

## A Supplementary material

### A.1 MCMC algorithm for the Thurstonian Model

In our estimation procedure, the goal is to draw samples for the latent variables  $x_{ij}$ ,  $z_j$ ,  $\mu_0$ ,  $\sigma_0$ ,  $\mu_i$ , and  $\sigma_i$  given the observed orderings  $\mathbf{y}_j$ . For this model, we can estimate all latent variables through Gibbs sampling. We first sample a value for each  $x_{ij}$  conditional on all other variables according to:

$$x_{ij} \mid \mu_i, \sigma_i, \mu_0, \sigma_0, z_j, x_{l,j}, x_{u,j} \sim \begin{cases} N_{truncated}(\mu_i, \sigma_i, x_{l,j}, x_{u,j}) & z_j = 1 \\ N_{truncated}(\mu_0, \sigma_0, x_{l,j}, x_{u,j}) & z_j = 0, \end{cases} \quad (\text{A1})$$

where the sampling distribution is the truncated normal with mean and standard deviation dependent on the latent state  $z_j$ ; with  $z_j = 1$ , and  $z_j = 0$ , the sample comes from the Thurstonian model and the guessing model respectively. The lower and upper bounds for truncated normal are determined by  $x_{l,j}$  and  $x_{u,j}$  respectively. The values  $x_{l,j}$  and  $x_{u,j}$  are based on the samples  $\mathbf{x}_j$  that are ordered right before and after the current value  $x_{ij}$  respectively. Specifically, if  $\pi(i)$  denotes the rank given to item  $i$  and  $\pi^{-1}(i)$  denotes the item assigned to rank  $i$ ,  $l = \pi^{-1}(\pi(i) - 1)$ , and  $u = \pi^{-1}(\pi(i) + 1)$ . We also define  $x_{l,j} = -\infty$  when  $\pi(i) = 1$ , and  $x_{u,j} = \infty$ , when  $\pi(i) = N$ . With these bounds, the observed data influences the possible locations for the samples. It is guaranteed that the ordering of samples  $\mathbf{x}_j$  is consistent with the observed ordering  $\mathbf{y}_j$  for individual  $j$ .

To sample  $\mu_i$ , and  $\sigma_i$  given  $\mathbf{x}$ , we have:

$$\sigma_i^2 \mid \mu_i, s_i^2, \mathbf{z} \sim \text{Inv-}\chi^2(M^{(z=1)} - 1, s_i^2) \quad (\text{A2})$$

$$\mu_i \mid \sigma_i, \bar{x}_i, \mathbf{z} \sim N(\bar{x}_i, \sigma_i / \sqrt{M^{(z=1)}}), \quad (\text{A3})$$

where  $s_i^2$  and  $\bar{x}_i$  are the variance and mean of all samples  $\mathbf{x}_i$  (restricted to individuals assigned to the Thurstonian model) for item  $i$  respectively, and  $M^{(z=1)} = \sum_j z_j$ , the number of individuals assigned to the Thurstonian model. Similar update equations were used to update  $\mu_0$  and  $\sigma_0$  based on the samples of the individuals assigned to the guessing route:

$$\sigma_0^2 \mid \mu_0, s_0^2, \mathbf{z} \sim \text{Inv-}\chi^2(M^{(z=0)} - 1, s_0^2) \quad (\text{A4})$$

$$\mu_0 \mid \sigma_0, \bar{x}_0, \mathbf{z} \sim N(\bar{x}_0, \sigma_0 / \sqrt{M^{(z=0)}}). \quad (\text{A5})$$

In order to prevent a drift in the item positions during estimation (as there is no natural zero point), we fixed the minimum of  $\mu_i$  to 0 and the maximum of  $\mu_i$  to 1, and scaled the other variables accordingly.

Finally, to sample the assignment of individuals to modeling routes, we use

$$p(z_j = k \mid \mu_i, \sigma_i, \mu_0, \sigma_0, x_{l,j}, x_{u,j}) \propto \begin{cases} \prod_{i=1}^N f(x_{ij} \mid \mu_i, \sigma_i) & k = 1 \\ \prod_{i=1}^N f(x_{ij} \mid \mu_0, \sigma_0) & k = 0. \end{cases} \quad (\text{A6})$$

where  $f(x \mid \mu, \sigma)$  is the normal probability density function. In our procedure, we ran 20 chains with a burn-in of 200 iterations. From each chain, we drew 20 samples with an interval of 10 iterations. In total, we collected 400 samples. To construct a single group answer, we analyzed the ordering of the items according to  $\mu_i$ , separately for each sample, and then picked the mode of this distribution. This corresponds to the most likely order in the distribution over orders inferred by the model.

## A.2 MCMC algorithm for Mallows Model

In our MCMC algorithm for Mallows model, we use a combination of Metropolis-Hastings (MH) and Gibbs sampling steps. To estimate  $\omega$ , we use the MH algorithm based on Lebanon and Lafferty (2002). The idea is to move the group estimate  $\omega$  by transposing any randomly chosen pair of items. The proposal distribution  $q(\omega^*|\omega)$  is

$$q(\omega^*|\omega) = \begin{cases} 1/\binom{n}{2} & \text{if } S(\omega^*, \omega) = 1 \\ 0 & \text{otherwise,} \end{cases} \quad (\text{A7})$$

where  $S(\omega', \omega)$  is the Cayley distance. The Metropolis-Hastings acceptance ratio is

$$\min \left[ 1, \frac{q(\omega|\omega') p(y|\omega', \theta, z)}{q(\omega'|\omega) p(y|\omega, \theta, z)} \right]. \quad (\text{A8})$$

Note that the first likelihood ratio for the proposal distribution equals one because of the symmetry in the proposals. Also, in Eq 1., the normalization constant does not depend on  $\omega$ , which can be used to simplify the acceptance ratio to:

$$\min \left[ 1, \exp \left( -\theta \sum_{z_j=1} d(\mathbf{y}_j, \omega') - d(\mathbf{y}_j, \omega) \right) \right], \quad (\text{A9})$$

where the sum is taken over all individuals currently assigned to Mallows model. To facilitate the inference for  $\theta$ , we used a discretized set of 1000  $\theta$  values, logarithmically spaced between  $10^{-4}$  and 2. Let  $v_k$  refer to the  $k$ th value in this set. We use a Gibbs sampling step for  $\theta$  by sampling from the discrete distribution

$$p(\theta = v_k | \omega, z, \mathbf{y}) \propto \exp \left[ -v_k \sum_{z_j=1} d(\mathbf{y}_j, \omega) - \sum_{z_j=1} \log \Psi(v_k) \right]. \quad (\text{A10})$$

Finally, we use a Gibbs sampling step to estimate the latent state  $z_j$

$$p(z_j = k | \theta, \omega, z_{-j}, \mathbf{y}_j) = \begin{cases} 1/N! & k = 0 \\ \exp[-\theta d(\mathbf{y}_j, \omega) - \log \Psi(\theta)] & k = 1. \end{cases} \quad (\text{A11})$$

In the MCMC procedure, we ran 20 chains with a burn-in of 200 iterations. From each chain, we drew 20 samples with an interval of 10 iterations. In total, we collected 400 samples. To construct a single group answer, we picked the most frequently occurring sampled ordering  $\omega$ .