{j}\right|^{2}=g(x, \psi)\right\}
$$

However, by subtracting $\frac{1}{2}|x|^{2}$ on both sides of the equation inside the big parenthesis and then flipping the sign one sees that

$$
\left\{\left.x \in \mathbb{R}^{d}\left|x \cdot y_{j}+\psi_{j}-\frac{1}{2}\right| y_{j}\right|^{2}=g(x, \psi)\right\}=P_{j}(\psi)
$$

Namely we have obtained that

$$
\partial_{\psi_{i}} g(x, \psi)=\delta_{i, j} \text { if } x \in P_{j}(\psi)
$$

In particular, this implies that $\psi \rightarrow g(x, \psi)$ is 1-Lipschitz in $\psi$ uniformly with respect to $x$. Finally since $\rho(x)$ is a probability measure, the desired identity E. 2 follows from the equation above and the dominated convergence theorem. This completes the proof of the lemma.

With the lemmas above, we are ready to prove Theorem 4.2. In fact, according to Lemma E.1 and Lemma E. $2 . \psi=\left(\psi_{1}, \cdots, \psi_{n}\right)$ is a maximizer of the functional $\mathcal{F}$ if and only if

$$
\partial_{\psi_{i}} \mathcal{F}(\psi)=\nu_{i}-\mu\left(P_{i}(\psi)\right)=\nu_{i}-\int_{P_{i}(\psi)} \rho(x) d x=0
$$

Since the dual Kantorovich problem in the semi-discrete setting reduces to the problem of maximizing $\mathcal{F}$, it follows from Theorem E. 1 that the optimal transport map $T$ solving the semi-discrete Monge's problem (4.4) is given by $\overline{T(x)}=\nabla \bar{\varphi}(x)$ where $\bar{\varphi}(x)=\frac{1}{2}|x|^{2}-\varphi(x)$ and $\left.\varphi(x)=\min _{j} \frac{1}{2} \right\rvert\, x-$ $\left.y_{j}\right|^{2}-\psi_{j}$. Consequently,

$$
\begin{aligned}
\bar{\varphi}(x) & =\frac{1}{2}|x|^{2}-\varphi(x) \\
& =\frac{1}{2}|x|^{2}-\left(\min _{j}\left\{\frac{1}{2}\left|x-y_{j}\right|^{2}-\psi_{j}\right\}\right) \\
& =\max _{j}\left\{x \cdot y_{j}+m_{j}\right\}
\end{aligned}
$$

with $m_{j}=\psi_{j}-\frac{1}{2}\left|y_{j}\right|^{2}$. Moreover, noticing that $\varphi(x)$ can be rewritten as

$$
\varphi(x)=\frac{1}{2}\left|x-y_{j}\right|^{2}-\psi_{j} \text { if } x \in P_{j}(\psi)
$$

one obtains that $T(x)=\nabla \bar{\varphi}(x)=y_{j}$ if $x \in P_{j}(\psi)$.

## E. 3 Proof of Proposition 4.1

Let us first consider the case that $n=2^{k}$ for some $k \in \mathbb{N}$. Then

$$
\bar{\varphi}(x)=\max _{j=1, \cdots, 2^{k}}\left\{x \cdot y_{j}+m_{j}\right\}=\max _{j=1, \cdots, 2^{k-1}} \max _{i \in\{2 j-1,2 j\}}\left\{x \cdot y_{i}+m_{i}\right\}
$$

Let us define maps $\varphi_{n}: \mathbb{R}^{n} \rightarrow \mathbb{R}^{n / 2}$ and $\psi: \mathbb{R}^{d} \rightarrow \mathbb{R}^{n}$ by setting

$$
\left[\varphi_{n}(z)\right]_{i}=\max \left\{z_{2 i-1}, z_{2 i}\right\}, i=1, \cdots, n / 2 \text { and }[\psi(x)]_{j}=x \cdot y_{j}+m_{j}, j=1, \cdots, n
$$

Then by definition it is straightforward that

$$
\begin{equation*}
\bar{\varphi}(x)=\left(\varphi_{2} \circ \varphi_{4} \circ \cdots \circ \varphi_{n / 2} \circ \varphi_{n} \circ \psi\right)(x) \tag{E.4}
\end{equation*}
$$

By defining

$$
Y=\left(\begin{array}{c}
y_{1}^{T} \\
y_{2}^{T} \\
\vdots \\
y_{n}^{T}
\end{array}\right), m=\left(\begin{array}{c}
m_{1} \\
m_{2} \\
\vdots \\
m_{n}
\end{array}\right)
$$

we can write the map $\psi$ as

$$
\begin{equation*}
\psi(x)=Y \cdot x+m \tag{E.5}
\end{equation*}
$$

Moreover, thanks to the following simple equivalent formulation of the maximum function:

$$
\begin{aligned}
\max (a, b) & =\operatorname{ReLU}(a-b)+\operatorname{ReLU}(b)-\operatorname{ReLU}(-b) \\
& =h^{T} \cdot \operatorname{ReLU}\left(A\binom{a}{b}\right)
\end{aligned}
$$

where

$$
A=\left(\begin{array}{cc}
1 & -1 \\
0 & 1 \\
0 & -1
\end{array}\right), \quad h=\left(\begin{array}{c}
1 \\
1 \\
-1
\end{array}\right)
$$

we can express the map $\varphi_{n}$ in terms of a two-layer neural network as follows

$$
\begin{equation*}
\varphi_{n}(z)=H_{n / 2} \cdot \operatorname{ReLU}\left(A_{n} \cdot z\right) \tag{E.6}
\end{equation*}
$$

where $A_{n}=\oplus^{n} A$ and $H_{n}=\oplus^{n} h$. Finally, by combining (E.4), E.5) and (E.6), one sees that $\bar{\varphi}$ can be expressed in terms of a DNN of width $n$ and depth $\log n$ with parameters $\left(W^{\ell}, b^{\ell}\right)_{\ell=1}^{L+1}$ defined by

$$
\begin{aligned}
& W^{0}=A_{n} \cdot Y, \quad b^{0}=A_{n} \cdot m \\
& W^{1}=A_{n / 2} \cdot H_{n / 2}, \quad b^{1}=0 \\
& W^{2}=A_{n / 4} \cdot H_{n / 4}, \quad b^{2}=0 \\
& \cdots, \\
& W_{L-1}=A_{2} \cdot H_{2}, \quad b^{L-1}=0 \\
& W_{L}=H_{1}, \quad b^{L}=0
\end{aligned}
$$

In the general case where $\log _{2} n \notin \mathbb{N}$, we set $k=\left\lceil\log _{2} n\right\rceil$ so that $k$ is smallest integer such that $2^{k}>n$. By redefining $y_{j}=0$ and $m_{j}=0$ for $j=n+1, \cdots, 2^{k}$, we may still write $\bar{\varphi}(x)=\max _{j=1, \cdots, 2^{k}}\left\{x \cdot y_{j}+m_{j}\right\}$ so that the analysis above directly applies.

## F Proof of Main Theorem 2.1

The proof follows directly from Theorem D.1 and Theorem 4.1. Indeed, on the one hand, the quantitative estimate for the convergence of the empirical measure $P_{n}$ directly translates to the sample complexity bounds in Theorem 2.1 with a given error $\varepsilon$. On the other hand, Theorem 4.1 provides a push-forward from $p_{x}$ to the empirical measure $P_{n}$ based on the gradient of a DNN.

## References

[1] S. Bobkov and M. Ledoux. One-dimensional empirical measures, order statistics, and Kantorovich transport distances. Mem. Amer. Math. Soc., 261(1259):0, 2019.
[2] I. Borisov. Approximation of distributions of von mises statistics with multidimensional kernels. Siberian Mathematical Journal, 32(4):554-566, 1991.
[3] D. Hsu, S. Kakade, T. Zhang, et al. A tail inequality for quadratic forms of subgaussian random vectors. Electronic Communications in Probability, 17, 2012.
[4] J. Lei. Convergence and concentration of empirical measures under wasserstein distance in unbounded functional spaces. Bernoulli, 26(1):767-798, 2020.
[5] B. Lévy and E. L. Schwindt. Notions of optimal transport theory and how to implement them on a computer. Computers \& Graphics, 72:135-148, 2018.
[6] Q. Liu, J. Lee, and M. Jordan. A kernelized stein discrepancy for goodness-of-fit tests. In International conference on machine learning, pages 276-284, 2016.
[7] B. Sriperumbudur. On the optimal estimation of probability measures in weak and strong topologies. Bernoulli, 22(3):1839-1893, 2016.
[8] C. Villani. Topics in optimal transportation. Number 58. American Mathematical Soc., 2003.

