A Negative association properties of ESR measures

Proposition 1. There exist ESR measures that are not SR.

Proof. Recall [11, Thm. 4.1] that a real bivariate affine polynomial p(x, y) is stable if and only if

$$\partial_x p \partial_y p - p \,\partial_{xy} p \ge 0.$$

For an E-DPP with a kernel $L \in \mathbb{R}^{2 \times 2}$, this is exactly $l_{11}^p l_{22}^p \ge \det(L)^p$, which is clearly true for all $p \ge 0$ as L must be positive semi-definite.

However, stability does not hold in general. To obtain a counterexample, consider the generating polynomial p(x, y, z) for an E-DPP of dimension 3 (writing $d = \det(L)$ as a shorthand):

 $p(x, y, z) := d^p + l_{11}^p yz + l_{22}^p xz + l_{33}^p xy + \det(L[1, 2])^p z + \det(L[1, 3])^p y + \det(L[2, 3])^p x + xyz.$

If p(x, y, z) is SR, p must be stable through conditioning [11, Theorem 4.1]; hence, p(x, y, 1) must also be stable. Writing $d_{ij} = \det(L[i, j])$, this requires that

$$p(x, y, 1) = d^p + d^p_{12} + (d^p_{22} + d^p_{23})x + (d^p_{11} + d^p_{13})y + (d^p_{33} + 1)xy$$

be stable. Since p(x, y, 1) is a real bivariate affine polynomial, we must then have

$$(d_{22}^p + d_{23}^p)(d_{11}^p + d_{13}^p) \ge (d^p + d_{12}^p)(d_{33}^p + 1).$$

Finally, one can verify that this last inequality is easily violated for several choices of (non-diagonal) positive semi-definite matrices. $\hfill \Box$

Theorem 1. There exists $\epsilon > 0$ such that for any $p \in [1 - \epsilon, 1 + \epsilon]$ and $n \in \mathbb{N}$, the set $\{L \in \mathbb{R}^{n \times n} : E\text{-DPP}(L, p) \text{ is SR}\}$ is strictly greater than the set of block-diagonal matrices with 2×2 blocks.

Proof. We write $d_S^i = \det(L[S \cup \{i\}))^p$ and when possible $d_{ijk} = \det(L[\{i, j, k\}])^p$.

Let $L \succeq 0 \in \mathbb{R}^{n \times n}$. The associated E-DPP is SR if and only if $P_{ij}(z) \ge 0$ where $P_{ij}(z)$ is defined for any $1 \le i \ne j \le n$ as

$$P_{r}(z) = \frac{\partial f}{\partial z_{i}}(z)\frac{\partial f}{\partial z_{j}}(z) - f(z)\frac{\partial^{2} f}{\partial z_{i}\partial z_{j}}(z)$$
$$= \Big(\sum_{S\in\mathcal{Y}'}d_{S}^{i}z^{S}\Big)\Big(\sum_{S\in\mathcal{Y}'}d_{S}^{j}z^{S}\Big) - \Big(\sum_{S\in\mathcal{Y}'}d_{S}^{ij}z^{S}\Big)\Big(\sum_{S\in\mathcal{Y}'}d_{S}z^{S}\Big).$$

where we write $\mathcal{Y}' = \{S \in [n], i \notin S, j \notin S\}$ and $z = (z_1, \ldots, z_n)$ where components z_i and z_j are removed. Now, choose $k \in [n] \setminus \{i, j\}$, and write \tilde{z} the vector z without component z_k . Write also $\mathcal{Y}'_k = \{S \in \mathcal{Y}', k \notin S\}$. Then,

$$\begin{split} P_{ij}(z) =& (\sum_{k \notin S} d_S^i \tilde{z}^S + z_k \sum_{k \in S} d_S^i \tilde{z}^S) (\sum_{k \notin S} d_S^j \tilde{z}^S + z_k \sum_{k \in S} d_S^j \tilde{z}^S) \\ &- (\sum_{k \notin S} d_S^{ij} \tilde{z}^S + z_k \sum_{k \in S} d_S^{ij} \tilde{z}^S) (\sum_{k \notin S} d_S \tilde{z}^S + z_k \sum_{k \in S} d_S \tilde{z}^S) \\ &= \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ik} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{jk} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{k} \tilde{z}^S \right) z_k^2 \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ik} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^j \tilde{z}^S + \sum_{S \in \mathcal{Y}'_k} d_S^i \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{jk} \tilde{z}^S \right) \\ &- \sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ij} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{k} \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ij} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{k} \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{j} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ij} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{k} \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{j} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{j} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S^{ijkk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{jkk} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{jkk} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S^{ijk} \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S^{jkk} \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S - \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum_{S \in \mathcal{Y}'_k} d_S \tilde{z}^S \right) z_k \\ &+ \left(\sum$$

Hence, L yields a SR E-DPP measure if and only if the following inequalities hold for all i, j, k:

$$(B/2)^2 \le AC; \quad A \ge 0; \quad C \ge 0.$$
 (A.1)

When n = 3, Eq. (A.1) reduces to the following arithmetic-geometric inequality, as $\mathcal{Y}'_k = \{\emptyset\}$:

$$\left(\frac{d_i d_{jk} + d_j d_{ik} - d_{ij} d_k - d_{ijk}}{2}\right)^2 \le (d_i d_j - d_{ij})(d_{jk} d_{ik} - d_k d_{ijk}) \tag{A.2}$$

One can easily obtain 3 positive semi-definite matrices L which verify Eq. (A.2) strictly for p = 1; in particular, by continuity, there exists $\epsilon > 0$ such that the E-DPP generated by the kernel L and power $p \in [1 - \epsilon, 1 + \epsilon]$ will still verify Eq. (A.2).

Then, as a block-diagonal matrix such that each diagonal block yields SR E-DPPs also yields a SR E-DPP, we can thus generate block-diagonal matrices of any size n such that the blocks are either L or 2×2 matrices, which all yield SR E-DPPs for $p \in [1 - \epsilon, 1 + \epsilon]$.

B *r*-closeness

Proposition 2. Let μ be an SR measure over $2^{[n]}$, and define ν to be the ESR measure such that $\mu(S) = \alpha \mu(S)^p$ for a given $\alpha \in \mathbb{R}$. Then

$$r(\mu, \nu) \leq \max_{S \in \operatorname{supp}(\nu)} \left[\mu(S)^{-|p-1|} \right] < \infty.$$

Proof. Let μ, ν be as in the proposition statement, and consider $S \in \text{supp}(\nu)$: $\nu(S) > 0$. Recall that $\sum_{T} \nu(T) = 1$.

$$\frac{\nu(S)}{\mu(S)} = \frac{\mu(S)^p}{\mu(S)\sum_T \mu(T)^p}$$

If $p\leq 1,$ we have $\sum_T \mu(T)^p\geq 1$ and $\mu(S)^p\geq \mu(S),$ and so

$$(\min_{S} \mu(S))^{p-1} \le \frac{\sum_{T} \mu(T)}{\sum_{T} \mu(T)^{p}} = \frac{1}{\sum_{T} \mu(T)^{p}} \le \frac{\mu(S)^{p}}{\mu(S) \sum_{T} \mu(T)^{p}} \le \mu(S)^{p-1} \le \frac{1}{\min_{S} \mu(S)^{1-p}}$$

where the left inequality is obtained by noticing that $\frac{a+b}{c+d} \ge \min\left(\frac{a}{c}, \frac{b}{d}\right)$.

Similarly, for $p\geq 1,$ we have $\sum_T \mu(T)^p\leq 1$ and $\mu(T)^p\leq \mu(T),$ and so

$$\min_{S} \mu(S)^{p-1} \le \frac{\mu(S)^p}{\mu(S)} \le \frac{\mu(S)^p}{\mu(S)\sum_T \mu(T)^p} \le \frac{\sum_T \mu(T)}{\sum_T \mu(T)^p} \le \max \mu(S)^{1-p} = \frac{1}{\min \mu(S)^{p-1}}.$$

C Bounds on mixing times

Theorem 2. Let μ, ν be measures over $2^{[n]}$ such that μ is SR and ν is ESR. Sampling from ν via Alg. 1 with μ as a proposal distribution has a mixing time $\tau(\epsilon)$ such that

$$\tau_S(\epsilon) \le 2r(\mu, \nu^p)\log\frac{1}{\epsilon}.$$

Proof. Alg. 1 has a state-independent proposal distribution μ , and hence its mixing time is governed by a ratio of probabilities: Cai [14] showed that, after t iterations,

$$\max_{S,T} d_{\text{TV}}(\nu_{(t)}(\cdot \mid S), \nu_{(t)}(\cdot \mid T)) = \left(1 - \frac{1}{\max_U \nu(U)/\mu(U)}\right)^t$$

where $\nu_{(t)}(U \mid S)$ is the probability of being in state U after t iterations when starting from set S, and d_{TV} is the total variation distance.

Hence, following [14, Cor.1], we obtain
$$\tau_S(\epsilon) \le 2 \max_U \frac{\nu(U)}{\mu(U)} \log \frac{1}{\epsilon} \le 2r(\mu, \nu^p) \log \frac{1}{\epsilon}$$
.

Theorem 3. Let ν be a k-homogeneous ESR measure over $2^{[n]}$. The mixing time for Alg. 2 with initialization S is bounded in expectation by

$$\tau_S(\epsilon) \le \inf_{\mu \in SR} 2nk r(\mu, \nu)^2 \log \frac{1}{\epsilon \nu(S)}$$

Proof. This bound is based on a comparison method [17], and relates the mixing time to the spectral gap. Let $1 = \mu_1 \ge \mu_2 \ge \ldots \ge -1$ be the eigenvalues of the state transition matrix of the chain. The *spectral gap* is $\gamma = 1 - \max\{|\mu|; \mu \text{ is an eigenvalue and } \mu \ne 1\}$. γ directly translates into a bound on the mixing time [18]:

$$au_S(\gamma) \leq rac{1}{\gamma} \log\left(rac{1}{\epsilon \nu(S)}
ight).$$

The comparison method yields a bound on γ if we know a bound on $\tilde{\gamma}$ for a related chain with stationary distribution μ . Specifying [17, Thm 2.1] to this case yields $\gamma \geq \tilde{\gamma}\alpha_1/\alpha_2$, where

$$\alpha_{1} = \min_{S} \frac{\mu(S)}{\nu(S)} \ge \frac{1}{r(p)}$$

$$\alpha_{2} = \max_{T,U} \frac{\mu(T)}{\nu(T)} \frac{\min\{1, \mu(U)/\mu(T)\}}{\min\{1, \nu(U)/\nu(T)\}} \le \max_{T} \frac{\mu(T)}{\nu(T)} \le r(\mu, \nu)$$

Anari et al. [5] show that $\tilde{\gamma} \geq \frac{1}{2nk}$. Hence, we obtain $\tau_S(\gamma) \leq 2nk \cdot r(\mu, \nu)^2 \cdot \log \frac{1}{\epsilon \nu(S)}$.

D Bounds for E-DPPs with *L^p*-kernel proposal

We require the following power-mean inequality:

Theorem 8 (Specht [45]). Let $x_i > 0$ and $w_i \ge 0$ for $1 \le i \le N$ such that $\sum_i w_i = 1$. Let $p < q \in \mathbb{R}$ such that $pq \ne 0$. Then, letting $\kappa = \frac{\max x_i}{\min x_i}$,

$$1 \leq \frac{M_q(\boldsymbol{w}; \boldsymbol{x})}{M_p(\boldsymbol{w}; \boldsymbol{x})} \leq \left(\frac{q-p}{q} \frac{\kappa^q - 1}{\kappa^q - \kappa^p}\right)^{\frac{1}{p}} \left(\frac{p}{q-p} \frac{\kappa^q - \kappa^p}{\kappa^p - 1}\right)^{\frac{1}{q}}$$

where the power mean $M_p(\boldsymbol{w}; \boldsymbol{x})$ is defined as

$$M_p(\boldsymbol{w};\boldsymbol{x}) := \left(\sum_{i=1}^N w_i x_i^p\right)^{\frac{1}{p}}.$$

Theorem 9. Let $L \in \mathbb{P}^n$ be a positive definite matrix and $S \subseteq [n]$. Then,

$$\det(L[S])^p \ge \det(L^p[S]), \qquad 0 \le p \le 1,$$

$$\det(L[S])^p \le \det(L^p[S]), \qquad p \ge 1.$$

Proof. From Lemma 1, there exists a vector w in the probability simplex, of size $\binom{n}{|S|}$, such that

$$\det(L[S]) = \sum_{J \subseteq [n], |J| = |S|} w_J \prod_{i \in J} \lambda_i.$$

Since $t \to t^p$ is convex for $p \ge 1$, Jensen's inequality shows that

$$\det(L[S])^p \le \sum_{J\subseteq [n], |J|=|S|} w_J \prod_{i\in J} \lambda_i^p = \det(L^p[S]),$$

where the latter equality follows due to L and L^p sharing the same eigenbasis. The same reasoning for p < 1 gives the other side of the inequality.

Theorem 7. Let μ be the distribution induced by a DPP with kernel L^p , and ν be the corresponding E-DPP such that $\nu(S) \triangleq \det(L[S])^p/Z_p$. Then $r(\mu, \nu) \leq r(\kappa_{\lfloor n/2 \rfloor}, p)$ where $r(\kappa, p)$ is defined by

$$r(\kappa, p) = \begin{cases} \left(\frac{p(\kappa-1)}{\kappa^{p}-1}\right)^{p} \left(\frac{(1-p)(\kappa-1)}{\kappa-\kappa^{p}}\right)^{1-p} & \text{for } 0 1 \end{cases}$$

Proof. We show the result for general DPPs; the result for k-DPPs follows the same exact reasoning. For $0 , it follows from Thm. 9 that <math>\det(L[S])^p \ge \det(L^p[S])$. Hence,

$$Z_p = \sum_{S \subseteq [n]} \det(L[S])^p \ge \sum_{S \subseteq [n]} \det(L^p[S]) = \det(I + L^p),$$

which entails $\frac{\det(I+L^p)}{Z_p} \leq 1$ whereby it remains to bound $\frac{\det(L[S])^p}{\det(L^p[S])}$.

Let $S \subseteq [n]$ of size k, and let λ be the vector of L's eigenvalues. We write $\lambda^S = \prod_{i \in S} \lambda_i$, and denote by $\lambda^{\wedge k}$ the $\binom{n}{k}$ -vector $(\lambda^S)_{S \subseteq [n], |S|=k}$. Using Lemma 1, there exists $w \in \mathbb{R}^{\binom{n}{k}}$ that sums to 1 such that

$$\frac{\det(L[S])^p}{\det(L^p[S])} = \frac{\left(\sum_{|S|=k} w_S \lambda^S\right)^p}{\sum_{|S|=k} w_S(\lambda^p)^S} = \left(\frac{M_1(\boldsymbol{w}; \boldsymbol{\lambda}^{\wedge k})}{M_p(\boldsymbol{w}; \boldsymbol{\lambda}^{\wedge k})}\right)^p < r(p)^p.$$

Where the last inequality follows from Thm. 8. To lower bound $r_p(S)$, the same reasoning gives us $\frac{\det(L[S]^p)}{\det(L^p[S])} \ge 1$ and hence

$$r_p(S) \ge \frac{\det(I+L^p)}{Z_p} \ge \min_{S} \frac{\det L^p[S]}{(\det L[S])^p}$$
$$\ge \left(\frac{M_p(\boldsymbol{w'};\boldsymbol{\lambda}^{\wedge k})}{M_1(\boldsymbol{w'};\boldsymbol{\lambda}^{\wedge k})}\right)^p \ge r(p)^{-p},$$

where the second inequality follows $\frac{a+b}{c+d} \ge \min(\frac{a}{c}, \frac{b}{d})$. The same reasoning yields the result for p > 1.

Finally, some algebra shows that for fixed p, r is an increasing function of κ , and so $r(\kappa_k, p)$ is upper bounded by $r(\kappa_{\lfloor n/2 \rfloor})$.

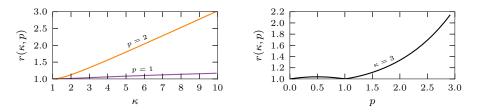
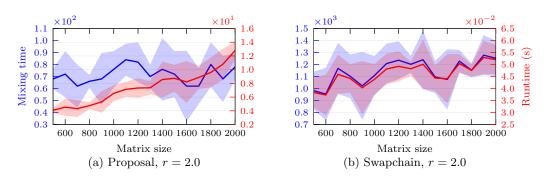


Figure 4: Evolution of the upper bound for $r(p, \kappa)$ from Thm. 7, which measures the *r*-closeness between the E-DPP with kernel L and the DPP with kernel L^p .



E Mixing time as a function of ground set size for E-DPPs

Figure 5: Influence of ground set size n on mixing time

F Additional Nystrom sampling results

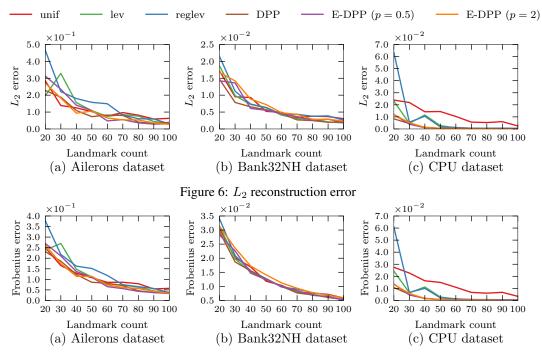


Figure 7: Frobenius norm reconstruction error