# <span id="page-0-14"></span><span id="page-0-5"></span><span id="page-0-3"></span>Supplementary Material

# <span id="page-0-11"></span><span id="page-0-10"></span><span id="page-0-9"></span><span id="page-0-2"></span>Adversarially Robust Optimization with Gaussian Processes

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# <span id="page-0-1"></span>A Illustration of STABLEOPT's Execution

The following figure gives an example of the selection procedure of STABLEOPT at two different time steps:

<span id="page-0-16"></span>

<span id="page-0-8"></span>Figure 4: An execution of STABLEOPT on the running example. We observe that after  $t = 15$  steps,  $\tilde{\mathbf{x}}_t$  obtained in Eq.  $\boxed{13}$  corresponds to  $\mathbf{x}_{\epsilon}^*$ .

<span id="page-0-15"></span><span id="page-0-13"></span><span id="page-0-12"></span><span id="page-0-4"></span>The intermediate time steps are illustrated as follows:

<span id="page-0-17"></span><span id="page-0-7"></span><span id="page-0-6"></span>

### B Proofs of Theoretical Results

#### **B.1** Proof of Theorem  $\boxed{1}$  (upper bound)

Recall that  $\tilde{\mathbf{x}}_t$  is the point computed by STABLEOPT in  $\sqrt{13}$  at time *t*, and that  $\delta_t$  corresponds to the perturbation obtained in STABLEOPT (Line 3) at time *t*. In the following, we condition on the event in Lemma [1](#page-0-2) holding true, meaning that  $ucb_t$  and  $lc_b$  provide valid confidence bounds as per  $(15)$ . As stated in the lemma, this holds with probability at least  $1 - \xi$ .

By the definition of  $\epsilon$ -instant regret, we have

$$
r_{\epsilon}(\tilde{\mathbf{x}}_t) = \max_{\mathbf{x} \in D} \min_{\boldsymbol{\delta} \in \Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x} + \boldsymbol{\delta}) - \min_{\boldsymbol{\delta} \in \Delta_{\epsilon}(x_t)} f(\tilde{\mathbf{x}}_t + \boldsymbol{\delta})
$$
(32)

<span id="page-1-0"></span>
$$
\leq \max_{\mathbf{x}\in D} \min_{\boldsymbol{\delta}\in\Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x}+\boldsymbol{\delta}) - \min_{\boldsymbol{\delta}\in\Delta_{\epsilon}(\tilde{\mathbf{x}}_t)} \mathrm{lcb}_{t-1}(\tilde{\mathbf{x}}_t+\boldsymbol{\delta})
$$
(33)

<span id="page-1-2"></span>
$$
= \max_{\mathbf{x} \in D} \min_{\delta \in \Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x} + \delta) - \mathrm{lcb}_{t-1}(\tilde{\mathbf{x}}_t + \delta_t)
$$
(34)

<span id="page-1-1"></span>
$$
\leq \max_{\mathbf{x} \in D} \min_{\boldsymbol{\delta} \in \Delta_{\epsilon}(\mathbf{x})} \mathrm{ucb}_{t-1}(\mathbf{x} + \boldsymbol{\delta}) - \mathrm{lcb}_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t)
$$
(35)

<span id="page-1-4"></span><span id="page-1-3"></span>
$$
= \min_{\delta \in \Delta_{\epsilon}(\tilde{\mathbf{x}}_t)} \text{ucb}_{t-1}(\tilde{\mathbf{x}}_t + \delta) - \text{lcb}_{t-1}(\tilde{\mathbf{x}}_t + \delta_t)
$$
(36)

$$
\leq \text{ucb}_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t) - \text{lcb}_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t)
$$
\n(37)

<span id="page-1-5"></span>
$$
=2\beta_t^{1/2}\sigma_{t-1}(\tilde{\mathbf{x}}_t+\boldsymbol{\delta}_t),\tag{38}
$$

where  $(33)$  and  $(35)$  follow from Lemma  $\overline{1}$ ,  $(34)$  follows since  $\delta_t$  minimizes lcb<sub>t-1</sub> by definition, [\(36\)](#page-1-3) follows since  $\tilde{\mathbf{x}}_t$  maximizes the robust upper confidence bound by definition, [\(37\)](#page-1-4) follows by upper bounding the minimum by the specific choice  $\delta_t \in \Delta_{\epsilon}(\mathbf{x}_t)$ , and [\(38\)](#page-1-5) follows since the upper and lower confidence bounds are separated by  $2\beta_t^{1/2} \sigma_{t-1}(\cdot)$  according to their definitions in [\(12\)](#page-0-4).

In fact, the analysis from  $(33)$  to  $(38)$  shows that the following *pessimistic estimate* of  $r_{\epsilon}(\tilde{\mathbf{x}}_t)$  is upper bounded by  $2\beta_t^{1/2} \sigma_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t)$ :

<span id="page-1-6"></span>
$$
\overline{r}_{\epsilon}(\tilde{\mathbf{x}}_t) = \max_{\mathbf{x} \in D} \min_{\delta \in \Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x} + \delta) - \min_{\delta \in \Delta_{\epsilon}(\tilde{\mathbf{x}}_t)} \text{lcb}_{t-1}(\tilde{\mathbf{x}}_t + \delta). \tag{39}
$$

Unlike  $r_{\epsilon}(\tilde{\mathbf{x}}_t)$ , the algorithm has the required knowledge to identify the value of  $t \in \{1, \ldots, T\}$  with the smallest  $\overline{r}_{\epsilon}(\tilde{\mathbf{x}}_t)$ . Specifically, the first term on the right-hand side of [\(39\)](#page-1-6) does not depend on *t*, so the smallest  $\overline{r}_{\epsilon}(\tilde{\mathbf{x}}_t)$  is achieved by  $\mathbf{x}^{(T)}$  defined in [\(17\)](#page-0-5). Since the minimum is upper bounded by the average, it follows that

$$
r_{\epsilon}(\mathbf{x}^{(T)}) \le \overline{r}_{\epsilon}(\mathbf{x}^{(T)})
$$
\n<sup>(40)</sup>

$$
\leq \frac{1}{T} \sum_{t=1}^{T} 2\beta_t^{1/2} \sigma_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t)
$$
\n(41)

<span id="page-1-10"></span><span id="page-1-8"></span><span id="page-1-7"></span>
$$
\leq \frac{2\beta_T^{1/2}}{T} \sum_{t=1}^T \sigma_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t), \tag{42}
$$

where [\(41\)](#page-1-7) uses [\(38\)](#page-1-5), and [\(42\)](#page-1-8) uses the monotonicity of  $\beta_T$ . Next, we claim that

$$
2\sum_{t=1}^{T} \sigma_{t-1}(\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t) \le \sqrt{C_1 T \gamma_T},
$$
\n(43)

where  $C_1 = 8/\log(1 + \sigma^{-2})$ . In fact, this is a special case of the well-known result [31], Lemma 5.[4](#page-1-9)],<sup>4</sup> which upper bounds the sum of posterior standard deviations of sampled points in terms of the information gain  $\gamma_T$  (recall that STABLEOPT samples at location  $\tilde{\mathbf{x}}_t + \boldsymbol{\delta}_t$ ). Combining [\(42\)](#page-1-8)–[\(43\)](#page-1-10) and re-arranging, we deduce that after *T* satisfies  $\frac{T}{\beta_T \gamma_T} \ge \frac{C_1}{\eta^2}$ , the  $\epsilon$ -instant regret is at most  $\eta$ , thus completing the proof.

<span id="page-1-9"></span><sup>&</sup>lt;sup>4</sup>More precisely,  $\left[31\right]$  Lemma 5.4] alongside an application of the Cauchy-Schwarz inequality as in  $\left[31\right]$ .

<span id="page-2-0"></span>![](_page_2_Figure_0.jpeg)

Figure 6: Illustration of functions  $f_1, \ldots, f_5$  equal to a common function shifted by various multiples of a given parameter  $w$ . In the  $\epsilon$ -stable setting, there is a wide region (shown in gray for the dark blue curve  $f_3$ ) within which the perturbed function value equals  $-2\eta$ .

#### **B.2** Proof of Theorem  $2$  (lower bound)

Our lower bounding analysis builds heavily on that of the non-robust optimization setting with  $f \in \mathcal{F}_k(B)$  studied in  $[27]$ , but with important differences. Roughly speaking, the analysis of  $[27]$  is based on the difficulty of finding a very narrow "bump" of height  $2\eta$  in a function whose values are mostly close to zero. In the  $\epsilon$ -stable setting, however, even the points around such a bump will be adversarially perturbed to another point whose function value is nearly zero. Hence, all points are essentially equally bad.

To overcome this challenge, we consider the reverse scenario: Most of the function values are still nearly zero, but there exists a narrow *valley* of depth  $-2\eta$ . This means that every point within an  $\epsilon$ -ball around the function minimizer will be perturbed to the point with value  $-2\eta$ . Hence, a constant fraction of the volume is still  $2\eta$ -suboptimal, and it is impossible to avoid this region with high probability unless the time horizon *T* is sufficiently large. An illustration is given in Figure  $\overline{6}$ , with further details below.

We now proceed with the formal proof.

#### B.2.1 Preliminaries

Recall that we are considering an arbitrary given (deterministic) GP optimization algorithm. More precisely, such an algorithm consists of a sequence of decision functions that return a sampling location  $x_t$  based on  $y_1, \ldots, y_{t-1}$ , and an additional decision function that reports the final point  $x^{(T)}$ based on  $y_1, \ldots, y_T$ . The points  $x_1, \ldots, x_{t-1}$  (or  $x_1, \ldots, x_T$ ) do not need to be treated as additional inputs to these functions, since  $(\mathbf{x}_1, \ldots, \mathbf{x}_{t-1})$  is a deterministic function of  $(y_1, \ldots, y_{t-1})$ .

We first review several useful results and techniques from  $[27]$ :

- We lower bound the worst-case  $\epsilon$ -regret within  $\mathcal{F}_k(B)$  by the  $\epsilon$ -regret averaged over a suitablydesigned finite collection  $\{f_1, \ldots, f_M\} \subset \mathcal{F}_k(B)$  of size *M*.
- We choose each  $f_m(\mathbf{x})$  to be a shifted version of a common function  $g(\mathbf{x})$  on  $\mathbb{R}^p$ . Specifically, each  $f_m(\mathbf{x})$  is obtained by shifting  $g(\mathbf{x})$  by a different amount, and then cropping to  $D = [0, 1]^p$ . For our purposes, we require  $g(x)$  to satisfy the following properties:
	- 1. The RKHS norm in  $\mathbb{R}^p$  is bounded,  $||g||_k \leq B$ ;
	- 2. We have (i)  $g(x) \in [-2\eta, 2\eta]$  with minimum value  $g(0) = -2\eta$ , and (ii) there is a "width" *w* such that  $g(\mathbf{x}) > -\eta$  for all  $\|\mathbf{x}\|_{\infty} \geq w$ ;
	- 3. There are absolute constants  $h_0 > 0$  and  $\zeta > 0$  such that  $g(\mathbf{x}) = \frac{2\eta}{h_0} h\left(\frac{\mathbf{x}\zeta}{w}\right)$  for some function  $h(\mathbf{z})$  that decays faster than any finite power of  $\|\mathbf{z}\|_2^{-1}$  as  $\|\mathbf{z}\|_2 \to \infty$ .

Letting  $g(x)$  be such a function, we construct the *M* functions by shifting  $g(x)$  so that each  $f_m(\mathbf{x})$  is centered on a unique point in a uniform grid, with points separated by *w* in each dimension. Since  $D = [0, 1]^p$ , one can construct

<span id="page-3-2"></span><span id="page-3-1"></span>
$$
M = \left\lfloor \left(\frac{1}{w}\right)^p \right\rfloor \tag{44}
$$

such functions. We will use this construction with  $w \ll 1$ , so that there is no risk of having  $M = 0$ , and in fact M can be assumed larger than any desired absolute constant.

• It is shown in  $\left[\frac{27}{1}\right]$  that the above properties<sup>5</sup> can be achieved with

$$
M = \left\lfloor \left( \frac{r\sqrt{\log \frac{B(2\pi l^2)^{p/4}h(0)}{2\eta}}}{\zeta \pi l} \right)^p \right\rfloor \tag{45}
$$

in the case of the SE kernel, and with

<span id="page-3-3"></span>
$$
M = \left\lfloor \left(\frac{Bc_3}{\eta}\right)^{p/\nu} \right\rfloor \tag{46}
$$

in the case of the Matérn kernel, where

$$
c_3 := \left(\frac{r}{\zeta}\right)^{\nu} \cdot \left(\frac{c_2^{-1/2}}{2(8\pi^2)^{(\nu+p/2)/2}}\right),\tag{47}
$$

and where  $c_2 > 0$  is an absolute constant. Note that these values of M amount to choosing w in  $\frac{\sqrt{44}}{4}$ , and the assumption of sufficiently small  $\frac{\eta}{B}$  in the theorem statement ensures that  $M \gg 1$ (or equivalently  $w \ll 1$ ) as stated above.

• Property 2 above ensures that the "robust" function value  $\min_{\delta \in \Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x})$  equals  $-2\eta$  for any x whose  $\epsilon$ -neighborhood includes the minimizer  $x_{\min}$  of f, while being  $-\eta$  or higher for any input whose entire  $\epsilon$ -neighborhood is separated from  $\mathbf{x}_{\min}$  by at least *w*. For  $w \ll 1$  and  $\epsilon$  < 0.5, a point of the latter type is guaranteed to exist, which implies

<span id="page-3-4"></span>
$$
r_{\epsilon}(\mathbf{x}) \ge \eta \tag{48}
$$

for any x whose  $\epsilon$ -neighborhood includes  $x_{\min}$ .

In addition, we introduce the following notation, also used in [\[27\]](#page-0-8):

- The probability density function of the output sequence  $y = (y_1, \ldots, y_T)$  when the underlying function is  $f_m$  is denoted by  $P_m(y)$ . We also define  $f_0(x)=0$  to be the zero function, and define  $P_0(y)$  analogously for the case that the optimization algorithm is run on *f*0. Expectations and probabilities (with respect to the noisy observations) are similarly written as  $\mathbb{E}_m$ ,  $\mathbb{P}_m$ ,  $\mathbb{E}_0$ , and  $\mathbb{P}_0$  when the underlying function is  $f_m$  or  $f_0$ . On the other hand, in the absence of a subscript,  $\mathbb{E}[\cdot]$  and  $\mathbb{P}[\cdot]$  are taken with respect to the noisy observations *and* the random function *f* drawn uniformly from  $\{f_1, \ldots, f_M\}$  (recall that we are lower bounding the worst case by this average).
- Let  $\{R_m\}_{m=1}^M$  be a partition of the domain into *M* regions according the above-mentioned uniform grid, with  $f_m$  taking its minimum value of  $-2\eta$  in the centre of  $\mathcal{R}_m$ . Moreover, let *j*<sup>*t*</sup> be the index at time *t* such that  $x_t$  falls into  $\mathcal{R}_j$ <sup>*t*</sup>; this can be thought of as a quantization of  $\mathbf{x}_t$ .
- Define the maximum (absolute) function value within a given region  $\mathcal{R}_i$  as

$$
\overline{v}_m^j := \max_{\mathbf{x} \in \mathcal{R}_j} |f_m(\mathbf{x})|,\tag{49}
$$

and the maximum KL divergence to  $P_0$  within the region as

$$
\overline{D}_m^j := \max_{\mathbf{x} \in \mathcal{R}_j} D(P_0(\cdot|\mathbf{x}) \| P_m(\cdot|\mathbf{x})),\tag{50}
$$

where  $P_m(y|\mathbf{x})$  is the distribution of an observation *y* for a given selected point **x** under the function  $f_m$ , and similarly for  $P_0(y|\mathbf{x})$ .

<span id="page-3-0"></span><sup>&</sup>lt;sup>5</sup>Here  $g(x)$  plays the role of  $-g(x)$  in  $[27]$  due to the discussion at the start of this appendix, but otherwise the construction is identical.

• Let  $N_j \in \{0, \ldots, T\}$  be a random variable representing the number of points from  $\mathcal{R}_j$  that are selected throughout the *T* rounds.

Next, we present several useful lemmas. The following well-known change-of-measure result, which can be viewed as a form of Le Cam's method, has been used extensively in both discrete and continuous bandit problems.

<span id="page-4-2"></span>**Lemma 2.**  $[\mathbf{I}, \mathbf{p}, 27]$  *For any function*  $a(\mathbf{y})$  *taking values in a bounded range* [0, A]*, we have* 

$$
\left|\mathbb{E}_{m}[a(\mathbf{y})] - \mathbb{E}_{0}[a(\mathbf{y})]\right| \le A d_{\text{TV}}(P_0, P_m)
$$
\n(51)

<span id="page-4-0"></span>
$$
\leq A \sqrt{D(P_0 \| P_m)},\tag{52}
$$

where  $d_{\text{TV}}(P_0, P_m) = \frac{1}{2} \int_{\mathbb{R}^T} |P_0(\mathbf{y}) - P_m(\mathbf{y})| \, d\mathbf{y}$  *is the total variation distance.* 

We briefly remark on some slight differences here compared to [\[1,](#page-0-9) p. 27]. There, only  $\mathbb{E}_m[a(\mathbf{y})]$  –  $\mathbb{E}_{0}[a(\mathbf{y})]$  is upper bounded in terms of  $d_{\text{TV}}(P_0, P_m)$ , but one easily obtains the same upper bound on  $\mathbb{E}_{0}[a(\mathbf{y})] - \mathbb{E}_{m}[a(\mathbf{y})]$  by interchanging the roles of  $P_0$  and  $P_m$ . The step [\(52\)](#page-4-0) follows from Pinsker's inequality,  $d_{\text{TV}}(P_0, P_m) \le \sqrt{\frac{D(P_0||P_m)}{2}}$ , and by upper bounding  $\frac{1}{\sqrt{2}} \le 1$  to ease the notation.

The following result simplifies the divergence term in  $(52)$ .

<span id="page-4-4"></span>Lemma 3. [\[27,](#page-0-8) Eq. (44)] *Under the preceding definitions, we have*

$$
D(P_0 \| P_m) \le \sum_{j=1}^{M} \mathbb{E}_0[N_j] \overline{D}_m^j.
$$
\n
$$
(53)
$$

The following well-known property gives a formula for the KL divergence between two Gaussians. **Lemma 4.**  $[27, Eq. (36)]$  $[27, Eq. (36)]$  *For*  $P_1$  *and*  $P_2$  *being Gaussian with means*  $(\mu_1, \mu_2)$  *and a common variance*  $\sigma^2$ , we have

<span id="page-4-6"></span>
$$
D(P_1 \| P_2) = \frac{(\mu_1 - \mu_2)^2}{2\sigma^2}.
$$
\n(54)

Finally, we have the following technical result regarding the "needle-in-haystack" type function constructed above.

<span id="page-4-7"></span>**Lemma 5.** [\[27,](#page-0-8) Lemma 7] *The functions*  $\{f_m\}_{m=1}^M$  *corresponding to* [\(45\)](#page-3-2)–[\(46\)](#page-3-3) *are such that the quantities*  $\overline{v}_{m}^{j}$  *satisfy*  $\sum_{m=1}^{M} (\overline{v}_{m}^{j})^{2} = O(\eta^{2})$  *for all j.* 

#### B.2.2 Analysis of the average  $\epsilon$ -stable regret

Let  $J_{bad}(m)$  be the set of *j* such that all  $\mathbf{x} \in \mathcal{R}_j$  yield  $\min_{\delta \in \Delta_{\epsilon}(\mathbf{x})} f(\mathbf{x} + \delta) = -2\eta$  when the true function is  $f_m$ , and define  $\mathcal{R}_{bad}(m) = \bigcup_{j \in J_{bad}(m)} \mathcal{R}_j$ . By the  $\epsilon$ -regret lower bound in [\(48\)](#page-3-4), we have

$$
\mathbb{E}_{m}[r_{\epsilon}(\mathbf{x}^{(T)})] \geq \eta \mathbb{P}_{m}[\mathbf{x}^{(T)} \in \mathcal{R}_{\text{bad}}(m)]
$$
\n(55)

$$
\geq \eta \bigg( \mathbb{P}_0[\mathbf{x}^{(T)} \in \mathcal{R}_{\text{bad}}(m)] - \sqrt{D(P_0 \| P_m)} \bigg) \tag{56}
$$

<span id="page-4-3"></span><span id="page-4-1"></span>
$$
\geq \eta \bigg( \mathbb{P}_0[\mathbf{x}^{(T)} \in \mathcal{R}_{\text{bad}}(m)] - \sqrt{\sum_{j=1}^M \mathbb{E}_0[N_j] \overline{D}_m^j} \bigg),\tag{57}
$$

where [\(56\)](#page-4-1) follows from Lemma  $2$  with  $a(y) = 1\{x^{(T)} \in \mathcal{R}_{bad}(m)\}\$  and  $A = 1$  (recall that  $x^{(T)}$ ) is a function of  $y = (y_1, \ldots, y_T)$ , and  $(57)$  follows from Lemma [3.](#page-4-4) Averaging over *m* uniform on *{*1*,...,M}*, we obtain

<span id="page-4-5"></span>
$$
\mathbb{E}[r_{\epsilon}(\mathbf{x}^{(T)})] \ge \eta \left(\frac{1}{M} \sum_{m=1}^{M} \mathbb{P}_{0}[\mathbf{x}^{(T)} \in \mathcal{R}_{\text{bad}}(m)] - \frac{1}{M} \sum_{m=1}^{M} \sqrt{\sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}]\overline{D}_{m}^{j}}\right).
$$
(58)

We proceed by bounding the two terms separately.

*•* We first claim that

<span id="page-5-0"></span>
$$
\frac{1}{M} \sum_{m=1}^{M} \mathbb{P}_0[\mathbf{x}^{(T)} \in \mathcal{R}_{\text{bad}}(m)] \ge C_1
$$
\n(59)

for some  $C_1 > 0$ . To show this, it suffices to prove that any given  $\mathbf{x}^{(T)} \in D$  is in at least a constant fraction of the  $\mathcal{R}_{bad}(m)$  regions, of which there are  $\tilde{M}$ . This follows from the fact that the  $\epsilon$ -ball centered at  $\mathbf{x}_{m,\min} = \arg \min_{\mathbf{x} \in D} f_m(\mathbf{x})$  takes up a constant fraction of the volume of *D*, where the constant depends on both the stability parameter  $\epsilon$  and the dimension *p*. A small caveat is that because the definition of  $\mathcal{R}_{bad}$  insists that the *every* point in the region  $\mathcal{R}_j$ is within distance  $\epsilon$  of  $\mathbf{x}_{m,\text{min}}$ , the left-hand side of  $(59)$  may be slightly below the relevant ratio of volumes above. However, since Theorem  $\frac{2}{3}$  assumes that  $\frac{\eta}{B}$  is sufficiently small, the choices of *M* in  $(45)$  and  $(46)$  ensure that *M* is sufficiently large for this "quantization" effect to be negligible.

• For the second term in [\(58\)](#page-4-5), we claim that

<span id="page-5-5"></span>
$$
\frac{1}{M} \sum_{m=1}^{M} \sqrt{\sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}]\overline{D}_{m}^{j}} \leq C_{2} \frac{\eta}{\sigma} \sqrt{\frac{T}{M}}
$$
(60)

for some  $C_2 > 0$ . To see this, we write

$$
\frac{1}{M} \sum_{m=1}^{M} \sqrt{\sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}]\overline{D}_{m}^{j}}
$$
\n
$$
= O\left(\frac{1}{\sigma}\right) \cdot \frac{1}{M} \sum_{m=1}^{M} \sqrt{\sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}](\overline{v}_{m}^{j})^{2}}
$$
\n(61)

<span id="page-5-2"></span><span id="page-5-1"></span>
$$
\leq O\left(\frac{1}{\sigma}\right) \cdot \sqrt{\frac{1}{M} \sum_{m=1}^{M} \sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}](\overline{v}_{m}^{j})^{2}}
$$
(62)

$$
= O\left(\frac{1}{\sigma}\right) \cdot \sqrt{\frac{1}{M} \sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}]\left(\sum_{m=1}^{M} (\overline{v}_{m}^{j})^{2}\right)}
$$
(63)

<span id="page-5-3"></span>
$$
= O\left(\frac{\eta}{\sqrt{M}\sigma}\right) \cdot \sqrt{\sum_{j=1}^{M} \mathbb{E}_{0}[N_{j}]}
$$
\n(64)

<span id="page-5-4"></span>
$$
=O\left(\frac{\sqrt{T}\eta}{\sqrt{M}\sigma}\right),\tag{65}
$$

where [\(61\)](#page-5-1) follows since the divergence  $D(P_0(\cdot|\mathbf{x})||P_m(\cdot|\mathbf{x}))$  associated with a point x having value  $v(\mathbf{x})$  is  $\frac{v(\mathbf{x})^2}{2\sigma^2}$  (*cf.*, [\(54\)](#page-4-6)), [\(62\)](#page-5-2) follows from Jensen's inequality, [\(64\)](#page-5-3) follows from Lemma  $\overline{5}$ , and [\(65\)](#page-5-4) follows from  $\sum_j N_j = T$ .

Substituting  $(59)$  and  $(60)$  into  $(58)$ , we obtain

$$
\mathbb{E}[r_{\epsilon}(\mathbf{x}^{(T)})] \ge \eta \Big( C_1 - C_2 \frac{\eta}{\sigma} \sqrt{\frac{T}{M}} \Big),\tag{66}
$$

which implies that the regret is lower bounded by  $\Omega(\eta)$  unless  $T = \Omega(\frac{M\sigma^2}{\eta^2})$ . Substituting M from  $\overline{45}$  and  $\overline{46}$ , we deduce that the conditions on *T* in the theorem statement are necessary to achieve average regret  $\mathbb{E}[r_{\epsilon}(\mathbf{x}^{(T)})] = O(\eta)$  with a sufficiently small implied constant.

#### B.2.3 From average to high-probability regret

Recall that we are considering functions whose values lie in the range  $[-2\eta, 2\eta]$ , implying that  $r_{\epsilon}(\mathbf{x}^{(T)}) \leq 4\eta$ . Letting  $T_{\eta}$  be the lower bound on *T* derived above for achieving average regret  $O(\eta)$  (i.e., we have  $\mathbb{E}[r^{(T_{\eta})}_{\epsilon}]=\Omega(\eta)$ ), it follows from the reverse Markov inequality (i.e., Markov's inequality applied to the random variable  $4\eta - r_{\epsilon}^{(T_{\eta})}$ ) that

$$
\mathbb{P}[r_{\epsilon}(\mathbf{x}^{(T_{\eta})}) \geq c\eta] \geq \frac{\Omega(\eta) - c\eta}{4\eta - c\eta} \tag{67}
$$

for any *c >* 0 sufficiently small for the numerator and denominator to be positive. The right-hand side is lower bounded by a constant for any such  $c$ , implying that the probability of achieving  $\epsilon$ -regret at most  $c\eta$  cannot be arbitrarily close to one. By renaming  $c\eta$  as  $\eta'$ , it follows that in order to achieve some target  $\epsilon$ -stable regret  $\eta'$  with probability sufficiently close to one, a lower bound of the same form as the average regret bound holds. In other words, the conditions on *T* in the theorem statement remain necessary also for the high-probability regret.

We emphasize that Theorem  $\sqrt{2}$  concerns the high-probability regret when "high probability" means *sufficiently close to one* as a function of  $\epsilon$ ,  $p$ , and the kernel parameters (but still constant with respect to  $T$  and  $\eta$ ). We do not claim a lower bound under any particular *given* success probability (e.g.,  $\eta$ -optimality with probability at least  $\frac{3}{4}$ ).

#### C Details on Variations from Section [4](#page-0-10)

We claim that the STABLEOPT variations and theoretical results outlined in Section  $\overline{A}$  are in fact special cases of Algorithm  $\prod$  and Theorem  $\prod$ , despite being seemingly quite different. The idea behind this claim is that Algorithm  $\prod$  and Theorem  $\prod$  allow for the "distance" function  $d(\cdot, \cdot)$  to be completely arbitrary, so we may choose it in rather creative/unconventional ways.

In more detail, we have the following:

• For the unknown parameter setting  $\max_{\mathbf{x} \in D} \min_{\theta \in \Theta} f(\mathbf{x}, \theta)$ , we replace x in the original setting by the concatenated input  $(x, \theta)$ , and set

$$
d((\mathbf{x}, \boldsymbol{\theta}), (\mathbf{x}', \boldsymbol{\theta}')) = \|\mathbf{x} - \mathbf{x}'\|_2.
$$
 (68)

If we then set  $\epsilon = 0$ , we find that the input x experiences no perturbation, whereas  $\theta$  may be perturbed arbitrarily, thereby reducing  $\overline{Q}$  to  $\max_{\mathbf{x}\in D} \min_{\theta\in\Theta} f(\mathbf{x}, \theta)$  as desired.

• For the robust estimation setting, we again use the concatenated input  $(x, \theta)$ . To avoid overloading notation, we let  $d_0(\bm{\theta},\bm{\theta}')$  denote the distance function (applied to  $\bm{\theta}$  alone) adopted for this case in Section  $\overline{4}$ . We set

$$
d((\mathbf{x}, \boldsymbol{\theta}), (\mathbf{x}', \boldsymbol{\theta}')) = \begin{cases} d_0(\boldsymbol{\theta}, \boldsymbol{\theta}') & \mathbf{x} = \mathbf{x}' \\ \infty & \mathbf{x} \neq \mathbf{x}'. \end{cases}
$$
(69)

Due to the second case, the input x experiences no perturbation, since doing so would violate the distance constraint of  $\epsilon$ . We are then left with  $\mathbf{x} = \mathbf{x}'$  and  $d_0(\theta, \theta') \leq \epsilon$ , as required.

• For the grouped setting  $\max_{G \in \mathcal{G}} \min_{\mathbf{x} \in G} f(\mathbf{x})$ , we adopt the function

$$
d(\mathbf{x}, \mathbf{x}') = \mathbf{1}\{\mathbf{x} \text{ and } \mathbf{x}' \text{ are in different groups}\},\tag{70}
$$

and set  $\epsilon = 0$ . Considering the formulation in  $\overline{D}$ , we find that any two inputs x and x' yield the same  $\epsilon$ -stable objective function, and hence, reporting a point x is equivalent to reporting its group *G*. As a result, [\(7\)](#page-0-12) reduces to the desired formulation  $\max_{G \in \mathcal{G}} \min_{\mathbf{x} \in G} f(\mathbf{x})$ .

The variations of STABLEOPT described in  $(20)$ – $(26)$ , as well as the corresponding theoretical results outlined in Section  $\overline{A}$ , follow immediately by substituting the respective choices of  $d(\cdot, \cdot)$  and  $\epsilon$  above into Algorithm  $\Pi$  and Theorem  $\Pi$ . It should be noted that in the first two examples, the definition of  $\gamma_t$ in [\(14\)](#page-0-15) is modified to take the maximum over not only  $x_1, \dots, x_t$ , but also  $\theta_1, \dots, \theta_t$ .

## D Lake Data Experiment

We consider an application regarding environmental monitoring of inland waters, using a data set containing 2024 in situ measurements of chlorophyll concentration within a vertical transect plane, collected by an autonomous surface vessel in Lake Zürich. This data set was considered in previous

<span id="page-7-0"></span>![](_page_7_Figure_0.jpeg)

Figure 7: Experiment on the Zürich lake dataset; In the later rounds STABLEOPT is the only method that reports a near-optimal  $\epsilon$ -stable point.

works such as  $[7, 15]$  $[7, 15]$  to detect regions of high concentration. In these works, the goal was to locate all regions whose concentration exceeds a pre-defined threshold.

Here we consider a different goal: We seek to locate a region of a given size such that the concentration throughout the region is as high as possible (in the max-min sense). This is of interest in cases where high concentration only becomes relevant when it is spread across a sufficiently wide area. We consider rectangular regions with different pre-specified lengths in each dimension:

$$
\Delta_{\epsilon_D,\epsilon_L}(\mathbf{x}) = \{\mathbf{x}' - \mathbf{x} : \mathbf{x}' \in D, \ |x_D - x'_D| \le \epsilon_D \cap |x_L - x'_L| \le \epsilon_L\},\tag{71}
$$

where  $\mathbf{x} = (x_D, x_L)$  and  $\mathbf{x}' = (x'_D, x'_L)$  indicate the depth and length, and we denote the corresponding stability parameters by  $(\epsilon_D, \epsilon_L)$ . This corresponds to  $d(\cdot, \cdot)$  being a weighted  $\ell_{\infty}$ -norm.

We evaluate each algorithm on a  $50 \times 50$  grid of points, with the corresponding values coming from the GP posterior that was derived using the original data. We use the Matérn-5/2 ARD kernel, setting its hyperparameters by maximizing the likelihood on a second (smaller) available dataset. The parameters  $\epsilon_D$  and  $\epsilon_L$  are set to 1.0 and 100.0, respectively. The stability requirement changes the global maximum and its location, as can be observed in Figure  $\frac{1}{\sqrt{2}}$ . The number of sampling rounds is  $T = 120$ , and each algorithm is initialized with the same 10 random data points and corresponding observations. The performance is averaged over 100 different runs, where every run corresponds to a different random initialization. In this experiment, STABLE-GP-UCB achieves the smallest  $\epsilon$ -regret in the early rounds, while in the later rounds STABLEOPT is the only method that reports a near-optimal  $\epsilon$ -stable point.