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A Proofs for 2-Layer LNNs

We will start by proving the following proposition.

Proposition A.1. *Assume that all weight values and the data points are independent random variable. And Further assume that their k -th order moments are bounded when $k \leq 4$. Within a weight matrix, all entries have the same moment value. All data points also have the same moment value. Then for any pair $i, j \in \{1, \dots, n\}$:*

$$E(\langle \nabla f_i, \nabla f_j \rangle) = \begin{cases} \Theta(K^2 d^2) & \text{if } i = j \\ \Theta(Kd(K+d)) & \text{if } i \neq j \end{cases}$$

To prove the above proposition, we first define some notations. Let $M_{i,j}$ be the j -th moment of the entries of W_i , and $W_{x,j}$ be the j -th moment of the entries of data point x_i .

Let us prove the two cases separately. We first consider the case where $i = j$. We can write the inner product as

$$E(\|\nabla f_i\|^2) = \sum_{p=1}^K \sum_{q=1}^d E(\|\frac{\partial f_i}{\partial W_{1,p,q}}\|^2) + \sum_{q=1}^K E(\|\frac{\partial f_i}{\partial W_{2,1,q}}\|^2)$$

We will show in Lemma 2 that the first expectation is $\Theta(Kd)$, and in Lemma 3 that the second expectation is $\Theta(Kd^2)$. Plugging these results into the above equation gives the desired result.

Lemma 2. $E(\|\frac{\partial f_i}{\partial W_{1,p,q}}\|^2) = \Theta(Kd)$.

Proof. Note that $\frac{\partial f_i}{\partial W_{1,p,q}} = (\hat{y}_i - y_i)W_{2,1,p}x_{i,q} = (W_2W_1 - W_2^*W_1^*)x_iW_{2,1,p}x_{i,q}$. We have

$$\begin{aligned} E(\|\frac{\partial f_i}{\partial W_{1,p,q}}\|^2) &= E((W_2W_1 - W_2^*W_1^*)x_iW_{2,1,p}x_{i,q})^2 \\ &= E((W_2W_1)x_iW_{2,1,p}x_{i,q})^2 + E((W_2^*W_1^*)x_iW_{2,1,p}x_{i,q})^2 \\ &= E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_i)^2 W_{2,1,p}^2 x_{i,q}^2\right) + E\left(\sum_{s=1}^K W_{2,1,s}^{*2} (W_{1,s,:}^*x_i)^2 W_{2,1,p}^2 x_{i,q}^2\right) \\ &= E\left(\sum_{s=1}^K W_{2,1,s}^2 \left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}^2\right) W_{2,1,p}^2 x_{i,q}^2\right) + E\left(\sum_{s=1}^K W_{2,1,s}^{*2} \left(\sum_{t=1}^d W_{1,s,t}^{*2} x_{i,t}^2\right) W_{2,1,p}^2 x_{i,q}^2\right) \\ &= M_{1,2}((K-1)M_{2,2}^2 + M_{2,4})((d-1)M_{x,2}^2 + M_{x,4}) + M_{1^*,2}KM_{2^*,2}((d-1)M_{x,2}^2 + M_{x,4}) \end{aligned}$$

This concludes the proof. \square

Lemma 3. $E(\|\frac{\partial f_i}{\partial W_{2,1,q}}\|^2) = \Theta(Kd^2)$.

Proof. Note that $\frac{\partial f_i}{\partial W_{2,1,q}} = (\hat{y}_i - y_i)W_{1,q,:}x_i = (W_2W_1 - W_2^*W_1^*)x_iW_{1,q,:}x_i$. We have

$$\begin{aligned} E(\|\frac{\partial f_i}{\partial W_{2,1,q}}\|^2) &= E((W_2W_1 - W_2^*W_1^*)x_iW_{1,q,:}x_i)^2 \\ &= E((W_2W_1)x_iW_{1,q,:}x_i)^2 + E((W_2^*W_1^*)x_iW_{1,q,:}x_i)^2 \\ &= E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_i)^2 (W_{1,q,:}x_i)^2\right) + E\left(\sum_{s=1}^K W_{2,1,s}^{*2} (W_{1,s,:}^*x_i)^2 (W_{1,q,:}x_i)^2\right). \end{aligned}$$

For the first term, we have

$$\begin{aligned} E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_i)^2 (W_{1,q,:}x_i)^2\right) &= E\left(\sum_{s=1}^K W_{2,1,s}^2 \left(\sum_{t=1}^d W_{1,s,t}x_{i,t}\right)^2 \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}\right)^2\right) \\ &= \sum_{s=1}^K M_{2,2}E\left(\left(\sum_{t=1}^d W_{1,s,t}x_{i,t}\right)^2 \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}\right)^2\right) \end{aligned}$$

We now distinguish two cases. If $s \neq q$,

$$\begin{aligned} E\left(\left(\sum_{t=1}^d W_{1,s,t}x_{i,t}\right)^2 \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}\right)^2\right) &= E\left(\left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}^2\right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u}^2\right)\right) \\ &= dM_{1,2}^2M_{x,4} + d(d-1)M_{1,2}^2M_{x,2}^2 = dM_{1,2}^2(M_{x,4} + (d-1)M_{x,2}^2) = \Theta(d^2) \end{aligned}$$

If $s = q$,

$$\begin{aligned}
& E \left(\left(\sum_{t=1}^d W_{1,s,t} x_{i,t} \right)^2 \left(\sum_{u=1}^d W_{1,q,u} x_{i,u} \right)^2 \right) \\
&= E \left(\left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}^2 \right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u}^2 \right) \right) \\
&+ E \left(\left(\sum_{t=1, t \neq v}^d W_{1,s,t} x_{i,t} W_{1,s,v} x_{i,v} \right) \left(\sum_{u=1, u \neq w}^d W_{1,q,u} x_{i,u} W_{1,q,w} x_{i,w} \right) \right) \\
&= E \left(\left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}^2 \right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u}^2 \right) \right) + E \left(\left(2 \sum_{t=1, t \neq v}^d W_{1,s,t}^2 x_{i,t}^2 W_{1,s,v}^2 x_{i,v}^2 \right) \right) \\
&= dM_{1,2}^2 M_{x,4} + 3d(d-1)M_{1,2}^2 M_{x,2}^2 = \Theta(d^2)
\end{aligned}$$

Combining the two cases, we have

$$E \left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:} x_i)^2 (W_{1,q,:} x_i)^2 \right) = \Theta(d^2)$$

For the second term, we have

$$\begin{aligned}
& E \left(\sum_{s=1}^K W_{2^*,1,s}^2 (W_{1^*,s,:} x_i)^2 (W_{1,q,:} x_i)^2 \right) \\
&= E \left(\sum_{s=1}^K W_{2^*,1,s}^2 \left(\sum_{t=1}^d W_{1^*,s,t} x_{i,t} \right)^2 \left(\sum_{u=1}^d W_{1,q,u} x_{i,u} \right)^2 \right) \\
&= \sum_{s=1}^K M_{2^*,2} E \left(\left(\sum_{t=1}^d W_{1^*,s,t} x_{i,t} \right)^2 \left(\sum_{u=1}^d W_{1,q,u} x_{i,u} \right)^2 \right) \\
&= \sum_{s=1}^K M_{2^*,2} E \left(\left(\sum_{t=1}^d W_{1^*,s,t}^2 x_{i,t}^2 \right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u}^2 \right) \right) \\
&= \sum_{s=1}^K M_{2^*,2} (dM_{1^*,2} M_{1,2} M_{x,4} + d(d-1)M_{1^*,2} M_{1,2} M_{x,2}^2) \\
&= \Theta(Kd^2).
\end{aligned}$$

Combining the first term and the second term, we obtain the desired result. \square

We next consider the case $i \neq j$. In this case, we can write the inner product as

$$E(\langle \nabla f_i, \nabla f_j \rangle) = \sum_{p=1}^K \sum_{q=1}^d E(\langle \frac{\partial f_i}{\partial W_{1,p,q}}, \frac{\partial f_j}{\partial W_{1,p,q}} \rangle) + \sum_{q=1}^K E(\langle \frac{\partial f_i}{\partial W_{2,1,q}}, \frac{\partial f_j}{\partial W_{2,1,q}} \rangle)$$

As before, we will show in Lemma 4 that the first expectation is $\Theta(K)$, and in Lemma 5 that the second expectation is $\Theta(d(K+d))$. Plugging these results into the above equation gives the desired result.

Lemma 4. If $i \neq j$, $E(\langle \frac{\partial f_i}{\partial W_{1,p,q}}, \frac{\partial f_j}{\partial W_{1,p,q}} \rangle) = \Theta(K)$.

Proof. Note that $\frac{\partial f_i}{\partial W_{1,p,q}} = (\hat{y}_i - y_i)W_{2,1,p}x_{i,q} = (W_2W_1 - W_2^*W_1^*)x_iW_{2,1,p}x_{i,q}$. We have

$$\begin{aligned}
& E\left(\left\langle \frac{\partial f_i}{\partial W_{1,p,q}}, \frac{\partial f_j}{\partial W_{1,p,q}} \right\rangle\right) \\
&= E\left(\left\langle (W_2W_1 - W_2^*W_1^*)x_iW_{2,1,p}x_{i,q}(W_2W_1 - W_2^*W_1^*)x_jW_{2,1,p}x_{j,q} \right\rangle\right) \\
&= E\left(W_2W_1x_iW_{2,1,p}x_{i,q}W_2W_1x_jW_{2,1,p}x_{j,q}\right) + E\left(W_2^*W_1^*x_iW_{2,1,p}x_{i,q}W_2^*W_1^*x_jW_{2,1,p}x_{j,q}\right) \\
&= E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_iW_{1,s,:}x_j) W_{2,1,p}^2 x_{i,q}x_{j,q}\right) \\
&+ E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}^*x_iW_{1,s,:}^*x_j) W_{2,1,p}^2 x_{i,q}x_{j,q}\right) \\
&= E\left(\sum_{s=1}^K W_{2,1,s}^2 \left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}x_{j,t}\right) W_{2,1,p}^2 x_{i,q}x_{j,q}\right) \\
&+ E\left(\sum_{s=1}^K W_{2,1,s}^{*2} \left(\sum_{t=1}^d W_{1,s,t}^{*2} x_{i,t}x_{j,t}\right) W_{2,1,p}^2 x_{i,q}x_{j,q}\right) \\
&= M_{1,2}((K-1)M_{2,2}^2 + M_{2,4})M_{x,2}^2 + M_{1^*,2}KM_{2,2}M_{2^*,2}M_{x,2}^2.
\end{aligned}$$

This concludes the proof. \square

Lemma 5. If $i \neq j$, $E\left(\left\langle \frac{\partial f_i}{\partial W_{2,1,q}}, \frac{\partial f_j}{\partial W_{2,1,q}} \right\rangle\right) = \Theta(d(K+d))$.

Proof. Note that $\frac{\partial f_i}{\partial W_{2,1,q}} = (\hat{y}_i - y_i)W_{1,q,:}x_i = (W_2W_1 - W_2^*W_1^*)x_iW_{1,q,:}x_i$. We have

$$\begin{aligned}
& E\left(\left\langle \frac{\partial f_i}{\partial W_{2,1,q}}, \frac{\partial f_j}{\partial W_{2,1,q}} \right\rangle\right) \\
&= E\left(\left\langle (W_2W_1 - W_2^*W_1^*)x_iW_{1,q,:}x_i(W_2W_1 - W_2^*W_1^*)x_jW_{1,q,:}x_j \right\rangle\right) \\
&= E\left(W_2W_1x_iW_{1,q,:}x_iW_2W_1x_jW_{1,q,:}x_j\right) + E\left(W_2^*W_1^*x_iW_{1,q,:}x_iW_2^*W_1^*x_jW_{1,q,:}x_j\right) \\
&= E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_iW_{1,s,:}x_j)(W_{1,q,:}x_iW_{1,q,:}x_j)\right) \\
&+ E\left(\sum_{s=1}^K W_{2,1,s}^{*2} (W_{1,s,:}^*x_iW_{1,s,:}^*x_j)(W_{1,q,:}x_iW_{1,q,:}x_j)\right)
\end{aligned}$$

For the first term, we have

$$\begin{aligned}
& E\left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:}x_iW_{1,s,:}x_j)(W_{1,q,:}x_iW_{1,q,:}x_j)\right) \\
&= E\left(\sum_{s=1}^K W_{2,1,s}^2 \left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right)\right) \\
&= \sum_{s=1}^K M_{2,2}E\left(\left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right)\right).
\end{aligned}$$

We now distinguish two cases. If $s \neq q$,

$$\begin{aligned}
& E\left(\left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{t=1}^d W_{1,s,t}x_{i,t}x_{j,t}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right) \left(\sum_{u=1}^d W_{1,q,u}x_{i,u}x_{j,u}\right)\right) \\
&= E\left(\left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t}x_{j,t}\right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u}x_{j,u}\right)\right) \\
&= E\left(\sum_{t=1}^d W_{1,s,t}^2 W_{1,q,t}^2 x_{i,t}^2 x_{j,t}^2\right) \\
&= dM_{1,2}^2 M_{x,2}^2.
\end{aligned}$$

If $s = q$,

$$\begin{aligned}
& E \left(\left(\sum_{t=1}^d W_{1,s,t} x_{i,t} \right) \left(\sum_{t=1}^d W_{1,s,t} x_{j,t} \right) \left(\sum_{u=1}^d W_{1,q,u} x_{i,u} \right) \left(\sum_{u=1}^d W_{1,q,u} x_{j,u} \right) \right) \\
&= E \left(\left(\sum_{t=1}^d W_{1,s,t}^2 x_{i,t} x_{j,t} \right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u} x_{j,u} \right) \right) \\
&+ E \left(\left(\sum_{t=1, t \neq v}^d W_{1,s,t} x_{i,t} W_{1,s,v} x_{j,v} \right) \left(\sum_{u=1, u \neq w}^d W_{1,q,u} x_{i,u} W_{1,q,w} x_{j,w} \right) \right) \\
&= E \left(\sum_{t=1}^d W_{1,s,t}^4 x_{i,t}^2 x_{j,t}^2 \right) + E \left(\sum_{t=1, t \neq v}^d W_{1,s,t}^2 x_{i,t}^2 W_{1,s,v}^2 x_{j,v}^2 \right) \\
&= d M_{1,4} M_{x,2}^2 + d(d-1) M_{1,2}^2 M_{x,2}^2.
\end{aligned}$$

Combining the two cases,

$$\begin{aligned}
E \left(\sum_{s=1}^K W_{2,1,s}^2 (W_{1,s,:} x_i)^2 (W_{1,q,:} x_j)^2 \right) &= M_{2,2} ((K-1) d M_{1,2}^2 M_{x,2}^2 + d M_{1,4} M_{x,2}^2 + d(d-1) M_{1,2}^2 M_{x,2}^2) \\
&= \Theta(Kd + d^2)
\end{aligned}$$

For the second term, we have

$$\begin{aligned}
& E \left(\sum_{s=1}^K W_{2,1,s}^{*2} (W_{1,s,:}^* x_i W_{1,s,:}^* x_j) (W_{1,q,:} x_i W_{1,q,:} x_j) \right) \\
&= E \left(\sum_{s=1}^K W_{2,1,s}^{*2} \left(\sum_{t=1}^d W_{1^*,s,t} x_{i,t} \right) \left(\sum_{t=1}^d W_{1^*,s,t} x_{j,t} \right) \left(\sum_{u=1}^d W_{1,q,u} x_{i,u} \right) \left(\sum_{u=1}^d W_{1,q,u} x_{j,u} \right) \right) \\
&= E \left(\sum_{s=1}^K W_{2,1,s}^{*2} \left(\sum_{t=1}^d W_{1^*,s,t}^2 x_{i,t} x_{j,t} \right) \left(\sum_{u=1}^d W_{1,q,u}^2 x_{i,u} x_{j,u} \right) \right) \\
&= E \left(\sum_{s=1}^K W_{2,1,s}^{*2} \left(\sum_{t=1}^d W_{1^*,s,t}^2 W_{1,q,u}^2 x_{i,t}^2 x_{j,t}^2 \right) \right) \\
&= K d M_{2^*,2} M_{1^*,2} M_{1,2} M_{x,2}^2 = \Theta(Kd).
\end{aligned}$$

Combining the first term and the second term, we obtain the desired result. \square

By applying Proposition A.1 and linearity of expectation, we obtain the following result:

Theorem 4. *We have*

$$E \left(\sum_{i=1}^n \|\nabla f_i\|^2 \right) = \Theta(n K^2 d^2) \quad (\text{A.1})$$

$$E \left(\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle \right) = \Theta(n^2 K^2 d + n^2 K d^2) \quad (\text{A.2})$$

The above theorem computes the expectation of each term of the ratio. In order to obtain a result on the expectation of the ratio, we also need to show that the value of each term will be concentrated around the expectation with high probability. To prove such a result, we first compute the variance.

Theorem 5. *We have*

$$\text{Var} \left(\sum_{i=1}^n \|\nabla f_i\|^2 \right) = \Theta(n^2 K^4 d^3) \quad (\text{A.3})$$

$$\text{Var} \left(\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle \right) = \Theta(n^4 K^3 d^2) \quad (\text{A.4})$$

Theorem 1. *Consider a 2 LNN. Let the weights $W_{l,p,q}, W_{l,p,q}^*$ for $l \in \{1, 2\}$ and \mathbf{x}_i be independently drawn random variables, such that their k -th order moments for $k \leq 4$ are bounded in a positive interval. Then, with arbitrary constant probability, the following holds:*

$$B_S(\mathbf{w}) \geq \frac{\Theta(n K d)}{\Theta(K n + d n + K d)}$$

Proof. By Chebyshev's Inequality, we have

$$\Pr \left(\left| \sum_{i=1}^n \|\nabla f_i\|^2 - E \left(\sum_{i=1}^n \|\nabla f_i\|^2 \right) \right| \geq \epsilon \right) \leq \frac{\text{Var}(\sum_{i=1}^n \|\nabla f_i\|^2)}{\epsilon^2}$$

Using the above two theorems, and choosing parameter $\epsilon = \Theta(nK^2 d^{3/2} \delta^{-1/2})$, we have that with probability $1 - \delta$,

$$\Theta(nK^2 d^2) - \Theta(nK^2 d^{3/2} \delta^{-1/2}) \leq \sum_{i=1}^n \|\nabla f_i\|^2 \leq \Theta(nK^2 d^2) + \Theta(nK^2 d^{3/2} \delta^{-1/2})$$

We can similarly use Chebyshev's inequality to obtain that with probability $1 - \delta$,

$$\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle \leq \Theta(n^2 K d (K + d)) + \Theta(n^2 K^{3/2} d \delta^{-1/2})$$

By applying the union bound, we can now bound the ratio as desired:

$$\begin{aligned} \frac{n \sum_{i=1}^n \|\nabla f_i\|^2}{\|\sum_{i=1}^n \nabla f_i\|^2} &= \frac{n \sum_{i=1}^n \|\nabla f_i\|^2}{\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle + \sum_{i=1}^n \|\nabla f_i\|^2} \\ &\geq \frac{\Theta(n^2 K^2 d^2) - \Theta(\frac{n^2 K^2 d^{3/2}}{\sqrt{\delta}})}{\Theta(n^2 K d (K + d)) + \Theta(\frac{n^2 K^{3/2} d}{\sqrt{\delta}}) + \Theta(nK^2 d^2) - \Theta(\frac{nK^2 d^{3/2}}{\sqrt{\delta}})} \\ &= \frac{\Theta(nKd)}{\Theta(Kn + dn + Kd)} \end{aligned}$$

Here we assumed that δ is chosen to be some arbitrarily small constant. d and n is sufficiently large such that $d > \frac{1}{\delta}$ and $\Theta(n^2)$ dominates $\Theta(n)$. \square

B Proofs for 2-Layer Nonlinear Neural Networks

In this section we present the detailed proof of Theorem 2, which is restated as below (A few constants are stated explicitly for ease of proof.).

Theorem 2. Consider a 2-layer NN with a monotone activation function σ such that for every x we have: $-\sigma(x) = \sigma(-x)$, and both $|\sigma(x)|$ and $\sup_x \{x\sigma'(x)\}$ are bounded. Let the weights $W_{l,p,q}, W_{l,p,q}^*$ for $l \in \{1, 2\}$ and \mathbf{x}_i be i.i.d. random variables from $\mathcal{N}(0, 1)$. Then, with high probability, the following holds:

$$\frac{\mathbb{E}[n \sum_{i=1}^n \|\nabla f_i\|_2^2]}{\mathbb{E}[\|\sum_{i=1}^n \nabla f_i\|_2^2]} \geq \Omega\left(\frac{Kd^2}{Kd + K + d}\right).$$

where the expectation is over W_2, W_2^* .

Denote the bound of $|\sigma(x)|$ by c_{max} , and the bound of $\sup_x \{x\sigma'(x)\}$ by c_{sup} .

B.1 Notations and Models

We consider a 2-layer nonlinear neural network. Let W_1, W_2 be the coefficient matrix of the first and second layer, respectively. $W_{a,p,q}$ is the p, q element in matrix a . For ease of notations, let us further define

$$\begin{aligned} A_1 &= n \sum_{i=1}^n (\hat{y}_i - y_i)^2 \left(\sum_{p=1}^K \sum_{q=1}^d \left(\frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 \right), \\ A_2 &= n \sum_{i=1}^n (\hat{y}_i - y_i)^2 \left(\sum_{q=1}^d \left(\frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right), \\ B_1 &= \left(\sum_{p=1}^K \sum_{q=1}^d \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 \right), \\ B_2 &= \left(\sum_{q=1}^d \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right). \end{aligned}$$

B.2 Some Helper Lemmas

We first provide some lemmas.

Lemma 6. *Let X, Y, Z be three normal distribution. Let $\rho_{XY}, \rho_{YZ}, \rho_{XZ}$ be the correlation between those random variables. Let $f(\cdot)$ be a monotone, bounded, and differentiable function. More precisely, $f(x) \geq f(y)$ iff $x \geq y$, $|f(x)| \leq f_{\max}$. Further assume $\sup f'(x)x = c$. Then we have*

$$|E(f(X)f'(Y)Z)| \leq (1 - \rho_{XY}^2)^{-1} |\rho_{XZ} - \rho_{YZ}\rho_{XY}| \times f_{\max} f'_{\max} \sigma_x + (1 - \rho_{XY}^2)^{-1} |-\rho_{XZ}\rho_{XY} + \rho_{YZ}| \times f_{\max} c.$$

Proof. Given X and Y , random variable Z is a normal distributed random variable with mean $E(Z|X, Y) = (1 - \rho_{XY}^2)^{-1}(\rho_{XZ} - \rho_{YZ}\rho_{XY})X + (1 - \rho_{XY}^2)^{-1}(-\rho_{XZ}\rho_{XY} + \rho_{YZ})Y$. Thus it holds that

$$\begin{aligned} & |E(f(X)f'(Y)Z)| \\ &= |E(E(Z|X, Y)f(X)f'(Y))| \\ &= (1 - \rho_{XY}^2)^{-1} |E((\rho_{XZ} - \rho_{YZ}\rho_{XY})Xf(Xf'(Y)) + (-\rho_{XZ}\rho_{XY} + \rho_{YZ})Yf(X)f'(Y))| \\ &\leq (1 - \rho_{XY}^2)^{-1} |\rho_{XZ} - \rho_{YZ}\rho_{XY}| \times E|Xf(X)f'(Y)| + (1 - \rho_{XY}^2)^{-1} |-\rho_{XZ}\rho_{XY} + \rho_{YZ}| \times E|Yf(X)f'(Y)| \\ &\leq (1 - \rho_{XY}^2)^{-1} |\rho_{XZ} - \rho_{YZ}\rho_{XY}| \times f_{\max} f'_{\max} \sigma_x + (1 - \rho_{XY}^2)^{-1} |-\rho_{XZ}\rho_{XY} + \rho_{YZ}| \times f_{\max} c. \end{aligned}$$

□

Lemma 7. *Let $a_1, a_2, \dots, a_d, b_1, b_2, \dots, b_d$ be i.i.d. standard normal distribution. Then we have w.p. $1 - 3\delta$,*

$$\frac{\sum a_i b_i}{\sqrt{\sum_i a_i^2} \sqrt{\sum_i b_i^2}} \leq \frac{\sqrt{d} \log \frac{2}{\delta}}{d - \sqrt{d} \log \frac{2}{\delta}} \quad (\text{B.1})$$

Proof. Directly applying Chernoff bound for normal distribution. □

Lemma 8. *Let Z_1, Z_2 be two r.v.s with normal distribution. Let $\rho = \frac{V_{12}}{\sigma_1 \sigma_2}$, where V_{12} is the correlation between Z_1, Z_2 . Consider a function $\sigma(\cdot)$ such that $\sigma(x) = -\sigma(-x)$, $|\sigma(\cdot)| \leq \sigma_{\max}$, and $\sup_x \sigma(x)x = \alpha_G$. Then we have*

$$|E(\sigma(Z_1)\sigma(Z_2))| \leq \left(\frac{\sigma_{\max} + 2\sqrt{\rho}\sigma_{\max} + 4\alpha_G\sqrt{1-\rho}}{\sqrt{1-\rho^2}} \right) \sigma_{\max}\sqrt{\rho}.$$

Proof. Expanding the expectation, we have

$$\begin{aligned} E(\sigma(Z_1)\sigma(Z_2)) &= \int_{-\infty}^{+\infty} \sigma(z_1)\sigma(z_2) \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{\frac{1}{\sigma_1^2}z_1^2 - \frac{2\rho}{\sigma_1\sigma_2}z_1z_2 + \frac{1}{\sigma_2^2}z_2^2}{2(1-\rho^2)}\right) dz_1 dz_2 \\ &= \int_{-\infty}^{+\infty} \sigma(z_1)\sigma(z_2) \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{\frac{1}{\sigma_1^2}(z_1 - \frac{\rho\sigma_1}{\sigma_2}z_2)^2 + \frac{1-\rho^2}{\sigma_2^2}z_2^2}{2(1-\rho^2)}\right) dz_1 dz_2 \\ &= \int_{-\infty}^{+\infty} \sigma(z_3 + \frac{\rho\sigma_1}{\sigma_2}z_2)\sigma(z_2) \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-\rho^2}} \exp\left(-\frac{\frac{1}{\sigma_1^2}(z_3)^2 + \frac{1-\rho^2}{\sigma_2^2}z_2^2}{2(1-\rho^2)}\right) dz_3 dz_2 \\ &= \int_{-\infty}^{+\infty} \sigma(\sigma_1 u_3 + \rho\sigma_1 u_2)\sigma(\sigma_2 u_2) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 du_2 \end{aligned}$$

where we simply change the integration variable. Let $G(u) = \sigma(\sigma_1 u)$. Note that $|G(u)| \leq \sigma_{\max}$. The above integration becomes

$$\int_{-\infty}^{+\infty} G(u_3 + \rho u_2)G(\frac{\sigma_2}{\sigma_1}u_2) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 du_2.$$

Now First fix u_2 and decompose the integration into two parts. The first part is the integration on $(-\infty, -x) \cap (x, +\infty)$ and the second part is the integration on $[-x, x]$. First consider the case $u_2 \geq 0$. For the first part,

$$\begin{aligned} & \int_{-\infty}^{-x} + \int_x^{+\infty} G(u_3 + \rho u_2) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &= \int_x^{+\infty} (G(u_3 + \rho u_2) + G(-u_3 + \rho u_2)) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &= \int_x^{+\infty} (G(u_3 + \rho u_2) - G(u_3 - \rho u_2)) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3, \end{aligned}$$

where the last inequality is by the symmetry of the function G , i.e., $G(x) = -(G - x)$. Note that

$$|G'(y)| = |\sigma(\sigma_1 y) \sigma_1 y \frac{1}{y}| \leq \alpha_G \frac{1}{y}.$$

Let $x = \rho u_2 + w$, where $w \geq 0$. Then we have for all $u_3 \geq x$,

$$|G'(u_3 + \rho u_2)| \leq \alpha_G \frac{1}{u_3 + \rho u_2} \leq \alpha_G \frac{1}{w}$$

and

$$|G'(u_3 - \rho u_2)| \leq \alpha_G \frac{1}{u_3 - \rho u_2} \leq \alpha_G \frac{1}{w},$$

which in fact proves $G(y)$ is Lipschitz continuous with constant $\frac{\alpha_G}{w}$ for $y \geq w$. Thus, we now have

$$|G(u_3 + \rho u_2) - G(u_3 - \rho u_2)| \leq \alpha_G \frac{1}{w} 2\rho u_2.$$

Since G is monotone, we have

$$G(u_3 + \rho u_2) - G(u_3 - \rho u_2) \leq \alpha_G \frac{1}{w} 2\rho u_2.$$

Apply this inequality in the integration, we have

$$\begin{aligned} & \int_x^{+\infty} (G(u_3 + \rho u_2) - G(u_3 - \rho u_2)) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &\leq \int_x^{+\infty} \left(\frac{\alpha_G}{w} 2\rho u_2\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &\leq \int_{-\infty}^{+\infty} \left(\frac{\alpha_G}{w} 2\rho u_2\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &= \left(\frac{\alpha_G}{w} 2\rho u_2\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right). \end{aligned}$$

Now let us consider the second part of the integration.

$$\begin{aligned} & \int_{-x}^x G(u_3 + \rho u_2) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &\leq \int_{-x}^x \sigma_{\max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &\leq \int_{-x}^x \sigma_{\max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}\sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right) du_3 \\ &= 2x\sigma_{\max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}\sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right), \end{aligned}$$

where the first inequality is because $\sigma_{\max} \geq G$, the second inequality is because $\exp(-a^2) \leq 1$. Combing the the integration, we finally have

$$\begin{aligned} & \int_{-\infty}^{+\infty} G(u_3 + \rho u_2) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi\sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2)u_2^2}{2(1-\rho^2)}\right) du_3 \\ &\leq \left(\frac{\alpha_G}{w} 2\rho u_2\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right) + 2x\sigma_{\max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}\sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right), \end{aligned}$$

where $x = \rho u_2 + w$. Let $w = \frac{\alpha_G 2u_2 \sqrt{\rho} \sqrt{1-\rho}}{\sigma_{max}}$. We have

$$\begin{aligned}
& \int_{-\infty}^{+\infty} G(u_3 + \rho u_2) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi \sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2) u_2^2}{2(1-\rho^2)}\right) du_3 \\
& \left(\frac{\alpha_G}{w} 2\rho u_2\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right) + 2x \sigma_{max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi} \sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right) \\
& = \left(\frac{\sqrt{\rho} \sigma_{max}}{\sqrt{1-\rho^2}}\right) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right) \\
& \quad + 2\left(\rho u_2 + \frac{\alpha_G 2u_2 \sqrt{\rho} \sqrt{1-\rho}}{\sigma_{max}}\right) \sigma_{max} G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{\sqrt{2\pi} \sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right) \\
& \leq \left(\frac{\sqrt{\rho} \sigma_{max}^2}{\sqrt{1-\rho^2}}\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right) + 2\left(\rho u_2 + \frac{\alpha_G 2u_2 \sqrt{\rho} \sqrt{1-\rho}}{\sigma_{max}}\right) \sigma_{max}^2 \frac{1}{\sqrt{2\pi} \sqrt{1-\rho^2}} \exp\left(-\frac{u_2^2}{2}\right) \\
& = \left(\frac{\sqrt{\rho} \sigma_{max}^2 + 2\rho u_2 \sigma_{max}^2 + 4\alpha_G u_2 \sqrt{\rho} \sigma_{max} \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right),
\end{aligned}$$

where the first equation is because we plug in the expression of w , the inequality is due to $G \leq \sigma_{max}$. Similarly we can prove it for the case when $u \leq 0$,

$$\begin{aligned}
& \int_{-\infty}^{+\infty} G(u_3 + \rho u_2) G\left(\frac{\sigma_2}{\sigma_1} u_2\right) \frac{1}{2\pi \sqrt{1-\rho^2}} \exp\left(-\frac{u_3^2 + (1-\rho^2) u_2^2}{2(1-\rho^2)}\right) du_3 \\
& \leq \left(\frac{\sqrt{\rho} \sigma_{max}^2 + 2\rho(-u_2) \sigma_{max}^2 + 4\alpha_G(-u_2) \sqrt{\rho} \sigma_{max} \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right).
\end{aligned}$$

Thus, we have

$$\begin{aligned}
E(\sigma(Z_1)\sigma(Z_2)) & \leq \int_0^{+\infty} \left(\frac{\sqrt{\rho} \sigma_{max}^2 + 2\rho u_2 \sigma_{max}^2 + 4\alpha_G u_2 \sqrt{\rho} \sigma_{max} \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{u_2^2}{2}\right) du_2 \\
& = \left(\frac{\sqrt{\rho} \sigma_{max}^2 + 2\rho \sigma_{max}^2 + 4\alpha_G \sqrt{\rho} \sigma_{max} \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \\
& = \left(\frac{\sigma_{max} + 2\sqrt{\rho} \sigma_{max} + 4\alpha_G \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \sigma_{max} \sqrt{\rho}.
\end{aligned}$$

By symmetry, we have

$$-E(\sigma(Z_1)\sigma(Z_2)) \leq \left(\frac{\sigma_{max} + 2\sqrt{\rho} \sigma_{max} + 4\alpha_G \sqrt{1-\rho}}{\sqrt{1-\rho^2}}\right) \sigma_{max} \sqrt{\rho},$$

which completes the proof. \square

B.3 Main Proof

The main proof of the theorem consists of 4 lemmas, based on which the main theorem becomes straightforward.

Lemma 9. Suppose W, W^*, c_i are all i.i.d. random variables sampled from standard normal distribution. Then w.h.p,

$$E_{W_2, W_2^*} A_1 \geq \mathcal{O}(n^2 K^2 d).$$

Proof. Expanding the expression of A_1 , we have

$$\begin{aligned}
E_{W_2} A_1 &= E_{W_2} \left(n \sum_{i=1}^n (\hat{y}_i - y_i)^2 \left(\sum_{p=1}^K \sum_{q=1}^d \left(\frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 \right) \right) \\
&= E_{W_2} \left(n \sum_{i=1}^n (W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i))^2 \left(\sum_{p=1}^K \sum_{q=1}^d (W_{2,1,p} \sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \right) \\
&= n E_{W_2} \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i))^2 \left(\sum_{p=1}^K \sum_{q=1}^d (W_{2,1,p} \sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \right) \\
&= n E_{W_2} \left(\sum_{i=1}^n \sum_{r=1}^K (W_{2,1,r}^2 \sigma(W_1 x_i)^2 + W_{2,1,r,*}^2 \sigma(W_1^* x_i)^2) \left(\sum_{p=1}^K \sum_{q=1}^d (W_{2,1,p} \sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \right) \\
&\geq n \left(\sum_{i=1}^n \sum_{r=1}^K (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{p=1}^K \sum_{q=1}^d (\sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \right).
\end{aligned}$$

Since $\sigma(\cdot)$ is bounded by σ_{\max} , we have

$$\begin{aligned}
Pr(|\sigma(W_{1,r,:} x_i)^2 \sigma'(W_{1,p,:} x_i)^2 x_{i,q}^2| \geq t) &\leq Pr(|\sigma_{\max}^2 \sigma_{\max}'^2 x_{i,q}^2| \geq t) \\
&\leq 2 \exp\left(-\frac{t^2}{2\sigma_{\max}^4 \sigma_{\max}'^4}\right)
\end{aligned}$$

where the last equation is due to the fact that $x_{i,q}$ is standard normal distributed. This implies $\sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2$ is sub-exponential (where $x_{i,q}^2$ is chi-square). Thus, we can apply Bernstein inequality to $\sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2$, to obtain w.p. $1 - \delta$,

$$\sum_{i=1}^n \sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 - E_{x_i} \sum_{i=1}^n \sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 \geq -\sqrt{n} \sigma_{\max}^2 \sigma_{\max}'^2 \log \frac{2}{\delta}$$

and similarly w.p. $1 - \delta$,

$$\left| \sum_{i=1}^n \sigma(W_{1,r,:}^* x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 - E_{x_i} \sum_{i=1}^n \sigma(W_{1,r,:}^* x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 \right| \geq -\sqrt{n} \sigma_{\max}^2 \sigma_{\max}'^2 \log \frac{2}{\delta}$$

By union bound, we have w.p. $1 - 2\delta$, the above two are both true. Plug them in the expression of A_1 . Finally, we have w.p. $1 - 2\delta$,

$$\begin{aligned}
E_{W_2} A_1 &\geq n \left(\sum_{i=1}^n \sum_{r=1}^K (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{p=1}^K \sum_{q=1}^d (\sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \right) \\
&= n \sum_{p=1}^K \sum_{q=1}^d \sum_{r=1}^K \sum_{i=1}^n (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) (\sigma'(W_{1,p,:} x_i) x_{i,q})^2 \\
&= n \sum_{p=1}^K \sum_{q=1}^d \sum_{r=1}^K \sum_{i=1}^n \left(\sigma(W_{1,r,:} x_i)^2 (\sigma'(W_{1,p,:} x_i) x_{i,q})^2 + \sigma(W_{1,r,:}^* x_i)^2 (\sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right) \\
&\geq 2n^2 \left(\sum_{p=1}^K \sum_{q=1}^d \sum_{r=1}^K E_{x_i} \sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 + E_{x_i} \sigma(W_{1,r,:}^* x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 - a_0 \right) \\
&\geq 2n^2 d \left(\sum_{p=1}^K \sum_{r=1}^K E_{x_i} \sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 + E_{x_i} \sigma(W_{1,r,:}^* x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 - a_0 \right)
\end{aligned}$$

where $a_0 = 2 \frac{1}{\sqrt{n}} \sigma_{\max}^2 \sigma_{\max}'^2 \log \frac{2}{\delta}$ is the extra error term. Note that this term is small and typically can be ignored.

Since the term within summation is bounded, we can apply Hoeffding bound over p and r separately. Finally we will get w.p. $1 - 6\delta$,

$$\begin{aligned}
E_{W_2} A_1 &\geq 2n^2 K^2 d \left(E \sigma(W_{1,r,:} x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 + E \sigma(W_{1,r,:}^* x_i)^2 \sigma'^2(W_{1,p,:} x_i) x_{i,q}^2 - O\left(\frac{1}{\sqrt{n}} + \frac{1}{\sqrt{K}}\right) \log \frac{1}{\delta} \right) \\
&= O(n^2 K^2 d)
\end{aligned}$$

which completes the proof. \square

Lemma 10. Suppose W, W^*, x_i are all i.i.d. random variables sampled from standard normal distribution. Then w.h.p,

$$E_{W_2, W_2^*} A_2 \geq \mathcal{O}(n^2 K^2).$$

Proof.

$$\begin{aligned} E_{W_2, W_2^*} A_2 &= E_{W_2} \left(n \sum_{i=1}^n (\hat{y}_i - y_i)^2 \left(\sum_{q=1}^K \left(\frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right) \right) \\ &= E_{W_2} \left(n \sum_{i=1}^n \|W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i)\|^2 \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) \right) \\ &= n \sum_{i=1}^n \sum_{r=1}^K (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right). \end{aligned}$$

Now fix W_1 and thus r . Note that $\|(\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2\| \leq 2K\sigma_{\max}^4$. Applying Hoeffding bound to the term in the summation over the randomness of x_i , we have w.p. $(1 - \delta)$,

$$\begin{aligned} &\left| \sum_{i=1}^n (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) - n E_{x_i} (\sigma(W_{1,r,:} x_i)^2 \right. \\ &\quad \left. + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) \right| \leq \log \frac{2}{\delta} \sqrt{8n} K \sigma_{\max}^4. \end{aligned}$$

And thus we have w.p. $1 - \delta$,

$$\begin{aligned} E_{W_2, W_2^*} A_2 &\geq n \sum_{r=1}^K \left(n E_{x_i} (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) - \log \frac{2}{\delta} \sqrt{8n} K \sigma_{\max}^4 \right) \\ &= n^2 \sum_{r=1}^K \left(E_{x_i} (\sigma(W_{1,r,:} x_i)^2 + \sigma(W_{1,r,:}^* x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) - \log \frac{2}{\delta} \sqrt{\frac{8}{n}} K \sigma_{\max}^4 \right) \\ &\geq n^2 \sum_{r=1}^K \left(E_{x_i} (\sigma(W_{1,r,:} x_i)^2) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) - \log \frac{2}{\delta} \sqrt{\frac{8}{n}} K \sigma_{\max}^4 \right) \\ &\geq n^2 \left(\left(\sum_{r=1}^K E_{x_i} (\sigma(W_{1,r,:} x_i)^2) \right) \left(\sum_{q=1}^K (\sigma(W_{1,q,:} x_i))^2 \right) - \log \frac{2}{\delta} \sqrt{\frac{8}{n}} K^2 \sigma_{\max}^4 \right). \end{aligned}$$

Now let us apply Hoeffding bound over W_1 , we have w.p. $1 - \delta$,

$$\left| \sum_{r=1}^K \sigma^2(W_{1,r,:} x_i) - K E_{W_1} \sigma^2(W_{1,1,:} x_i) \right| \leq \log \frac{2}{\delta} \sqrt{K} \sigma_{\max}^2,$$

and similarly w.p. $1 - \delta$,

$$\left| \sum_{r=1}^K \sigma^2(W_{1,q,:} x_i) - K E_{W_1} \sigma^2(W_{1,1,:} x_i) \right| \leq \log \frac{2}{\delta} \sqrt{K} \sigma_{\max}^2.$$

Thus, w.p. $1 - 3\delta$, we have

$$\begin{aligned} E_{W_2, W_2^*} A_2 &\geq n^2 K^2 E_x \left((E_{W_1} \sigma^2(W_{1,1,:} x_i) - \log \frac{2}{\delta} \sigma_{\max}^2 \sqrt{\frac{1}{K}})^2 - \log \frac{2}{\delta} \sqrt{8n} K^2 \sigma_{\max}^4 \right) \\ &= n^2 K^2 E_x (E_{W_1} \sigma^2(W_{1,1,:} x_i))^2 - E_{W_1, x} \sigma^2(W_{1,1,:} x_i) (\log \frac{2}{\delta} \sigma_{\max}^2 \sqrt{K}) K n^2 \\ &\quad + n^2 (\log \frac{2}{\delta})^2 K \sigma_{\max}^4 - \log \frac{2}{\delta} \sqrt{8n} K \sigma_{\max}^4 \\ &\geq n^2 K^2 E_x (E_{W_1} \sigma^2(W_{1,1,:} x_i))^2 - (\log \frac{2}{\delta} \sigma_{\max}^4 \sqrt{K}) K n^2 \\ &\quad + n^2 (\log \frac{2}{\delta})^2 K \sigma_{\max}^4 - \log \frac{2}{\delta} \sqrt{8n} K \sigma_{\max}^4. \end{aligned}$$

Changing δ to $\frac{\delta}{3}$, we have w.p. $1 - \delta$,

$$\begin{aligned} E_{W_2, W_2^*} A_2 &\geq n^2 K^2 E_x (E_{W_1} \sigma^2(W_{1,1,:} x_i))^2 - (\log \frac{6}{\delta} \sigma_{max}^4 \sqrt{K}) K n^2 \\ &\quad + n^2 (\log \frac{6}{\delta})^2 K \sigma_{max}^4 - \log \frac{6}{\delta} \sqrt{8n} K \sigma_{max}^4 \\ &= \mathcal{O}(n^2 K^2). \end{aligned}$$

□

Lemma 11. Suppose W, W^*, x_i are all i.i.d. random variables sampled from standard normal distribution. Then w.h.p.,

$$E_{W_2, W_2^*} B_1 \leq \mathcal{O}(n^2 K^2).$$

Proof. Expanding the expression of B_2 , we have

$$\begin{aligned} E_{W_2, W_2^*} B_1 &= E_{W_2, W_2^*} \left(\sum_{p=1}^K \sum_{q=1}^d \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 \right) \\ &= E_{W_2, W_2^*} \left(\sum_{p=1}^K \sum_{q=1}^d \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \right) \\ &= \left(\sum_{p=1}^K \sum_{q=1}^d E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \right) \\ &\leq 2 \left(\sum_{p=1}^K \sum_{q=1}^d E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \right) \\ &\quad + 2 \left(\sum_{p=1}^K \sum_{q=1}^d E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_2^* \sigma(W_1^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \right) \end{aligned}$$

where the inequality is due to $(a + b)^2 \leq 2a^2 + 2b^2$. Expanding W_2 , we have

$$\begin{aligned} &E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &= E_{W_2, W_2^*} \left(\sum_{r=1}^K \sum_{i=1}^n (W_{2,1,r} \sigma(W_{1,r} x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &= \sum_{r=1}^K E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_{2,1,r} \sigma(W_{1,r} x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &\leq 3 \sum_{r=1}^K \left(\sum_{i=1}^n (\sigma(W_{1,r} x_i)) \sigma'(W_{1,p} x_i) x_{i,q} \right)^2, \end{aligned}$$

where the second equation is because $E(W_{2,1,r_1} W_{2,1,r_2} W_{2,1,p}^2) = 0$ as long as $r_1 \neq r_2$. The inequality is because $E(W_{2,1,r}^2 W_{2,1,p}^2) = 3$ if $r = p$ and $E(W_{2,1,r}^2 W_{2,1,p}^2) = 1 < 3$ if $r \neq p$. Similarly, we have

$$\begin{aligned} &E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_2^* \sigma(W_1^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &= E_{W_2, W_2^*} \left(\sum_{r=1}^K \sum_{i=1}^n (W_{2,1,r}^* \sigma(W_{1,r}^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &= \sum_{r=1}^K E_{W_2, W_2^*} \left(\sum_{i=1}^n (W_{2,1,r}^* \sigma(W_{1,r}^* x_i)) W_{2,1,p} \sigma'(W_{1,p} x_i) x_{i,q} \right)^2 \\ &= K \left(\sum_{i=1}^n \sigma(W_{1,r}^* x_i) \sigma'(W_{1,p} x_i) x_{i,q} \right)^2. \end{aligned}$$

Therefore, we obtain

$$E_{W_2, W_2^*} B_1 \leq 2 \sum_{p=1}^K \sum_{q=1}^d \left(3 \sum_{r=1}^K \left(\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right)^2 + K \left(\sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right)^2 \right).$$

Since $\sigma(\cdot) \leq \sigma_{\max}$ and $\sigma(\cdot)' \leq \sigma'_{\max}$, we have

$$\begin{aligned} Pr(|\sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q}| \geq t) &\leq Pr(|\sigma_{\max} \sigma'_{\max} x_{i,q}| \geq t) \\ &\leq 2 \exp\left(-\frac{t^2}{2\sigma_{\max}^2 \sigma_{\max}^{'2}}\right) \end{aligned}$$

which implies $\sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q}$ is sub-exponential. Thus, we can apply Bernstein inequality to $\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q}$, to obtain w.p. $1 - \delta$,

$$\begin{aligned} & \left| \sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} - n E_{x_i} \sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right| \\ & \leq \sqrt{\frac{n}{2} \log \frac{2}{\delta}} \sigma_{\max} \sigma'_{\max} + \sigma_{\max} \sigma'_{\max} \log \frac{2}{\delta} \end{aligned}$$

and similarly w.p. $(1 - \delta)$,

$$\begin{aligned} & \left| \sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} - n E_{x_i} \sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right| \\ & \leq \sqrt{\frac{n}{2} \log \frac{2}{\delta}} \sigma_{\max} \sigma'_{\max} + \sigma_{\max} \sigma'_{\max} \log \frac{2}{\delta}. \end{aligned}$$

By union bound, we have w.p. $1 - 2\delta$, the above two are both true. Plug them in the expression of B_1 and use $(a + b)^2 \leq 2a^2 + 2b^2$. Finally, we have w.p. $1 - 2\delta$,

$$\begin{aligned} & E_{W_2, W_2^*} B_1 \\ & \leq 2 \sum_{p=1}^K \sum_{q=1}^d \left(3 \sum_{r=1}^K \left(\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right)^2 + K \left(\sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} \right)^2 \right) \\ & \leq 2 \sum_{p=1}^K \sum_{q=1}^d \left(3 \sum_{r=1}^K (n E_{x_i} \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} + a_1)^2 \right. \\ & \quad \left. + K (n E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q} + a_1)^2 \right) \\ & \leq 4n^2 \sum_{p=1}^K \sum_{q=1}^d \left(3 \sum_{r=1}^K (E_{x_i} \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q})^2 \right. \\ & \quad \left. + K (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma'(W_{1,p,:} x_i) x_{i,q})^2 + 8K a_1^2 \right), \end{aligned}$$

where $a_1 = \sqrt{\frac{1}{2n} \log \frac{2}{\delta}} \sigma_{\max} \sigma'_{\max} + \frac{1}{n} \sigma_{\max} \sigma'_{\max} \log \frac{2}{\delta}$ is the extra error term. Note that this term is $O(\sqrt{\frac{1}{n}})$ and typically can be ignored. Now let us consider $E_{x_i} \sigma(W_{1,r,:} x_i) \sigma'(W_{1,p,:} x_i) x_{i,q}$. Abuse the notation a little bit, let $X = W_{1,r,:} x_i$, $Y = W_{1,p,:} x_i$, $Z = x_{i,q}$. Given W_1 , they are all normal distribution, and the correlation is

$$\begin{aligned} \rho_{XZ} &= \frac{W_{1,r,q}}{\sum_{b=1}^d W_{1,r,b}^2} \\ \rho_{YZ} &= \frac{W_{1,p,q}}{\sum_{b=1}^d W_{1,p,b}^2} \\ \rho_{XY} &= \frac{\sum_{b=1}^d W_{1,p,b} W_{1,r,b}}{\sqrt{\sum_{b=1}^d W_{1,p,b}^2 \sum_{b=1}^d W_{1,r,b}^2}} \\ \sigma_x &= \sqrt{\sum_{b=1}^d W_{1,r,b}^2}. \end{aligned}$$

Apply lemma 6, we can get,

$$\begin{aligned} & |E(\sigma(X)\sigma'(Y)Z)| \\ & \leq (1 - \rho_{XY}^2)^{-1} |\rho_{XZ} - \rho_{YZ}\rho_{XY}| \times \sigma_{max}\sigma'_{max}\sigma_x + (1 - \rho_{XY}^2)^{-1} |-\rho_{XZ}\rho_{XY} + \rho_{YZ}| \times \sigma_{max}\alpha_G. \end{aligned}$$

Note that by Chernoff bound and Lemma 7, w.p. $1 - 4\delta$,

$$\begin{aligned} |\rho_{XZ}| & \leq \frac{2\log \frac{1}{\delta}}{d - \sqrt{d}\log \frac{1}{\delta}} \\ |\rho_{YZ}| & \leq \frac{2\log \frac{1}{\delta}}{d - \sqrt{d}\log \frac{1}{\delta}} \\ |\rho_{XY}| & \leq \sqrt{d} \frac{2\log \frac{1}{\delta}}{d - \sqrt{d}\log \frac{1}{\delta}} \\ \sigma_x & \leq \sqrt{d + \sqrt{d}\log \frac{1}{\delta}}. \end{aligned}$$

Plug them in the above inequality, we have w.p. $1 - 4\delta$,

$$E_{x_i}\sigma(W_{1,r,:}x_i)\sigma'(W_{1,p,:}x_i)x_{i,q} = |E(\sigma(X)\sigma'(Y)Z)| \leq O\left(\frac{1}{\sqrt{d}}\right)\sigma_{max}\sigma'_{max} + O\left(\frac{1}{d}\right)\sigma_{max}\alpha_G.$$

Similarly we can get w.p. $1 - 4\delta$,

$$E_{x_i}\sigma(W_{1,r,:}^*x_i)\sigma'(W_{1,p,:}x_i)x_{i,q} \leq O\left(\frac{1}{\sqrt{d}}\right)\sigma_{max}\sigma'_{max} + O\left(\frac{1}{d}\right)\sigma_{max}\alpha_G.$$

Thus, we have w.p. $1 - 10\delta$,

$$\begin{aligned} E_{W_2, W_2^*} B_1 & \leq 4n^2 \sum_{p=1}^K \sum_{q=1}^d \left(3 \sum_{r=1}^K (E_{x_i}\sigma(W_{1,r,:}x_i)\sigma'(W_{1,p,:}x_i)x_{i,q})^2 \right. \\ & \quad \left. + K (E_{x_i}\sigma(W_{1,r,:}^*x_i)\sigma'(W_{1,p,:}x_i)x_{i,q})^2 \right) \\ & \leq 4n^2 \sum_{p=1}^K \sum_{q=1}^d \left(3K \left(\mathcal{O}\left(\frac{1}{\sqrt{d}}\sigma_{max}\sigma'_{max}\right) \right)^2 + K \left(\mathcal{O}\left(\frac{1}{\sqrt{d}}\sigma_{max}\sigma'_{max}\right) \right)^2 \right) \\ & = \mathcal{O}(K^2 n^2), \end{aligned}$$

which completes the proof. \square

Lemma 12. Suppose W, W^*, x_i are all i.i.d. random variables sampled from standard normal distribution. Then w.h.p,

$$E_{W_2, W_2^*} B_2 \leq \mathcal{O}\left(n^2 K^2 \left(\frac{1}{d} + \frac{1}{\sqrt{d}} + \frac{1}{K}\right)\right).$$

Proof. Expanding the expression of B_2 , we have

$$\begin{aligned} E_{W_2, W_2^*} B_2 & = E_{W_2, W_2^*} \left(\sum_{q=1}^K \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right) \\ & = E_{W_2, W_2^*} \left(\sum_{q=1}^K \left(\sum_{i=1}^n (W_2 \sigma(W_1 x_i) - W_2^* \sigma(W_1^* x_i)) \sigma(W_{1,q,:} x_i) \right)^2 \right) \\ & = \sum_{q=1}^K \sum_{r=1}^K \left(\left(\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i) \right)^2 + \left(\sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i) \right)^2 \right), \end{aligned}$$

where the third equation is because W_2, W_2^* are independent. Applying Hoeffding bound to $\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i)$, we have w.p. $(1 - \delta)$,

$$\left| \sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i) - n E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i) \right| \leq \sqrt{\frac{n}{2}} \sigma_{max}^2 \log \frac{2}{\delta}$$

and similarly w.p. $(1 - \delta)$,

$$\left| \sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i) - n E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i) \right| \leq \sqrt{\frac{n}{2}} \sigma_{max}^2 \log \frac{2}{\delta}$$

By union bound, we have w.p. $1 - 2\delta$, the above two are both true. Plug them in the expression of B_2 , we have w.p. $1 - 2\delta$,

$$\begin{aligned} E_{W_2, W_2^*} B_2 &= \sum_{q=1}^K \sum_{r=1}^K \left(\left(\sum_{i=1}^n \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i) \right)^2 + \left(\sum_{i=1}^n \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i) \right)^2 \right) \\ &\leq 2 \sum_{q=1}^K \sum_{r=1}^K \left(n^2 (E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 + n^2 (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i))^2 + n \sigma_{max}^4 \log^2 \frac{2}{\delta} \right) \\ &= 2n^2 \sum_{q=1}^K \sum_{r=1}^K \left((E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i))^2 \right) \\ &\quad + 2K^2 n \sigma_{max}^4 \log^2 \frac{2}{\delta} \\ &= 2n^2 \sum_{q=1}^K \sum_{r \neq q}^K \left((E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i))^2 \right) \\ &\quad + 2n^2 \sum_{q=1}^K \left((E_{x_i} \sigma^2(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma^2(W_{1,q,:}^* x_i))^2 \right) \\ &\quad + 2K^2 n \sigma_{max}^4 \log^2 \frac{2}{\delta}, \end{aligned}$$

where the first inequality is due to $(a + b)^2 \leq 2(a^2 + b^2)$.

For the first term, define

$$\rho_{q,r} = \frac{E_{x_i} (W_{1,r,:} x_i W_{1,q,:} x_i)}{E_{x_i} (W_{1,r,:} x_i)^2 E_{x_i} (W_{1,q,:} x_i)^2} = \frac{\sum_{y=1}^d W_{1,r,y} W_{1,q,y}}{\sqrt{\sum_{y=1}^d W_{1,r,y}^2} \sqrt{\sum_{y=1}^d W_{1,q,y}^2}}.$$

Apply lemma 8 to each $(E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2$. We have

$$\begin{aligned} &(E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 \\ &\leq \left(\left(\frac{\sigma_{max} + 2\sqrt{\rho_{q,r}} \sigma_{max} + 4\alpha_G \sqrt{1 - \rho_{q,r}}}{\sqrt{1 - \rho_{q,r}^2}} \right) \sigma_{max} \sqrt{\rho_{q,r}} \right)^2 \\ &= \left(\frac{\sigma_{max} + 2\sqrt{\rho_{q,r}} \sigma_{max} + 4\alpha_G \sqrt{1 - \rho_{q,r}}}{\sqrt{1 - \rho_{q,r}^2}} \right)^2 \sigma_{max}^2 \rho_{q,r} \\ &\leq 3 \left(\frac{\sigma_{max}^2 + 4\rho_{q,r} \sigma_{max}^2 + 16\alpha_G^2 (1 - \rho_{q,r})}{\sqrt{1 - \rho_{q,r}^2}} \right) \sigma_{max}^2 \rho_{q,r} \end{aligned}$$

where the last inequality is due to $3a^2 + 3b^2 + 3c^2 \geq (a + b + c)^2$. By Lemma 7, we have w.p. $(1 - 3\delta)$,

$$\rho_{q,r} \leq \frac{\sqrt{d} \log \frac{2}{\delta}}{d - \sqrt{d} \log \frac{2}{\delta}}.$$

Therefore, plug in this value into the above inequality, we have w.p. $1 - 3\delta$,

$$\begin{aligned} &(E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 \\ &\leq 3 \left(\frac{\sigma_{max}^2 + 4\rho_{q,r} \sigma_{max}^2 + 16\alpha_G^2 (1 - \rho_{q,r})}{\sqrt{1 - \rho_{q,r}^2}} \right) \sigma_{max}^2 \rho_{q,r} \\ &\simeq 3 \left(\frac{\sigma_{max}^2 + 4 \frac{\sqrt{d} \log \frac{2}{\delta}}{d} \sigma_{max}^2 + 16\alpha_G^2 (1 - \frac{\sqrt{d} \log \frac{2}{\delta}}{d})}{\sqrt{1 - (\frac{\sqrt{d} \log \frac{2}{\delta}}{d})^2}} \right) \sigma_{max}^2 \frac{\sqrt{d} \log \frac{2}{\delta}}{d} \\ &= \mathcal{O} \left(\frac{1}{\sqrt{d}} + \frac{1}{d} \right) \log \frac{2}{\delta}. \end{aligned}$$

Apply this for all (q, r) pairs, and then use union bound, we have w.p. $1 - \delta$, for all $q \neq r$,

$$(E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 \leq \mathcal{O}\left(\frac{1}{\sqrt{d}} + \frac{1}{d}\right) \log \frac{2K(K-1)}{\delta}$$

and thus

$$\begin{aligned} 2n^2 \sum_{q=1}^K \sum_{r \neq q} \left((E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i))^2 \right) \\ \leq \mathcal{O}\left(n^2 K^2 \left(\frac{1}{\sqrt{d}} + \frac{1}{d}\right) \log \frac{2K(K-1)}{\delta}\right) \\ \simeq \mathcal{O}\left(n^2 K^2 \left(\frac{1}{\sqrt{d}} + \frac{1}{d}\right) \log \frac{1}{\delta}\right). \end{aligned}$$

For the second term, noting that $\sigma() \leq \sigma_{\max}$, we have

$$2n^2 \sum_{q=1}^K \left((E_{x_i} \sigma^2(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma^2(W_{1,q,:}^* x_i))^2 \right) \leq 2n^2 \sum_{q=1}^K (2\sigma_{\max}^2) = \mathcal{O}(n^2 K).$$

Combing all those terms, we have w.p. $1 - \delta$,

$$\begin{aligned} E_{W_2, W_2^*} B_2 &\leq 2n^2 \sum_{q=1}^K \sum_{r \neq q} \left((E_{x_i} \sigma(W_{1,r,:} x_i) \sigma(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma(W_{1,r,:}^* x_i) \sigma(W_{1,q,:} x_i))^2 \right) \\ &\quad + 2n^2 \sum_{q=1}^K \left((E_{x_i} \sigma^2(W_{1,q,:} x_i))^2 + (E_{x_i} \sigma^2(W_{1,q,:}^* x_i))^2 \right) \\ &\quad + 2K^2 n \sigma_{\max}^4 \log^2 \frac{2}{\delta} \\ &\leq \mathcal{O}\left(\frac{1}{\sqrt{d}} + \frac{1}{d}\right) \log \frac{2}{\delta} + \mathcal{O}(n^2 K) + \mathcal{O}(nK^2) \simeq \mathcal{O}\left(n^2 K^2 \left(\frac{1}{\sqrt{d}} + \frac{1}{d}\right) + n^2 K\right). \end{aligned}$$

This completes the proof. \square

Now we are ready to prove Theorem 2.

Proof. Note that $f_i = (\hat{y}_i - y_i)$, we can apply the chain rule to have

$$\begin{aligned} n \sum_{i=1}^n \|\nabla f_i\|_2^2 &= n \sum_{i=1}^n \left(\sum_{p=1}^K \sum_{q=1}^d \left(\frac{\partial f_i}{\partial W_{1,p,q}} \right)^2 + \sum_{q=1}^K \left(\frac{\partial f_i}{\partial W_{2,1,q}} \right)^2 \right) \\ &= n \sum_{i=1}^n (\hat{y}_i - y_i)^2 \left(\sum_{p=1}^K \sum_{q=1}^d \left(\frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 + \sum_{q=1}^K \left(\frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right) \\ &= A_1 + A_2, \end{aligned}$$

and

$$\begin{aligned} \left\| \sum_{i=1}^n \nabla f_i \right\|_2^2 &= \left(\sum_{p=1}^K \sum_{q=1}^d \left(\sum_{i=1}^n \frac{\partial f_i}{\partial W_{1,p,q}} \right)^2 + \sum_{q=1}^K \left(\sum_{i=1}^n \frac{\partial f_i}{\partial W_{2,1,q}} \right)^2 \right) \\ &= \left(\sum_{p=1}^K \sum_{q=1}^d \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{1,p,q}} \right)^2 + \sum_{q=1}^K \left(\sum_{i=1}^n (\hat{y}_i - y_i) \frac{\partial \hat{y}_i}{\partial W_{2,1,q}} \right)^2 \right) \\ &= B_1 + B_2. \end{aligned}$$

The goal is now to understand the behavior of $\frac{E_{W_2, W_2^*} A_1 + A_2}{E_{W_2, W_2^*} B_1 + B_2}$.

By Lemma 9, w.h.p,

$$E_{W_2, W_2^*} A_1 \geq \mathcal{O}(n^2 K^2 d).$$

By Lemma 10, we have w.h.p,

$$E_{W_2, W_2^*} A_2 \geq \mathcal{O}(n^2 K^2).$$

By Lemma 11, we have w.h.p,

$$E_{W_2, W_2^*} B_1 \leq \mathcal{O}(n^2 K^2).$$

By Lemma 12, we have w.h.p,

$$E_{W_2, W_2^*} B_2 \leq \mathcal{O}\left(n^2 K^2 \left(\frac{1}{d} + \frac{1}{\sqrt{d}} + \frac{1}{K}\right)\right).$$

Combing these four results we directly obtain the desired theorem. \square

C Proofs for Multilayer Linear Neural Networks

We first present the main theorem for multilayer NNs.

Theorem 6. *Consider a LNN with $L \geq 2$ layers. Let the weight values $W_{l,p,q}$ for $l \in \{1, \dots, L\}$ and \mathbf{x}_i be independently drawn random variables from $\mathcal{N}(0, 1)$. Let*

$$M = n^2 \prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2)$$

Then:

$$\begin{aligned} \mathbb{E}[n \sum_{i=1}^n \|\nabla f_i\|^2] &= M \cdot L \left(1 + \prod_{\ell=1}^{L-1} \frac{K_\ell}{K_\ell + 2}\right), \\ \mathbb{E}\left[\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle\right] &= M \cdot \frac{n-1}{n} \left(\sum_{\phi=0}^{L-1} \frac{L-\phi}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{L}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2}\right). \end{aligned}$$

Given Theorem 6, Theorem 3, the main theorem for multilayer NNs, becomes a direct corollary.

Next, we prove Theorem 6. We start by stating a few general lemmas that will be necessary in the proof.

C.1 Models, Assumptions and Notations

Let us denote $W = \prod_{\ell=1}^L W_\ell$ and $W^* = \prod_{\ell=1}^L W_\ell^*$. We will also need the following notation:

$$r_{a,p} = \left(\prod_{\ell=a+2}^L W_\ell\right) W_{a+1, :, p} \quad l_{a,q} = W_{a-1, q, :} \left(\prod_{\ell=1}^{a-2} W_\ell\right) \quad l_{a,q,s}^b = W_{a-1, q, :} \left(\prod_{\ell=b}^{a-2} W_\ell\right) W_{b-1, :, s},$$

where

$$\left(\prod_{\ell=a+2}^L W_\ell\right) W_{a+1, :, p} = \begin{cases} W_{L, :, p} & \text{if } a = L-1 \\ 1 & \text{if } a = L \end{cases} \quad W_{a-1, q, :} \left(\prod_{\ell=1}^{a-2} W_\ell\right) = \begin{cases} e_q^T & \text{if } a = 1 \\ W_{1, q, :} & \text{if } a = 2 \end{cases},$$

where $e_q \in \mathbb{R}^d$ is the q -th unit vector in the d dimensional space.

Then we can write the differential as follows:

$$\frac{\partial f_i}{\partial W_{a,p,q}} = (\hat{y}_i - y_i) \left(\prod_{\ell=a+2}^L W_\ell\right) W_{a+1, :, p} W_{a-1, q, :} \left(\prod_{\ell=1}^{a-2} W_\ell\right) x_i = (W x_i - W^* x_i) r_{a,p} (l_{a,q} x_i).$$

Furthermore, let $W_{(a+2):L} = \prod_{\ell=a+2}^L W_\ell$.

By default, we define $\prod_{i=k}^n a_i = 1$ if $k > n$.

C.2 Some Helper Lemmas

We would need the Isserlis Theorem (author?) [49]. The following lemma can derived from the Isserlis Theorem.

Lemma 13. Let $x \in \mathbb{R}^d$ such that x_i is i.i.d. $\sim \mathcal{N}(0, 1)$. Then

$$\begin{aligned} E_x (a^\top x)^2 &= \|a\|_2^2 \\ E_x (a^\top x)^4 &= 3\|a\|_2^4 \\ E_x (a^\top x)^8 &= 105\|a\|_2^8 \\ E_x (a^\top x b^\top x) &= a \cdot b \\ E_x (a^\top x)^2 (b^\top x)^2 &= 2\|a \cdot b\|^2 + \|a\|_2^2 \|b\|_2^2 \\ E_x (a^\top x)^4 (b^\top x)^4 &\leq 105\|a\|_2^4 \|b\|_2^4 \leq 105 \left(E_x (a^\top x)^2 (b^\top x)^2 \right)^2. \end{aligned}$$

Lemma 14. Let $B = (b_1, b_2, \dots, b_{d_a}) \in \mathbb{R}^{d_c \times d_a}$ be a random matrix whose elements are all i.i.d $\mathcal{N}(0, 1)$, and $c = (c_1, c_2, \dots, c_{d_c}) \in \mathbb{R}^{d_c}$ be a constant. Let $a_i = c^\top b_i$. Then we have

$$E \left(\sum_{i=1}^{d_a} a_i^2 \right)^2 = d_a (d_a + 2) \left(\sum_{j=1}^{d_c} c_j^2 \right)^2. \quad (\text{C.1})$$

Proof. By linearity of Expectation, we have

$$\begin{aligned} E \left(\sum_{i=1}^{d_a} a_i^2 \right)^2 &= E \left(\sum_{i=1}^{d_a} \sum_{j=1, j \neq i}^{d_a} a_i^2 a_j^2 \right) + E \left(\sum_{i=1}^{d_a} a_i^4 \right) \\ &= \sum_{i=1}^{d_a} \sum_{j=1, j \neq i}^{d_a} E \left((c^\top b_i)^2 (c^\top b_j)^2 \right) + \sum_{i=1}^{d_a} E \left((c^\top b_i)^4 \right) \\ &= \sum_{i=1}^{d_a} \sum_{j=1, j \neq i}^{d_a} \left(\sum_{k=1}^{d_c} c_k^2 \right)^2 + \sum_{i=1}^{d_a} 3 \left(\sum_{k=1}^{d_c} c_k^2 \right)^2 \\ &= d_a (d_a + 2) \left(\sum_{j=1}^{d_c} c_j^2 \right)^2, \end{aligned} \quad (\text{C.2})$$

where the third equation is due to Lemma 13. \square

Lemma 15. Let $G = (g_1; g_2; \dots; g_{d_g}) \in \mathbb{R}^{d_g \times d_a}$ be a random matrix whose elements are all i.i.d $\mathcal{N}(0, 1)$, $x_s, x_t \in \mathbb{R}^{d_x}$ be i.i.d $\mathcal{N}(0, 1)$, and $A \in \mathbb{R}^{d_a \times d_x}$ be a constant. Then we have

$$E \left(\sum_{j=1}^{d_g} g_j A x_s g_j A x_t \right)^2 = \sum_{i=1}^{d_g} \sum_{j=1}^{d_g} E \left(\sum_{u=1}^{d_x} g_j A_{:,u} g_i A_{:,u} \right)^2. \quad (\text{C.3})$$

Proof. By linearity of expectation, we have

$$\begin{aligned} E \left(\sum_{j=1}^{d_g} g_j A x_s g_j A x_t \right)^2 &= \sum_{i=1}^{d_g} \sum_{j=1}^{d_g} E (g_j A x_s g_j A x_t g_i A x_s g_i A x_t) \\ &= \sum_{i=1}^{d_g} \sum_{j=1}^{d_g} E \left(\sum_{u=1}^{d_x} g_j A_{:,u} g_i A_{:,u} \sum_{v=1}^{d_x} g_j A_{:,v} g_i A_{:,v} \right) \\ &= \sum_{i=1}^{d_g} \sum_{j=1}^{d_g} E \left(\sum_{u=1}^{d_x} g_j A_{:,u} g_i A_{:,u} \right)^2, \end{aligned} \quad (\text{C.4})$$

where the second equality is taking expectation over x . \square

Lemma 16. if all elements in W are i.i.d standard normal distribution, we have

$$E \left(\sum_{s=1}^{K_a} r_{a,s}^2 r_{a,p}^2 \right) = (K_a + 2) F_a = (K_a + 2) \left(\prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) \right).$$

Proof.

$$\begin{aligned}
& E_{W_{a+1}} \left(\sum_{s=1}^{K_a} r_{a,s}^2 r_{a,p}^2 \right) \\
&= E_{W_{a+1}} \left(\sum_{s=1}^{K_a} (W_{(a+2):L} W_{a+1,:s})^2 (W_{(a+2):L} W_{a+1,:p})^2 \right) \\
&= E_{W_{a+1}} \left(\sum_{s=1, s \neq p}^{K_a} (W_{(a+2):L} W_{a+1,:s})^2 (W_{(a+2):L} W_{a+1,:p})^2 + (W_{(a+2):L} W_{a+1,:p})^4 \right) \\
&= (K_a + 2) \left(E_{W_{a+1}} \left((W_{(a+2):L} W_{a+1,:p})^2 \right) \right)^2 \\
&= (K_a + 2) \left(\sum_{r=1}^{K_{a+1}} (W_{(a+3):L} W_{a+2,:,r})^2 \right)^2.
\end{aligned}$$

The first equation expands the expression of the original formula. The second equation split the summation over s into two parts, the case when $p = s$ and the case $p \neq s$. The third equation uses the fact that $E \|a^T x\|_2^4 = 3 \|a\|_2^4$ from Lemma 13, and that all $W_{a+1, :, s}$ are symmetric and thus the expectation of the sum is essentially $K_a - 1$ times the expectation of each value.

By Lemma 14, we have

$$E_{W_{a+2}} \left(\left(\sum_{r=1}^{K_{a+1}} (W_{(a+3):L} W_{a+2,:,r})^2 \right)^2 \right) = K_{a+1} (K_{a+1} + 2) \left(\sum_{r=1}^{K_{a+2}} (W_{a+4:L} W_{a+3,:,r})^2 \right)^2.$$

Note that now the formula on the right side has the same form of that on the left side. This actually means that we can use induction over a to further simplify it. Formally, let

$$F_a = \left(\sum_{r=1}^{K_{a+1}} (W_{(a+3):L} W_{a+2,:,r})^2 \right)^2.$$

Then the above equation becomes for all $a \leq L - 4$,

$$E_{W_{a+2}} (F_a) = K_{a+1} (K_{a+1} + 2) F_{a+1},$$

which implies

$$E(F_a) = K_{a+1} (K_{a+1} + 2) E(F_{a+1}).$$

Now we prove that by induction,

$$E(F_a) = \left(\prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) \right), \forall a \leq L - 3.$$

When $a = L - 3$, we have

$$\begin{aligned}
E(F_{L-3}) &= E \left(\sum_{r=1}^{K_{L-2}} (W_L W_{L-1, :, r})^2 \right)^2 \\
&= E \left(\sum_{r=1}^{K_{L-2}} \sum_{v=1}^{K_{L-2}} (W_L W_{L-1, :, r})^2 (W_L W_{L-1, :, v})^2 \right) \\
&= E \left(\sum_{r=1}^{K_{L-2}} \sum_{v=1, v \neq r}^{K_{L-2}} (W_L W_{L-1, :, r})^2 (W_L W_{L-1, :, v})^2 + \sum_{r=1}^{K_{L-2}} (W_L W_{L-1, :, r})^4 \right) \\
&= \sum_{r=1}^{K_{L-2}} \sum_{v=1, v \neq r}^{K_{L-2}} E(W_L W_{L-1, :, r})^2 (W_L W_{L-1, :, v})^2 + \sum_{r=1}^{K_{L-2}} E(W_L W_{L-1, :, r})^4 \\
&= \sum_{r=1}^{K_{L-2}} \sum_{v=1, v \neq r}^{K_{L-2}} E_{W_L} E_{W_{L-1}} (W_L W_{L-1, :, r})^2 (W_L W_{L-1, :, v})^2 \\
&\quad + \sum_{r=1}^{K_{L-2}} E_{W_L} E_{W_{L-1}} (W_L W_{L-1, :, r})^4 \\
&= \sum_{r=1}^{K_{L-2}} \sum_{v=1, v \neq r}^{K_{L-2}} E_{W_L} (E_{W_{L-1}} (W_L W_{L-1, :, r})^2)^2 + \sum_{r=1}^{K_{L-2}} E_{W_L} E_{W_{L-1}} (W_L W_{L-1, :, r})^4 \\
&= \sum_{r=1}^{K_{L-2}} \sum_{v=1, v \neq r}^{K_{L-2}} E_{W_L} \left(\sum_{t=1}^{K_{L-1}} (W_{L, :, t})^2 \right)^2 + 3 \sum_{r=1}^{K_{L-2}} E_{W_L} \left(\sum_{t=1}^{K_{L-1}} (W_{L, :, t})^2 \right)^2 \\
&= K_{L-2} (K_{L-2} + 2) E_{W_L} \left(\sum_{t=1}^{K_{L-1}} (W_{L, :, t})^2 \right)^2 \\
&= K_{L-2} (K_{L-2} + 2) K_{L-1} (K_{L-1} + 2).
\end{aligned}$$

The first and second equations expand the expression of F_L . The third equation splits the summation over v into 2 parts, the case when $v = r$ and the case when $v \neq r$. The forth equation uses the linearity of expectation and the fifth equation uses conditional expectation. The sixth equation uses the fact that W_{L-1} are all i.i.d standard normal distribution. The seventh equation uses the fact that $E\|a^T x\|_2^4 = 3\|a\|^4$ from Lemma 13. The last two equations are simple algebra.

Assume that when $a = \theta$,

$$E(F_\theta) = \left(\prod_{\ell=\theta+1}^{L-1} K_\ell (K_\ell + 2) \right).$$

When $a = \theta - 1$, we have

$$\begin{aligned}
E(F_{\theta-1}) &= K_\theta (K_\theta + 2) E(F_\theta) \\
&= \left(\prod_{\ell=\theta}^{L-1} K_\ell (K_\ell + 2) \right).
\end{aligned}$$

Thus, by induction,

$$E(F_a) = \left(\prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) \right), \forall a \leq L - 3.$$

Therefore,

$$E \left(\sum_{s=1}^{K_a} r_{a,s}^2 r_{a,p}^2 \right) = (K_a + 2) E(F_a) = (K_a + 2) \left(\prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) \right).$$

□

Lemma 17. *If all elements in W, W^*, x are i.i.d standard normal distribution, then*

$$E \left(\sum_{t=1}^{K_{a-1}} (l_{a,q} x_i)^2 (l_{a,t} x_i)^2 \right) = (K_{a-1} + 2) \left(\prod_{\ell=0}^{a-2} K_{\ell} (K_{\ell} + 2) \right).$$

Proof. Similar to the proof of Lemma 16. □

Lemma 18. *If all elements in W, W^*, x are i.i.d standard normal distribution, then*

$$E \left(\sum_{t=1}^{K_{a-1}} (l_{a,q,v})^2 (l_{a,t,v})^2 \right) = (K_{a-1} + 2) \left(\prod_{\ell=1}^{a-2} K_{\ell} (K_{\ell} + 2) \right).$$

Proof. Similar to the proof of Lemma 16. □

Lemma 19.

$$E \left(\sum_{s,t=1}^{K_{b-2}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) = \left(\prod_{\ell=b-2}^{a-1} K_{\ell} (K_{\ell} - 1) \right) \left(\sum_{\phi=b-2}^{a-1} \frac{1}{K_{\phi} - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1} \right),$$

and in particular,

$$E \left(\sum_{s,t=1}^{K_0} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^0 l_{a,v,t}^0 \right)^2 \right) = \left(\prod_{\ell=0}^{a-1} K_{\ell} (K_{\ell} - 1) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_{\phi} - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1} \right).$$

Proof. We will prove the result using recurrent formula. Let us first note that

$$\begin{aligned} & E \left(\sum_{s,t=1}^{K_{b-2}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) \\ &= K_{b-2} E \left(\left(\sum_{v=1}^{K_{a-1}} (l_{a,v,s}^b)^2 \right)^2 \right) + K_{b-2} (K_{b-2} - 1) E \left(\left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) \\ &= K_{b-2} \prod_{\ell=b-1}^{a-1} K_{\ell} (K_{\ell} + 2) + K_{b-2} (K_{b-2} - 1) E \left(\left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right), \end{aligned}$$

where the first equation splits the summation over s into two cases, the case when $s = t$ and the case when $s \neq t$. Note that $l_{a,v,s}^b$ are symmetric over all s , and thus the summation in the first case becomes K_{b-2} times a single term. Similarly, in the second term, we have $K_{b-2} K_{b-2} - 1$ as the coefficient. The second equation essentially plugs in the value of the $E \left(\left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b \right)^2 \right)$, which is a sum of squares and we already know how to compute it using lemma 14.

Now let us turn to the term $E \left(\left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right)$. Let us further split W_b in this is term, and we obtain

$$\begin{aligned} E \left(\left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) &= E \left(\sum_{v_1=1}^{K_{a-1}} \sum_{v_2=1}^{K_{a-1}} l_{a,v_1,s}^b l_{a,v_1,t}^b l_{a,v_2,s}^b l_{a,v_2,t}^b \right) \\ &= E \left(\sum_{v_1=1}^{K_{a-1}} \sum_{v_2=1}^{K_{a-1}} l_{a,v_1,s}^{b+1} W_{b, :, s} l_{a,v_1,t}^{b+1} W_{b, :, t} l_{a,v_2,s}^{b+1} W_{b, :, s} l_{a,v_2,t}^{b+1} W_{b, :, t} \right) \\ &= E \left(\sum_{v_1=1}^{K_{a-1}} \sum_{v_2=1}^{K_{a-1}} \sum_{s=1}^{K_{b-1}} \sum_{t=1}^{K_{b-1}} l_{a,v_1,s}^{b+1} l_{a,v_1,t}^{b+1} l_{a,v_2,s}^{b+1} l_{a,v_2,t}^{b+1} \right) \\ &= E \left(\sum_{s,t=1}^{K_{b-1}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^{b+1} l_{a,v,t}^{b+1} \right)^2 \right), \end{aligned}$$

where the first equation is simply expanding the square of summations, the second equation is splitting the expression of $\ell_{a,v,s}^b$, the third equation is computing the expectation over W_b , and the final equation is changing the order of summation. Combining the above two main equations, we effectively obtain the following equation.

$$E \left(\sum_{s,t=1}^{K_{b-2}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) = K_{b-2} \prod_{\ell=b-1}^{a-1} K_{\ell} (K_{\ell} + 2) + K_{b-2} (K_{b-2} - 1) E \left(\sum_{s,t=1}^{K_{b-1}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^{b+1} l_{a,v,t}^{b+1} \right)^2 \right).$$

This holds for every $b \leq a - 1$. Now Let us define $f(b) = E \left(\sum_{s,t=1}^{K_{b-2}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) / \prod_{\ell=b-2}^{a-1} K_{\ell} (K_{\ell} - 1)$. The above equation now becomes

$$f(b) = \frac{1}{K_{b-2} - 1} \prod_{\ell=b-1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1} + f(b + 1).$$

One can easily check that $f(a + 1) = \frac{1}{K_{a-1} - 1}$ and that $f(a) = \frac{K_{a-1} + 2}{K_{a-1} - 1} \frac{1}{K_{a-2} - 1} + \frac{1}{K_{a-1} - 1}$. Therefore, by induction, we can easily obtain

$$\begin{aligned} f(b) &= \frac{1}{K_{b-2} - 1} \prod_{\ell=b-1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1} + f(b + 1) \\ &= \dots \\ &= \sum_{\phi=b-2}^{a-1} \frac{1}{K_{\phi} - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1}, \end{aligned}$$

where we use the notation $\prod_{u=i}^j = 1, i > j$ for simplicity. By plugging in back $f(b)$ to the expression of the expectation, we have

$$E \left(\sum_{s,t=1}^{K_{b-2}} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^b l_{a,v,t}^b \right)^2 \right) = \left(\prod_{\ell=b-2}^{a-1} K_{\ell} (K_{\ell} - 1) \right) \left(\sum_{\phi=b-2}^{a-1} \frac{1}{K_{\phi} - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_{\ell} + 2}{K_{\ell} - 1} \right).$$

This completes the proof. \square

C.3 Computing the Expectation

Theorem 7. If $\forall a, W_{a,p,q}^*, W_{a,p,q}, x_i$ are all i.i.d $\sim \mathcal{N}(0, 1)$, then

$$E \left(\left\| \frac{\partial f_i}{\partial W_{a,p,q}} \right\|^2 \right) = \frac{K_0 (K_0 + 2)}{K_a K_{a-1}} \left(\prod_{\ell=1}^{L-1} K_{\ell} (K_{\ell} + 2) + \prod_{\ell=1}^{L-1} K_{\ell}^2 \right).$$

Proof. We start by writing the expectation as follows.

$$\begin{aligned} E \left(\left\| \frac{\partial f_i}{\partial W_{a,p,q}} \right\|^2 \right) &= E \left((W x_i - W^* x_i) r_{a,p}(l_{a,q} x_i) \right)^2 \\ &= E \left((W x_i)^2 r_{a,p}^2(l_{a,q} x_i) \right) + E \left((W^* x_i)^2 r_{a,p}^2(l_{a,q} x_i) \right), \end{aligned}$$

where we plug in the expression of the derivative in the first equation. The second equation uses the fact that W and W^* are 0-means independent random variables.

For the first term, computing the expectation over W_a , we have

$$\begin{aligned} E_{W_a} \left((W x_i)^2 r_{a,p}^2(l_{a,q} x_i) \right) &= \sum_{s=1}^{K_a} \sum_{t=1}^{K_{a-1}} r_{a,s}^2 r_{a,p}^2(l_{a,q} x_i)^2 (l_{a,t} x_i)^2 \\ &= \sum_{s=1}^{K_a} r_{a,s}^2 r_{a,p}^2 \sum_{t=1}^{K_{a-1}} (l_{a,q} x_i)^2 (l_{a,t} x_i)^2, \end{aligned}$$

where the first equation uses the fact that W_a only appears in W where $W = W_{(a+1):L} W_a W_{(1):(a-1)}$, and all elements in W_a are i.i.d 0-means. Note that $r_{i,j}$ and $l_{k,\ell}$ are independent, so we can compute their expectation separately. By Lemma 16, we have

$$E \left(\sum_{s=1}^{K_a} r_{a,s}^2 r_{a,p}^2 \right) = (K_a + 2) F_a = (K_a + 2) \left(\prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) \right).$$

By Lemma 17, we have

$$E \left(\sum_{t=1}^{K_{a-1}} (l_{a,q} x_i)^2 (l_{a,t} x_i)^2 \right) = (K_{a-1} + 2) \left(\prod_{\ell=0}^{a-2} K_\ell (K_\ell + 2) \right).$$

Combining those two equations we have

$$E \left((W x_i)^2 r_{a,p}^2 (l_{a,q} x_i)^2 \right) = \left(\prod_{\ell=0, \ell \notin \{a, a-1\}}^{L-1} K_\ell (K_\ell + 2) \right) (K_{a-1} + 2) (K_a + 2).$$

For the second term, note that $(W^* x_i)^2$, $r_{a,p}^2$ and $(l_{a,q} x_i)^2$ are independent given x_i . Thus, we can compute the conditional expectation separately.

$$\begin{aligned} E_{W^*} \left((W^* x_i)^2 \right) &= E_{W^*} \left(\left(\prod_{t=1}^L W_t^* x_i \right)^2 \right) \\ &= E_{W^*} \left(\left(W_L^* \prod_{t=1}^{L-1} W_t^* x_i \right)^2 \right) \\ &= E_{W^*} \left(\sum_{\alpha=1}^{K_{L-1}} W_{L,:,\alpha}^{*,2} \left(W_{L-1,\alpha,:}^* \prod_{t=1}^{L-2} W_t^* x_i \right)^2 \right) \\ &= K_{L-1} E_{W^*} \left(\left(W_{L-1,\alpha,:}^* \prod_{t=1}^{L-2} W_t^* x_i \right)^2 \right) \\ &= \dots \\ &= \prod_{\ell=1}^{L-1} K_\ell \sum_{k=1}^{K_0} x_{i,k}^2, \end{aligned}$$

where the first two equations are simply plugging in the expression of W^* . The third equation uses the fact that W_a^* are i.i.d. 0-mean. The fourth equation uses the fact that $W_{L,:,\alpha}$ are symmetric. The fifth equation uses induction to finally obtain the last equation. Similarly,

$$E_{l_{a,q}} \left((l_{a,q} x_i)^2 \right) = \prod_{\ell=1}^{a-2} K_\ell \sum_{k=1}^{K_0} x_{i,k}^2,$$

and

$$E_{r_{a,p}} \left((r_{a,p})^2 \right) = \prod_{\ell=a+1}^{L-1} K_\ell.$$

Hence, the second term becomes

$$\begin{aligned} E \left((W^* x_i)^2 r_{a,p}^2 (l_{a,q} x_i)^2 \right) &= E_{x_i} \left(E_{W^*} (W^* x_i)^2 E_{r_{a,p}} (r_{a,p}^2) E_{l_{a,q}} (l_{a,q} x_i)^2 \right) \\ &= E_{x_i} \left(\prod_{\ell=1}^{L-1} K_\ell \sum_{k=1}^{K_0} x_{i,k}^2 \prod_{\ell=1}^{a-2} K_\ell \sum_{k=1}^{K_0} x_{i,k}^2 \prod_{\ell=a+1}^{L-1} K_\ell \right) \\ &= \frac{1}{K_{a-1} K_a} \prod_{\ell=1}^{L-1} K_\ell^2 E_{x_i} \left(\left(\sum_{k=1}^{K_0} x_{i,k}^2 \right)^2 \right) \\ &= \frac{1}{K_{a-1} K_a} \prod_{\ell=0}^{L-1} K_\ell^2. \end{aligned}$$

Combining both terms finishes the proof. \square

Theorem 8. If $\forall \ell, p, q, i, W_{\ell,p,q}^*, W_{\ell,p,q}, x_i$ are all i.i.d $\sim \mathcal{N}(0, 1)$, then we have

$$E(\|\nabla f_i\|^2) = L \left(K_0 (K_0 + 2) \left(\prod_{\ell=1}^{L-1} K_\ell (K_\ell + 2) + \prod_{\ell=1}^{L-1} K_\ell^2 \right) \right)$$

Proof. This can be directly obtained from the last theorem by summing over a, p, q . \square

Remarks: One can verify that when $L = 2$, this reduces to the 2-layer case and we have $E(\|\frac{\partial f_i}{\partial W_{a,p,q}}\|^2) = 2d(d+2)(2K+2)K$, which agrees with the 2-layer analysis.

Theorem 9. If $W_{\ell,p,q}, W_{\ell,p,q}^*, x_i, x_j, i \neq j$ are all i.i.d $\sim \mathcal{N}(0, 1)$, then we have

$$\begin{aligned} E\left(\frac{\partial f_i}{\partial W_{a,p,q}}\right) \left(\frac{\partial f_j}{\partial W_{a,p,q}}\right) \\ = \frac{1}{K_a K_{a-1}} \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{1}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right), \end{aligned}$$

Proof. Note that

$$\begin{aligned} E\left(\frac{\partial f_i}{\partial W_{a,p,q}}\right) \left(\frac{\partial f_j}{\partial W_{a,p,q}}\right) &= E((Wx_i - W^*x_i)(Wx_j - W^*x_j) r_{a,p}^2(l_{a,q}x_i)(l_{a,q}x_j)) \\ &= E((W^*x_i)(W^*x_j) r_{a,p}^2(l_{a,q}x_i)(l_{a,q}x_j)) \\ &\quad + E((Wx_i)(Wx_j) r_{a,p}^2(l_{a,q}x_i)(l_{a,q}x_j)), \end{aligned}$$

where we plug in the expression of the derivative into the first equation, and the second equation uses the fact that $E(WW^*) = 0$ since W, W^* are independent i.i.d. random variables.

For the first term, we have

$$\begin{aligned} &E((W^*x_i)(W^*x_h) r_{a,p}^2(l_{a,q}x_i)(l_{a,q}x_h)) \\ &= E\left(r_{a,p}^2 \sum_{s=1}^{K_0} \sum_{t=1}^{K_0} W_{2:L}^* W_{1, :, s}^* W_{2:L}^* W_{1, :, t}^* W_{a-1,q, :} W_{2:a-2} W_{1, :, s} W_{a-1,q, :} W_{2:a-2} W_{1, :, t}\right) \\ &= E\left(r_{a,p}^2 \sum_{s=1}^{K_0} (W_{2:L}^* W_{1, :, s}^*)^2 (W_{a-1,q, :} W_{2:a-2} W_{1, :, s})^2\right), \end{aligned}$$

where the first equation is because of taking expectation over x and x_i, x_j are i.i.d 0-mean, while the second equation is because we take the expectation over W_1 where again W_1 are independent and 0-mean.

Since r, W, W^* are independent, we have

$$\begin{aligned} &E\left(r_{a,p}^2 \sum_{s=1}^{K_0} (W_{2:L}^* W_{1, :, s}^*)^2 (W_{a-1,q, :} W_{2:a-2} W_{1, :, s})^2\right) \\ &= E(r_{a,p}^2) \sum_{s=1}^{K_0} E(W_{2:L}^* W_{1, :, s}^*)^2 E(W_{a-1,q, :} W_{2:a-2} W_{1, :, s})^2. \end{aligned}$$

Applying the fact that $E_x(\|a^T x\|_2^2) = \|a\|_2^2$ from Lemma 13, we have

$$\begin{aligned}
E(r_{a,p}^2) &= E\left(\left(\prod_{\ell=a+2}^L W_\ell\right) W_{a+1, :, p}\right)^2 \\
&= E\left(\left\|\prod_{\ell=a+2}^L W_\ell\right\|^2\right) \\
&= E\left(\sum_{v=1}^{K_{a+1}} \left(\prod_{\ell=a+3}^L W_\ell W_{a+2, :, v}\right)^2\right) \\
&= K_{a+1} E\left(\left(\prod_{\ell=a+3}^L W_\ell W_{a+2, :, v}\right)^2\right) \\
&= K_{a+1} E(r_{a+1, v}^2) \\
&= K_{a+1} K_{a+2} E(r_{a+2, p}^2) \\
&= \dots \\
&= \prod_{\ell=a+1}^{L-1} K_\ell.
\end{aligned}$$

Similarly, we have

$$E(W_{2:L}^* W_{1, :, s}^*)^2 = \prod_{\ell=1}^{L-1} K_\ell$$

and

$$E(W_{a-1, q, :} W_{2:a-2} W_{1, :, s})^2 = \prod_{\ell=1}^{a-2} K_\ell.$$

Hence,

$$E\left(r_{a,p}^2 \sum_{s=1}^{K_0} (W_{2:L}^* W_{1, :, s}^*)^2 (W_{a-1, q, :} W_{2:a-2} W_{1, :, s})^2\right) = K_0 \prod_{\ell=1}^{L-1} K_\ell^2 \cdot \frac{1}{K_{a-1} K_a}.$$

For the second term, we have

$$\begin{aligned}
&E((W x_i) (W x_j) r_{a,p}^2 (l_{a,q} x_i) (l_{a,q} x_j)) \\
&= E\left(\sum_{s=1}^{K_0} \sum_{t=1}^{K_0} r_{a,p}^2 W_{2:L} W_{1, :, s} W_{2:L} W_{1, :, t} l_{a,q, s} l_{a,q, t}\right) \\
&= E\left(\sum_{s, t=1}^{K_0} \sum_{u=1}^{K_a} \sum_{v=1}^{K_{a-1}} r_{a,p}^2 r_{a,u}^2 l_{a, v, s} l_{a, v, t} l_{a, q, s} l_{a, q, t}\right),
\end{aligned}$$

where we use similar tricks as in the first term, i.e., the first equation is due to taking expectation over x , and the last equation is by taking expectation over W_a . Note that r and l are independent, we can compute their expectation separately. For computation convenience, let us now take into account of summation over p, q as well. This is essentially compute the sum of the derivative over W_a instead of $W_{a,p,q}$. By Lemma 16,

$$\sum_p E\left(\sum_{u=1}^{K_a} r_{a,p}^2 r_{a,u}^2\right) = K_a (K_a + 2) \prod_{\ell=a+1}^{L-1} K_\ell (K_\ell + 2) = \prod_{\ell=a}^{L-1} K_\ell (K_\ell + 2),$$

which implies

$$E\left(\sum_{u=1}^{K_a} r_{a,p}^2 r_{a,u}^2\right) = \frac{1}{K_a} \prod_{\ell=a}^{L-1} K_\ell (K_\ell + 2).$$

Now let us consider l .

$$\sum_{q=1}^{K_{a-1}} E\left(\sum_{s, t=1}^{K_0} \sum_{v=1}^{K_{a-1}} l_{a, v, s} l_{a, v, t} l_{a, q, s} l_{a, q, t}\right) = E\left(\sum_{s, t=1}^{K_0} \left(\sum_{v=1}^{K_{a-1}} l_{a, v, s} l_{a, v, t}\right)^2\right).$$

By Lemma 19, we have

$$E \left(\sum_{s,t=1}^{K_0} \left(\sum_{v=1}^{K_{a-1}} l_{a,v,s}^0 l_{a,v,t}^0 \right)^2 \right) = \left(\prod_{\ell=0}^{a-1} K_\ell (K_\ell - 1) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_\ell + 2}{K_\ell - 1} \right),$$

which implies

$$E \left(\sum_{s,t=1}^{K_0} \sum_{v=1}^{K_{a-1}} l_{a,v,s} l_{a,v,t} l_{a,q,s} l_{a,q,t} \right) = \frac{1}{K_{a-1}} \left(\prod_{\ell=0}^{a-1} K_\ell (K_\ell - 1) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_\ell + 2}{K_\ell - 1} \right),$$

Combing those two terms, we have

$$\begin{aligned} & E \left((W x_i) (W x_j) r_{a,p}^2 (l_{a,q} x_i) (l_{a,q} x_j) \right) \\ &= E \left(\sum_{s,t=1}^{K_0} \sum_{u=1}^{K_a} \sum_{v=1}^{K_{a-1}} r_{a,p}^2 r_{a,u}^2 l_{a,v,s} l_{a,v,t} l_{a,q,s} l_{a,q,t} \right) \\ &= \left(\frac{1}{K_a} \prod_{\ell=a}^{L-1} K_\ell (K_\ell + 2) \right) \left(\frac{1}{K_{a-1}} \left(\prod_{\ell=0}^{a-1} K_\ell (K_\ell - 1) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_\ell + 2}{K_\ell - 1} \right) \right) \\ &= \left(\frac{1}{K_a K_{a-1}} \prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\left(\prod_{\ell=0}^{a-1} \frac{K_\ell - 1}{K_\ell + 2} \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=\phi+1}^{a-1} \frac{K_\ell + 2}{K_\ell - 1} \right) \right) \\ &= \left(\frac{1}{K_a K_{a-1}} \prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} \right). \end{aligned}$$

Summing the two terms from the original expression, we finally have

$$\begin{aligned} & E \left(\frac{\partial f_i}{\partial W_{a,p,q}} \right) \left(\frac{\partial f_j}{\partial W_{a,p,q}} \right) \\ &= E \left((W^* x_i) (W^* x_j) r_{a,p}^2 (l_{a,q} x_i) (l_{a,q} x_j) \right) \\ &+ E \left((W x_i) (W x_j) r_{a,p}^2 (l_{a,q} x_i) (l_{a,q} x_j) \right) \\ &= K_0 \prod_{\ell=1}^{L-1} K_\ell^2 \cdot \frac{1}{K_{a-1} K_a} + \left(\frac{1}{K_a K_{a-1}} \prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} \right) \\ &= \frac{1}{K_a K_{a-1}} \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{1}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right), \end{aligned}$$

which completes the proof. \square

Theorem 10. If $W_{\ell,p,q}, x_i, x_j, i \neq j$ are all i.i.d $\sim \mathcal{N}(0, 1)$, then we have

$$E \left(\langle \nabla f_i, \nabla f_j \rangle \right) = \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{L - \phi}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell + 2}{K_\ell - 1} + \frac{L}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right).$$

Proof. From Theorem 9, we have

$$\begin{aligned} & E \left(\frac{\partial f_i}{\partial W_{a,p,q}} \right) \left(\frac{\partial f_j}{\partial W_{a,p,q}} \right) \\ &= \frac{1}{K_a K_{a-1}} \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{1}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right). \end{aligned}$$

Summing over p, q , we have

$$\begin{aligned} & \sum_{p=1}^{K_a} \sum_{q=1}^{K_{a-1}} E \left(\frac{\partial f_i}{\partial W_{a,p,q}} \right) \left(\frac{\partial f_j}{\partial W_{a,p,q}} \right) \\ &= \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{1}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right). \end{aligned}$$

Thus, we have

$$\begin{aligned}
E(\langle \nabla f_i, \nabla f_j \rangle) &= \sum_{a=1}^L \sum_{p=1}^{K_a} \sum_{q=1}^{K_{a-1}} E\left(\frac{\partial f_i}{\partial W_{a,p,q}}\right) \left(\frac{\partial f_j}{\partial W_{a,p,q}}\right) \\
&= \sum_{a=1}^L \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{1}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right) \\
&= \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{a=1}^L \sum_{\phi=0}^{a-1} \frac{1}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{L}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right) \\
&= \left(\prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2) \right) \left(\sum_{\phi=0}^{a-1} \frac{L - \phi}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{L}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right).
\end{aligned}$$

□

Finally we arrive at the main theorem.

Theorem 6. Consider a LNN with $L \geq 2$ layers. Let the weight values $W_{l,p,q}$ for $l \in \{1, \dots, L\}$ and \mathbf{x}_i be independently drawn random variables from $\mathcal{N}(0, 1)$. Let

$$M = n^2 \prod_{\ell=0}^{L-1} K_\ell (K_\ell + 2)$$

Then:

$$\begin{aligned}
\mathbb{E}[n \sum_{i=1}^n \|\nabla f_i\|^2] &= M \cdot L \left(1 + \prod_{\ell=1}^{L-1} \frac{K_\ell}{K_\ell + 2} \right), \\
\mathbb{E} \left[\sum_{i=1, j \neq i}^n \langle \nabla f_i, \nabla f_j \rangle \right] &= M \cdot \frac{n-1}{n} \left(\sum_{\phi=0}^{L-1} \frac{L - \phi}{K_\phi - 1} \prod_{\ell=0}^{\phi} \frac{K_\ell - 1}{K_\ell + 2} + \frac{L}{K_0} \prod_{\ell=0}^{L-1} \frac{K_\ell}{K_\ell + 2} \right).
\end{aligned}$$

Proof. This can be directly achieved from Theorem 8 and Theorem 10.

□