
Supplement to “Statistical bounds for entropic optimal transport: sample complexity and the central limit theorem”

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Throughout the supplement, the symbol C will be used to indicate an unspecified positive constant whose value may change from line to line. Subscripts will be used to indicate if C depends on any other parameters.

A Supplementary results

Lemma A.1. *If P is σ^2 subgaussian, then*

$$E_P \|X\|^{2k} \leq (2d\sigma^2)^k k!$$

for all nonnegative integers k , and

$$E_P e^{v \cdot X} \leq E_P e^{\|v\| \|X\|} \leq 2e^{\frac{d\sigma^2}{2} \|v\|^2}$$

for all $v \in \mathbb{R}^d$.

Proof. For the first claim, it suffices to take expectations of both sides of the inequality $\frac{\|X\|^{2k}}{(2d\sigma^2)^k k!} \leq e^{\frac{\|X\|^2}{2d\sigma^2}} - 1$ and use the assumption that P is σ^2 -subgaussian. To prove the second claim, we use the inequality $v \cdot X \leq \|v\| \|X\| \leq \frac{d\sigma^2}{2} \|v\|^2 + \frac{1}{2d\sigma^2} \|X\|^2$ and apply subgaussianity. \square

Proposition A.1. *Let P and Q be two σ^2 -subgaussian distributions. Then there exist smooth optimal potentials (f, g) for $S(P, Q)$ such that*

$$\begin{aligned} -d\sigma^2(1 + \frac{1}{2}(\|x\| + \sqrt{2d}\sigma)^2) - 1 &\leq f(x) \leq \frac{1}{2}(\|x\| + \sqrt{2d}\sigma)^2 \\ -d\sigma^2(1 + \frac{1}{2}(\|y\| + \sqrt{2d}\sigma)^2) - 1 &\leq g(y) \leq \frac{1}{2}(\|y\| + \sqrt{2d}\sigma)^2 \end{aligned}$$

and the dual optimality conditions (4) hold for all $x, y \in \mathbb{R}^d$.

Proof. Let (f_0, g_0) be any pair of optimal potentials. Since $(f_0 + K, g_0 - K)$ also satisfy the optimality conditions and $f_0 \in L_1(P)$ and $g_0 \in L_1(Q)$, we can assume without loss of generality that $E_P f_0(X) = E_Q g_0(Y) = \frac{1}{2}S(P, Q) \geq 0$. We define

$$\begin{aligned} f(x) &= -\log \int e^{g_0(y) - \frac{1}{2}\|x-y\|^2} dQ(y) \\ g(y) &= -\log \int e^{f(x) - \frac{1}{2}\|x-y\|^2} dP(x), \end{aligned}$$

for all $x, y \in \mathbb{R}^d$.

We need to check that these integrals are well defined. First, Jensen's inequality implies

$$\begin{aligned} g_0(y) &= -\log \int e^{f_0(x) - \frac{1}{2}\|x-y\|^2} dP(x) \\ &\leq -E_P f_0(X) + \frac{1}{2}E_P \|X - y\|^2 \\ &\leq \frac{1}{2}E_P \|X - y\|^2 \end{aligned}$$

for Q -a.e. y . Therefore

$$e^{g_0(y) - \frac{1}{2}\|x-y\|^2} \leq e^{\frac{1}{2}E_P \|X-y\|^2 - \frac{1}{2}\|x-y\|^2}$$

for Q -a.e. y . By Lemma A.1, $E_P \|X\|^2 \leq 2d\sigma^2$, which implies that $e^{g_0(y) - \frac{1}{2}\|x-y\|^2}$ is dominated by $e^{d\sigma^2 + (\|x\| + \sqrt{2}d\sigma)\|y\|}$. Subgaussianity implies

$$\int e^{d\sigma^2 + (\|x\| + \sqrt{2}d\sigma)\|y\|} dQ(y) \leq 2e^{d\sigma^2(1 + \frac{1}{2}(\|x\| + \sqrt{2}d\sigma)^2)} < \infty$$

Therefore $f(x)$ is well defined for all $x \in \mathbb{R}^d$. The same argument used to bound g_0 holds for f as well, which implies that g is also well defined. Therefore our definitions of f and g are valid on the whole space, and moreover the claimed lower bounds on f and g hold. Jensen's inequality combined with the inequalities $E_Q g_0(Y) \geq 0$ and $E_P f(X) \geq 0$ yield the upper bounds. The smoothness of f and g follows from an easy application of dominated convergence.

We now show that (f, g) are optimal potentials. By construction $\int e^{f(x) + g(y) - \frac{1}{2}\|x-y\|^2} dP(x) = 1$ for all $y \in \mathbb{R}^d$. Now, note that

$$\begin{aligned} \int e^{f(x) + g(y) - \frac{1}{2}\|x-y\|^2} dP(x) dQ(y) &= \int e^{f(x) + g_0(y) - \frac{1}{2}\|x-y\|^2} dP(x) dQ(y) \\ &= \int e^{f_0(x) + g_0(y) - \frac{1}{2}\|x-y\|^2} dP(x) dQ(y). \end{aligned}$$

Jensen's inequality yields

$$\begin{aligned} \int (f - f_0)(x) dP(x) + \int (g - g_0)(y) dQ(y) &\geq -\log \int e^{f_0(x) - f(x)} dP(x) - \log \int e^{g_0(y) - g(y)} dQ(y) \\ &= -\log \int e^{f_0(x) + g_0(y) - \frac{1}{2}\|x-y\|^2} dP(x) dQ(y) \\ &\quad - \log \int e^{f(x) + g_0(y) - \frac{1}{2}\|x-y\|^2} dP(x) dQ(y) \\ &= 0. \end{aligned}$$

Since (f_0, g_0) maximizes (3), so does (f, g) . Therefore (f, g) are optimal potentials. In particular, this implies that $\int (g - g_0)(y) dQ(y) = \log \int e^{g_0(y) - g(y)} dQ(y)$, and hence $g = g_0$ Q -almost surely by the strict concavity of the logarithm function. We obtain that $\int e^{f(x) + g(y) - \frac{1}{2}\|x-y\|^2} dQ(y) = \int e^{f(x) + g_0(y) - \frac{1}{2}\|x-y\|^2} dQ(y) = 1$ for all $x \in \mathbb{R}^d$. \square

Proposition A.2. *Let P_n, Q_n be empirical measures, P and Q both assumed subgaussian. There exist (f_n, g_n) optimal potentials for (P_n, Q_n) such that (f_n, g_n) converges uniformly in compacts to optimal potentials (f, g) for P and Q .*

Proof. The proof is inspired by Feydy et al. (2019) and we divide it in two steps:

Step 1 By using the following extended version of the Arzela-Ascoli theorem we find a convergent subsequence: suppose h_n is a sequence of functions in \mathbb{R}^d satisfying

(a) Local equicontinuity: for each $x_0 \in \mathbb{R}^d$ and $\epsilon > 0$, there is a $\delta > 0$ such that

$$\|x - x_0\| < \delta \quad \text{implies} \quad |h_n(x) - h_n(x_0)| < \epsilon \quad \text{for all } n$$

(b) Pointwise boundedness: for each x , the sequence $h_n(x)$ is bounded.

Then, there exist a subsequence h_{n_j} that converges uniformly on compacts to a continuous function h .

Step 2 We prove the limit functions are optimal for (P, Q) and conclude the entire sequence converges by a uniqueness argument.

Proof of Step 1: By Lemma A.2, there exists a (random) σ^2 such that the measures $\{P_n\}$ are uniformly σ^2 -subgaussian. We choose (f_n, g_n) and (f, g) as in Proposition A.1.

By Proposition A.1, (f_n, g_n) are pointwise bounded by a quantity independent of n . Likewise, Proposition 1 implies that the derivatives of f_n and g_n are also pointwise bounded, which implies local equicontinuity.

We conclude for a certain subsequence n_j , (f_{n_j}, g_{n_j}) converges to some (f_∞, g_∞) .

Proof of Step 2: It is easy to verify (by Jensen's inequality and dominated convergence) that Proposition 11 in Feydy et al. (2019), holds in arbitrary domains (not necessarily bounded), and we can assume (f, g) are unique $(P \otimes Q)$ -a.s. once we fix $E_P f(X) = E_Q g(Y)$. Notice that if $f_\infty = f, g_\infty = g$, P -a.s. and Q -a.s. we can conclude: on each compact we apply the above argument starting with any arbitrary subsequence n_k and find a subsequence such that $f_{n_{k_j}} \rightarrow f, g_{n_{k_j}} \rightarrow g$; therefore $f = \lim f_n(x)$ and $g(y) = \lim g_n$, uniformly in compacts.

It therefore suffices to show that that i) (f_∞, g_∞) satisfy the dual optimality conditions and that f_∞ (respectively g_∞) is P (respectively Q) integrable, with $E_P f_\infty(X) = E_Q g_\infty(Y)$. Let's prove i. Passing to a subsequence, we assume $f_n \rightarrow f$ and $g_n \rightarrow g$ uniformly on compact sets. We have

$$\begin{aligned} e^{-f_\infty(x)} &= \lim_{n \rightarrow \infty} \int e^{g_n(y) - \frac{1}{2}\|x-y\|^2} dQ_n(y) \\ e^{-g_\infty(y)} &= \lim_{n \rightarrow \infty} \int e^{f_n(x) - \frac{1}{2}\|x-y\|^2} dP_n(x). \end{aligned}$$

It suffices to show that the order of the limit and integral on the right side can be swapped. For a fixed x we observe that Proposition A.1 implies that the integrand is dominated by a uniformly integrable function. Therefore for an arbitrary $\varepsilon > 0$ there exists a compact set K such that

$$\begin{aligned} \int_{K^c} e^{g_\infty(y) - \frac{1}{2}\|x-y\|^2} dQ(y) &\leq \varepsilon \\ \int_{K^c} e^{g_n(y) - \frac{1}{2}\|x-y\|^2} dQ_n(y) &\leq \varepsilon \quad \forall n \geq 0. \end{aligned}$$

Write $v_n(y) = e^{g_n(y) - \frac{1}{2}\|x-y\|^2}$ and $v_\infty = e^{g_\infty(y) - \frac{1}{2}\|x-y\|^2}$. Since g_n converges uniformly in compacts so does v_n ; in particular, there exists n_0 such that if $n \geq n_0$,

$$|v_n(y) - v_\infty(y)| \leq \varepsilon, \forall y \in K. \quad (1)$$

Also, since v_∞ is Q -integrable, by the strong law of large numbers, almost surely there exists an n_1 such that if $n \geq n_1$,

$$\left| \int v_\infty(y) dQ_n(y) - \int v_\infty(y) dQ(y) \right| \leq \varepsilon, \quad (2)$$

We obtain that for n sufficiently large,

$$\left| \int v_n(y) dQ_n(y) - \int v_\infty(y) dQ(y) \right| \leq 4\varepsilon.$$

Since ε was arbitrary, we obtain

$$e^{-f_\infty(x)} = \int v_\infty(y) dQ(y) = \int e^{g_\infty(y) - \frac{1}{2}\|x-y\|^2} dQ(y).$$

Repeating the proof for g_∞ , we obtain that (f_∞, g_∞) satisfy the dual optimality conditions.

Clearly (f_∞, g_∞) are integrable by dominated convergence, and an argument analogous to the one used to show dual optimality establishes that $E_P f_\infty(X) = E_Q g_\infty(Y)$. The claim is therefore proved. \square

Lemma A.2. Suppose P is a σ^2 -subgaussian measure. Then, there exists a (random) $\sigma_u < \infty$ such that $\{P_n\}, P$ are uniformly σ_u^2 -subgaussian P almost surely.

Proof. By definition, there exists $\sigma > 0$ such that $E_P \left(e^{\frac{\|X\|^2}{2\sigma^2 d}} \right) \leq 2$. By the strong law of large numbers we have that P almost surely

$$\lim_{n \rightarrow \infty} E_{P_n} \left(e^{\frac{\|X\|^2}{2\sigma^2 d}} \right) = E_P \left(e^{\frac{\|X\|^2}{2\sigma^2 d}} \right) \leq 2.$$

In particular, this implies the sequence $E_{P_n} \left(e^{\frac{\|X\|^2}{2\sigma^2 d}} \right)$ is bounded by a random positive number. By the equivalence of definitions of subgaussianity, this implies that P_n are uniformly subgaussian, with a new parameter that we call σ_u^2 . \square

Proposition A.3. Assume P and Q are subgaussian. Let (f, g) be the corresponding optimal dual potentials constructed in Proposition A.1, and define

$$R_n = S(P_n, Q) - \int f(x) dP_n(x).$$

Then,

$$\lim_{n \rightarrow \infty} n \text{Var}(R_n) = 0.$$

Our proof relies on the tensorization property for the variance (Efron and Stein, 1981; Boucheron et al., 2013; van Handel, 2014), also known as Efron-Stein inequality: Let X_1, \dots, X_n be i.i.d r.v's with distribution P and X'_1, \dots, X'_n be independent copies of X_1, \dots, X_n . Also, let w be an arbitrary measurable function of the sample that is symmetric on its coordinates, and define $Z = w(X_1, \dots, X_n)$ and $Z' = w(X'_1, X_2, \dots, X_n)$. Then,

$$\text{Var}(Z) \leq \frac{n}{2} E(Z - Z')_+^2. \quad (3)$$

Proof of Proposition A.3. Denote by P'_n the empirical distribution of X'_1, X_2, \dots, X_n , and let

$$R'_n = S(P'_n, Q) - \int f(x) dP'_n(x).$$

by Efron-Stein, it suffices to show $\lim_{n \rightarrow \infty} n^2 E(R_n - R'_n)_+^2 = 0$. We divide the proof in the verification of two statements. First, we show $\lim_{n \rightarrow \infty} n(R_n - R'_n)_+ = 0$. We will then show that $n^2(R_n - R'_n)_+^2$ is uniformly integrable.

Call (f_n, g_n) the optimal potentials associated to (P_n, Q) . Since P_n is subgaussian by Lemma A.2, Proposition A.1 implies that we can assume that (f_n, g_n) satisfy the dual optimality conditions for all $x, y \in \mathbb{R}^d$. Therefore

$$\begin{aligned} S(P_n, Q) &= \int f_n(x) dP_n(x) + \int g_n(y) dQ(y), \\ S(P'_n, Q) &\geq \int f_n(x) dP'_n(x) + \int g_n(y) dQ(y) - \iint e^{f_n(x) + g_n(y) - \frac{1}{2}\|x-y\|^2} dP'_n(x) dQ(y) + 1 \\ &= \int f_n(x) dP'_n(x) + \int g_n(y) dQ(y). \end{aligned}$$

Therefore,

$$n(R_n - R'_n)_+ \leq (f_n(X_1) - f(X_1)) - (f_n(X'_1) - f(X'_1)).$$

By Proposition A.2, (f_n, g_n) converges pointwise to (f, g) almost surely, so $\lim_{n \rightarrow \infty} n(R_n - R'_n)_+ = 0$ almost surely.

To show uniform integrability, we note that $n(R_n - R'_n) = n(S(P_n, Q) - S(P'_n, Q)) - (f(X_1) - f(X'_1))$ and by Proposition A.1 and the subgaussianity of $P, f(X_1), f(X'_1)$ have finite second moments. It therefore suffices to show that $n^2(S(P_n, Q) - S(P'_n, Q))_+^2$ is uniformly integrable.

Let π' be the underlying optimal entropic coupling between P'_n and Q that we disintegrate in terms of Q and the (random) kernel $\{P'(\cdot|y)\}_y$ of conditional distributions over the sample P'_n given y , i.e.

$$d\pi'(x, y) = dQ(y) \left(P'(x|y)\delta_{X'_1}(x) + \sum_{i=2}^n P'(x|y)\delta_{X_i}(x) \right).$$

We now slightly modify π' to make it have P_n as first marginal; specifically, we define

$$d\bar{\pi}(x, y) = dQ(y) \left(\sum_{i=1}^n \bar{P}(x|y)\delta_{X_i}(x) \right), \text{ with } \bar{P}(x|y) = \begin{cases} P'(X'_1|y) & x = X_1 \\ P'(X_i|y) & x = X_i, i \neq 1 \end{cases}.$$

By the definitions of $S(P_n, Q)$ and $S(P'_n, Q)$, it is easily verified that

$$S(P_n, Q) \leq \sum_{i=1}^n \int \frac{\|X_i - y\|^2}{2} \bar{P}(X_i|y) dQ(y) + I(\bar{\pi}),$$

and that

$$S(P'_n, Q) = \int \frac{\|X'_1 - y\|^2}{2} P'(X'_1|y) dQ(y) + \sum_{i=2}^n \int \frac{\|X_i - y\|^2}{2} P'(X_i|y) dQ(y) + I(\pi'),$$

where $I(\cdot)$ denotes mutual information. Therefore,

$$S(P_n, Q) - S(P'_n, Q) \leq I(\bar{\pi}) - I(\pi') + \int \frac{\|X_1 - y\|^2 - \|X'_1 - y\|^2}{2} P'(X'_1|y) dQ(y). \quad (4)$$

Observe that $I(\bar{\pi}) = I(\pi')$ since $I(\pi')$ doesn't depend on the sample values, but only in the way the conditionals $P'(\cdot|y)$ split over the sample, which by construction is the same for both $\bar{\pi}$ and π' . Therefore, we only need to bound the (expected squared) integral in (4), and we proceed as in Del Barrio and Loubes (2019). Specifically, we have

$$\begin{aligned} S(P_n, Q) - S(P'_n, Q) &\leq \int \frac{\|X_1 - y\|^2 - \|X'_1 - y\|^2}{2} P'(X'_1|y) dQ(y) \\ &\leq \frac{1}{2} \|X_1 - X'_1\| \left(\frac{\|X_1\| + \|X'_1\|}{n} + 2 \int \|y\| P'(X'_1|y) dQ(y) \right), \end{aligned} \quad (5)$$

from which it follows that

$$n^2(S(P_n, Q) - S(P'_n, Q))^2_+ \leq (\|X_1 - X'_1\|^2 \|X_1\|^2) + n^2 \|X_1 - X'_1\|^2 \left(\int \|y\| P'(X'_1|y) dQ(y) \right)^2. \quad (6)$$

The first term is clearly uniformly integrable since P has moments of all orders, so we focus on the second term.

By Cauchy-Schwartz,

$$\begin{aligned} E \left(\|X_1 - X'_1\|^4 \left(\int \|y\| P'(X'_1|y) dQ(y) \right)^4 \right)^2 &\leq E \left(\|X_1 - X'_1\|^8 \right) \times \\ &\quad E \left(\left(\int \|y\| P'(X'_1|y) dQ(y) \right)^8 \right). \end{aligned}$$

And now, by Hölder's inequality,

$$\begin{aligned} \left(\int \|y\| P'(X'_1|y) dQ(y) \right)^8 &\leq \left(\int P'(X'_1|y) dQ(y) \right)^7 \left(\int \|y\|^8 P'(X'_1|y) dQ(y) \right) \\ &= \frac{1}{n^7} \left(\int \|y\|^8 P'(X'_1|y) dQ(y) \right). \end{aligned}$$

Also, notice that the r.v.'s $\int \|y\|^8 P'(X'_i|y)dQ(y)$ are equally distributed, and therefore

$$E \left(\int \|y\|^8 P'(X'_1|y)dQ(y) \right) = \frac{1}{n} E \left(\sum_{i=1}^n \int \|y\|^8 P'(X'_i|y)dQ(y) \right) = \frac{1}{n} E \left(\int \|y\|^8 dQ(y) \right).$$

We obtain

$$E \left(\left(\int \|y\| P'(X'_1|y)dQ(y) \right)^8 \right) \leq \frac{1}{n^8} \int \|y\|^8 dQ(y). \quad (7)$$

Together, (7) and (7) imply that the quantity $n^2 \|X_1 - X'_1\|^2 \left(\int \|y\| P'(X'_1|y)dQ(y) \right)^2$ has uniformly bounded second moments, and is therefore uniformly integrable. Therefore $n^2 (S(P_n, Q) - S(P'_n, Q))^2_+$ is uniformly integrable as well, and combining this with the almost sure convergence implies the claim. \square

Lemma A.3. *Let μ_β be defined as in (8). Then*

$$|\mu_\beta| \leq C_{|\beta|,d} \begin{cases} \sigma^{|\beta|}(\sigma + \sigma^2)^{|\beta|} & \|x\| \leq \sqrt{d}\sigma \\ \sigma^{|\beta|}(\sqrt{\sigma}\|x\| + \sigma\|x\|)^{|\beta|} & \|x\| > \sqrt{d}\sigma. \end{cases}$$

Proof. To bound μ_β , we split the integral in the numerator according to the norm of y . Let $A = \{y : \|y\| \leq \tau\}$, where τ is a threshold to be chosen. Then

$$\mu_\beta = \frac{\int \mathbf{1}_A y^\beta e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}{\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)} + \frac{\int \mathbf{1}_{\bar{A}} y^\beta e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}{\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}.$$

The first term is clearly bounded by τ^β . For the second, we apply Proposition A.1 to show

$$\left(\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y) \right)^{-1} = e^{-\frac{1}{2}\|x\|^2} e^{f(x)} \leq e^{d\sigma^2 + \sqrt{d}\sigma\|x\|}$$

and

$$e^{g(y) - \frac{1}{2}\|y\|^2} \leq e^{d\sigma^2 + \sqrt{d}\sigma\|y\|}.$$

We obtain

$$\begin{aligned} \frac{\int \mathbf{1}_{\bar{A}} y^\beta e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}{\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)} &\leq e^{2d\sigma^2 + \sqrt{d}\sigma\|x\|} \int \mathbf{1}_{\bar{A}} y^\beta e^{\sqrt{d}\sigma\|y\| + x \cdot y} dQ(y) \\ &\leq e^{2d\sigma^2 + \sqrt{d}\sigma\|x\|} \left(\int \mathbf{1}_{\bar{A}} y^{2\beta} dQ(y) \right)^{1/2} \left(\int e^{2(\sqrt{d}\sigma + \|x\|)\|y\|} dQ(y) \right)^{1/2} \end{aligned}$$

Since Q is subgaussian, Lemma A.1 and the definition of A imply

$$\left(\int \mathbf{1}_{\bar{A}} y^{2\beta} dQ(y) \right)^{1/2} \leq e^{-\frac{\tau^2}{8d\sigma^2}} \left(\int e^{\frac{\|y\|^2}{4d\sigma^2}} y^{2\beta} dQ(y) \right)^{1/2} \leq \sqrt{2} e^{-\frac{\tau^2}{8d\sigma^2}} (2|\beta|)!^{1/4} (\sqrt{2d}\sigma)^{|\beta|}.$$

Lemma A.1 also implies

$$\int e^{2(\sqrt{d}\sigma + \|x\|)\|y\|} dQ(y) \leq 2e^{2d\sigma^2(\|x\| + \sqrt{d}\sigma)^2}.$$

Therefore, if we choose $\tau^2 \geq C_{|\beta|,d}(\sigma^4 + \sigma^6)$ if $\|x\| \leq \sqrt{d}\sigma$ and $\tau^2 \geq C_{|\beta|,d}(\sigma^3\|x\| + \sigma^4\|x\|^2)$ if $\|x\| > \sqrt{d}\sigma$ for a sufficiently large constant $C_{|\beta|,d}$, then we will have

$$\frac{\int \mathbf{1}_{\bar{A}} y^\beta e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}{\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)} \leq C_{|\beta|,d} (\sqrt{d}\sigma)^{|\beta|}$$

Combining this with the bound on the first term yields the claim. \square

Lemma A.4. *Let $\tilde{\sigma}$ be defined as in the proof of Theorem 2. Then for any positive integer k ,*

$$E \tilde{\sigma}^{2k} \leq 2k^k \sigma^{2k}.$$

Proof. First, let P be an arbitrary probability distribution, and let $\alpha > 0$. We first show that if $t = E_P e^{\frac{\|X\|^2}{\alpha}}$ is finite, then P is $t\frac{\alpha}{2d}$ -subgaussian. To see this, set $\tau^2 = t\frac{\alpha}{2d}$. Then

$$E e^{\frac{\|X\|^2}{2d\tau^2}} \leq \left(E e^{\frac{\|X\|^2}{\alpha}} \right)^{\frac{\alpha}{2d\tau^2}} = t^{1/t} \leq e^{1/e} < 2,$$

where the first step uses Jensen's inequality and the fact that $t \geq 1$.

The above considerations imply that if Q is σ^2 subgaussian and we set

$$\tau^2 = \max\{E_{P_n} e^{\frac{\|X\|^2}{2kd\sigma^2}} k\sigma^2, E_{Q_n} e^{\frac{\|Y\|^2}{2kd\sigma^2}} k\sigma^2\},$$

then P_n, Q_n, P , and Q are all τ^2 subgaussian, which implies that $\tilde{\sigma}^2 \leq \tau^2$. Therefore, by Jensen's inequality,

$$\tilde{\sigma}^{2k} \leq E_{P_n} e^{\frac{\|X\|^2}{2d\sigma^2}} k^k \sigma^{2k} + E_{Q_n} e^{\frac{\|Y\|^2}{2d\sigma^2}} k^k \sigma^{2k},$$

and taking expectations with respect to P and Q yields

$$E\tilde{\sigma}^{2k} \leq E_P e^{\frac{\|X\|^2}{2d\sigma^2}} k^k \sigma^{2k} + E_Q e^{\frac{\|Y\|^2}{2d\sigma^2}} k^k \sigma^{2k} \leq 4k^k \sigma^{2k}.$$

□

B Omitted proofs

B.1 Proof of Proposition 1

We choose the potentials f and g as in Proposition A.1. That establishes the $k = 0$ case.

For convenience, write $\bar{f}(x) = f(x) - \frac{1}{2}\|x\|^2$. We seek to bound $|D^\alpha \bar{f}(x)|$.

Our calculation is similar to classical calculations which relate the cumulants of a distribution to its moments (see McCullagh, 1987, Section 2.3). Given a multi-index β , write

$$\mu_\beta = \frac{\int y^\beta e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}{\int e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y} dQ(y)}. \quad (8)$$

We use the convention that $y^\beta = \prod_{i=1}^d y_i^{\beta_i}$. The notation μ_β is chosen to remind the reader that these quantities are moments of y under the tilted measure whose density with respect to Q is proportional to $e^{g(y) - \frac{1}{2}\|y\|^2 + x \cdot y}$.

By the multivariate Faà di Bruno formula (see, e.g. Constantine and Savits, 1996),

$$D^\alpha \bar{f}(x) = -D^\alpha \log(e^{-\bar{f}(x)}) = \sum_{\substack{\beta_1, \dots, \beta_k \\ \beta_1 + \dots + \beta_k = \alpha}} \lambda_{\alpha, \beta_1, \dots, \beta_k} \prod_{j=1}^k \mu_{\beta_j}, \quad (9)$$

where the coefficients $\lambda_{\alpha, \beta_1, \dots, \beta_k}$ are combinatorial quantities related to partitions of $[k]$ whose precise value is unimportant.

Applying Lemma A.3 yields the claim.

B.2 Proof of Proposition 3

We use the symbol C , decorated with subscripts, to indicate constants whose value may change from line to line. We apply van der Vaart and Wellner (1996, Corollary 2.7.4). Denote by L the quantity $\frac{1}{n} \sum_{i=1}^n e^{\|x_i\|^2 / 2d\sigma^2}$. The subgaussianity of P implies that $EL \leq 2$. We partition \mathbb{R}^d into sets B_j defined by $B_0 = [-\sigma, \sigma]^d$ and $B_j = [-2^j\sigma, 2^j\sigma] \setminus [-2^{j-1}\sigma, 2^{j-1}\sigma]$. Note that for each j , the Lebesgue measure of $\{x : d(x, B_j) \leq 1\}$ is bounded by $C_d(1 + \sigma^d 2^{dj})$. Moreover, by Markov's inequality, the mass that P_n assigns to each B_j is at most $Le^{-2^{2j-3}}$. Finally, by definition of the class \mathcal{F}^s , the functions in \mathcal{F}^s have $\mathcal{C}^s(B_0)$ norm at most $C_{s,d}(1 + \sigma^s)$, and on B_j for $j \geq 1$ have $\mathcal{C}^s(B_j)$ norm at most $C_{s,d}2^{js}(1 + \sigma^s)$, where $\mathcal{C}^s(\Omega)$ represents the Hölder space on Ω of smoothness s .

Applying van der Vaart and Wellner (1996, Corollary 2.7.4) with $V = d/s$ and $r = 2$ yields

$$\begin{aligned} \log N(\varepsilon, \mathcal{F}^s, L_2(P_n)) &\leq C_d \varepsilon^{-d/s} L^{d/2s} \left(\sum_{j \geq 0} (1 + \sigma^d 2^{dj})^{\frac{2s}{d+2s}} 2^{\frac{2dj s}{d+2s}} (1 + \sigma^s)^{\frac{2d}{d+2s}} e^{-\frac{d 2^{2j-3}}{d+2s}} \right)^{\frac{d+2s}{2s}} \\ &\leq C_d \varepsilon^{-d/s} L^{d/2s} (1 + \sigma^{2d}) \left(\sum_{j \geq 0} 2^{\frac{4dj s}{d+2s}} e^{-\frac{d 2^{2j-3}}{d+2s}} \right)^{\frac{d+2s}{2s}} \\ &\leq C_d \varepsilon^{-d/s} L^{d/2s} (1 + \sigma^{2d}), \end{aligned}$$

where the final step follows because the series is summable with value independent of σ and L .

To show the second claim, we note that $E_{P_n} \|X\|^4 \leq C_d L \sigma^4$ by the same argument used to bound the moments of P in Lemma A.1. The definition of the class \mathcal{F}^s implies

$$\max_{f \in \mathcal{F}^s} \|f\|_{L_2(P_n)}^2 = \max_{f \in \mathcal{F}^s} E_{P_n} |f(X)|^2 \leq C_d E_{P_n} (1 + \|X\|^4) \leq C_d (1 + L \sigma^4).$$

□

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