
Supplementary materials for latent Bayesian melding for integrating individual and population models

Mingjun Zhong, Nigel Goddard, Charles Sutton

School of Informatics

University of Edinburgh

United Kingdom

{mzhong, nigel.goddard, csutton}@inf.ed.ac.uk

Theorem 1. *If $E_{p_S(S)} \left[\frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right] < \infty$, then a constant $c_\alpha < \infty$ exists such that $\int \tilde{p}_{S,\xi}(S, \xi) d\xi dS = 1$, for any fixed $\alpha \in [0, 1]$.*

Proof. If $\alpha = 1$, then $c_\alpha = 1$. If $\alpha = 0$, then,

$$\begin{aligned} \int p_S(S) \left(\frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau^*(f(S))} \right) d\xi dS &= \int p_S(S) \frac{p_\tau(f(S))}{p_\tau^*(f(S))} \frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau(f(S))} d\xi dS \\ &= \int p_S(S) \frac{p_\tau(f(S))}{p_\tau^*(f(S))} dS < \infty \end{aligned}$$

Now we look at $\alpha \in (0, 1)$. Firstly, for $x > 0$, if $\alpha \in (0, 1)$, then $g(x) = x^{1-\alpha}$ is a concave function, because $g(x)'' = -\alpha(1-\alpha)x^{-\alpha-1} < 0$. Similarly, x^α is also a concave function. Then we have

$$\begin{aligned} &\int p_S(S) \left(\frac{p_\tau(f(S)|\xi)p(\xi)}{p_\tau^*(f(S))} \right)^{1-\alpha} d\xi dS \\ &= \int p_S(S) \left(\frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} E_{p(\xi|f(S))} \left[\left(\frac{p_\tau(f(S))}{p_\tau(f(S)|\xi)p(\xi)} \right)^\alpha \right] dS \\ &\leq \int p_S(S) \left(\frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} \left[E_{p(\xi|f(S))} \left(\frac{p_\tau(f(S))}{p_\tau(f(S)|\xi)p(\xi)} \right) \right]^\alpha dS \\ &= \int p_S(S) \left(\frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right)^{1-\alpha} dS \\ &\leq \left[E_{p_S(S)} \left(\frac{p_\tau(f(S))}{p_\tau^*(f(S))} \right) \right]^{1-\alpha} \\ &< \infty \end{aligned}$$

where Jensen's inequality has been applied twice. Therefore, $c_\alpha < \infty$ exists, satisfying $\int \tilde{p}_{\xi,S}(\xi, S) d\xi dS = 1$. \square

Theorem 2. If $\lim_{\delta \rightarrow 0} p_\delta(\tau) = p_\tau^*(\tau)$, and $g_\delta(\tau)$ has bounded derivatives in any order, then $\lim_{\delta \rightarrow 0} \int p_\delta(\tau|S)g_\delta(\tau)d\tau = g(f(S))$.

Proof. Since τ is an Uniform distribution on $[f(S) - \delta, f(S) + \delta]$ conditional on S and δ , we could draw N samples for τ such that $\tau_i = f(S) + (2u_i - 1)\delta$ where u_i is a sample drawn from the standard Uniform distribution, where $i = 1, 2, \dots, N$. By using Monte Carlo approximation and Taylor's expansion, we have

$$\begin{aligned}
& \lim_{\delta \rightarrow 0} \int p_\delta(\tau|S)g_\delta(\tau)d\tau \\
&= \lim_{\delta \rightarrow 0} \int p(\tau \in (f(S) - \delta, f(S) + \delta))g_\delta(\tau)d\tau \\
&= \lim_{\delta \rightarrow 0} \lim_{N \rightarrow \infty} \frac{1}{N} \sum_{i=1}^N g_\delta(f(S) + (2u_i - 1)\delta) \\
&= \lim_{\delta \rightarrow 0} \lim_{N \rightarrow \infty} \left\{ g_\delta(f(S)) + g'_\delta(f(S))\delta \frac{1}{N} \sum_{i=1}^N (2u_i - 1) + \frac{1}{2!} g''_\delta(f(S))\delta^2 \frac{1}{N} \sum_{i=1}^N (2u_i - 1)^2 + \dots \right\} \\
&= \lim_{\delta \rightarrow 0} g_\delta(f(S)) \\
&= g(f(S)).
\end{aligned}$$

This holds, since $|2u_i - 1|^k \leq 1$ ($k = 1, 2, 3, \dots$) and $\lim_{N \rightarrow \infty} \sum_{i=1}^N \frac{1}{N} |2u_i - 1|^k \leq 1$, $\sum_{i=1}^N \frac{1}{N} (2u_i - 1)^k$ converges absolutely when $N \rightarrow \infty$. \square