

---

# Algorithms for Infinitely Many-Armed Bandit (Supplementary file)

---

**Yizao Wang\***

Department of Statistics - University of Michigan  
437 West Hall, 1085 South University, Ann Arbor, MI, 48109-1107, USA  
yizwang@umich.edu

**Jean-Yves Audibert**

Université Paris Est, Ecole des Ponts, ParisTech, Certis  
& Willow - ENS / INRIA, Paris, France  
audibert@certis.enpc.fr

**Rémi Munos**

INRIA Futurs Lille, SequeL project,  
40 avenue Halley, 59650 Villeneuve d'Ascq, France  
remi.munos@inria.fr

**Theorem 3** Any algorithm suffers a regret larger than  $cn^{\frac{\beta}{1+\beta}}$  for some small enough constant  $c$  depending on  $c_2$  and  $\beta$ .

**Proof of Theorem 3.** An elementary event of the probability space is characterized by the infinite sequence  $I_1, I_2, \dots$  of arms and by the infinite sequences of rewards corresponding to each of the arm:  $X_{I_1,1}, X_{I_1,2}, \dots, X_{I_2,1}, X_{I_2,2}, \dots$ , and so on. Arm  $I_1$  is the first arm drawn,  $I_2 \neq I_1$  is the second one, and so on. Let  $0 < \delta < \delta' < \mu^*$ . Let  $K^*$  denote the smallest  $\ell$  such that  $\mu_{I_\ell} > \mu^* - \delta$ . Let  $\bar{K}$  be the number of arms in  $\{I_1, \dots, I_{K^*-1}\}$  with expected reward smaller than or equal to  $\mu^* - \delta'$ . An algorithm will request a number of arms  $K$ , which is a random variable (possibly depending on the obtained rewards). Let  $\hat{\mu}$  be the expected reward of the best arm in  $\{I_1, \dots, I_K\}$ . Let  $\kappa > 0$  a parameter to be chosen. We have

$$\begin{aligned} R_n &= R_n \mathbf{1}_{\hat{\mu} \leq \mu^* - \delta} + R_n \mathbf{1}_{\hat{\mu} > \mu^* - \delta} \\ &\geq n\delta \mathbf{1}_{\hat{\mu} \leq \mu^* - \delta} + \bar{K} \delta' \mathbf{1}_{\hat{\mu} > \mu^* - \delta} \\ &\geq n\delta \mathbf{1}_{\hat{\mu} \leq \mu^* - \delta} + \kappa \delta' \mathbf{1}_{\hat{\mu} > \mu^* - \delta; \bar{K} \geq \kappa}, \end{aligned}$$

where the first inequality uses that  $\hat{\mu} > \mu^* - \delta$  implies that the arms  $I_1, \dots, I_{K^*}$  have been at least tried once. By taking expectations on both sides and taking  $\kappa = n\delta/\delta'$ , we get

$$\mathbb{E}R_n \geq n\delta \mathbb{P}(\hat{\mu} \leq \mu^* - \delta) + \kappa \delta' (\mathbb{P}(\hat{\mu} > \mu^* - \delta) - \mathbb{P}(\bar{K} < \kappa)) = \delta' \kappa \mathbb{P}(\bar{K} \geq \kappa).$$

Now the random variable  $\bar{K}$  follows a geometric distribution with parameter  $p = \frac{\mathbb{P}(\mu > \mu^* - \delta)}{\mathbb{P}(\mu \notin (\mu^* - \delta', \mu^* - \delta])}$ . So we have  $\mathbb{E}R_n \geq \delta' \kappa (1-p)^\kappa$ . Taking  $\delta = \delta' n^{-1/(\beta+1)}$  and  $\delta'$  a constant value in  $(0, \mu^*)$  (for instance  $(2c_2)^{-1/\beta}$  to ensure  $p \leq 2c_2 \delta^\beta$ ), we have  $\kappa = n^{\frac{\beta}{1+\beta}}$  and  $p$  is of order  $1/\kappa$  and obtain the desired result.

**Theorem 4** For any horizon time  $n \geq 2$ , the expected regret of the UCB-AIR algorithm satisfies

$$\mathbb{E}R_n \leq \begin{cases} C(\log n)^2 \sqrt{n} & \text{if } \beta < 1 \text{ and } \mu^* < 1 \\ C(\log n)^2 n^{\frac{\beta}{1+\beta}} & \text{otherwise, i.e. if } \mu^* = 1 \text{ or } \beta \geq 1 \end{cases} \quad (1)$$

with  $C$  a constant depending only on  $c_1, c_2$  and  $\beta$ .

---

\*The major part of this work was completed during the research internship at Certis and INRIA SequeL.

**Proof of Theorem 4.** We essentially need to adapt the proof of Theorem 1. We recall that  $K_n$  denote the number of arms played up to time  $n$ . Let  $I_1, \dots, I_{K_n}$  denote the selected arms:  $I_1$  is the first arm drawn,  $I_2$  the second, and so on. Let  $S_k$  denote the time arm  $k$  being played for the first time.  $1 = S_{I_1} < S_{I_2} < \dots < S_{I_{K_n}}$ . Since arms  $I_1, \dots, I_{K_n}$  progressively enter in competition, Lemma 1 no longer holds but an easy adaptation of its proof shows that for  $k \in \{I_1, \dots, I_{K_n}\}$ ,

$$\mathbb{E}(T_k(n)|I_1, \dots, I_{K_n}) \leq u + \sum_{t=u+1}^n \sum_{s=u}^t \mathbb{P}(B_{k,s,t} > \tau) + \Omega_k \quad (2)$$

with

$$\Omega_k = \sum_{t=u+1}^n \prod_{k' \neq k, S_{k'} \leq t} \mathbb{P}(\exists s' \in [0, t], B_{k',s',t} \leq \tau).$$

As in the proof of Theorem 1, since the exploration sequence satisfies  $\mathcal{E}_t \geq 2 \log(10 \log t)$ , we have  $\mathbb{P}(\exists s' \in [0, t], B_{k',s',t} \leq \tau) \leq 1/2$  for arms  $k'$  such that  $\mu_{k'} \geq \tau$ . Consequently, letting  $N_{\tau,k,t}$  denote the cardinal of the set  $\{k' : k' \neq k, \mu_{k'} \geq \tau, S_{k'} \leq t\}$ , we have

$$\Omega_k \leq \sum_{t=1}^n 2^{-N_{\tau,k,t}}.$$

Let us first consider the case  $\mu^* = 1$  or  $\beta \geq 1$ . In the case of UCB-AIR,  $S_{I_j}$  is the smallest integer strictly larger than  $(j-1)^{(\beta+1)/\beta}$ . To shorten notation, let us write  $S_j$  for  $S_{I_j}$ . According to the arm-increasing rule (try a new arm if  $K_{t-1} < t^{\beta/(\beta+1)}$ ),  $[S_j, S_{j+1})$  is the time interval in which the competing arms are  $I_1, I_2, \dots, I_j$ .

As in the proof of Theorem 1, we consider  $\tau = \mu^* - \Delta_k/2$ . We have

$$\begin{aligned} \mathbb{E}(\Omega_{I_\ell} | I_\ell = k) &\leq \sum_{j=1}^{K_n} \sum_{t=S_j}^{S_{j+1}-1} \mathbb{E}\left(2^{-N_{\tau,k,S_j}} | I_\ell = k\right) \\ &= \sum_{j=1}^{K_n} (S_{j+1} - S_j) \mathbb{E}\left(2^{-N_{\tau,k,S_j}} | I_\ell = k\right) \\ &\leq \sum_{j=1}^{K_n} (S_{j+1} - S_j) \mathbb{E}\left(2^{-N_{\tau,\infty,S_{j-1}}}\right). \end{aligned} \quad (3)$$

Since  $N_{\tau,\infty,S_{j-1}}$  follows a binomial distribution with parameter  $j-1$  and  $\mathbb{P}(\mu \geq \tau)$ , we have

$$\mathbb{E}\left(2^{-N_{\tau,\infty,S_{j-1}}}\right) = (1 - \mathbb{P}(\mu \geq \tau)/2)^{j-1},$$

and

$$\begin{aligned} \sum_{j=1}^{K_n} (S_{j+1} - S_j) \mathbb{E}\left(2^{-N_{\tau,\infty,S_{j-1}}}\right) &= \sum_{j=1}^{K_n} (S_{j+1} - S_j) (1 - \mathbb{P}(\mu \geq \tau)/2)^{j-1} \\ &\leq \sum_{j=1}^{K_n} \left(1 + \frac{\beta+1}{\beta} j^{\frac{1}{\beta}}\right) (1 - \tilde{c}[2(\mu^* - \tau)]^\beta)^{j-1}, \end{aligned} \quad (4)$$

where  $\tilde{c} = c_1 2^{-1-\beta}$ . Plugging (4) into (3), we obtain

$$\mathbb{E}(\Delta_{I_\ell} \Omega_{I_\ell}) \leq \frac{2\beta+1}{\beta} \sum_{j=1}^{K_n} j^{\frac{1}{\beta}} \mathbb{E}(\Delta_{I_\ell} [1 - \tilde{c}\Delta_{I_\ell}^\beta]^{j-1}).$$

Now this last expectation can be bounded by the same computations as for  $\mathbb{E}\chi(\Delta_1)$  in the proof of Theorem 1. We have, for appropriate positive constants  $C_1$  and  $C_2$  depending on  $c_1$  and  $\beta$ ,

$$\mathbb{E}(\Delta_{I_\ell} \Omega_{I_\ell}) \leq C_1 \sum_{j=1}^{K_n} j^{\frac{1}{\beta}} j^{-\frac{1}{\beta}} \frac{\log j}{j} \leq C_2 (\log K_n)^2. \quad (5)$$

Using (2) and  $\mathbb{E}R_n = \sum_{\ell=1}^{K_n} \mathbb{E}(\Delta_{I_\ell} \Omega_{I_\ell})$ , we obtain

$$\mathbb{E}R_n \leq K_n \mathbb{E}\left\{ \left[ 50 \left( \frac{V(\Delta_1)}{\Delta_1} + 1 \right) \log n \right] \wedge (n\Delta_1) + C_2 (\log K_n)^2 \right\}, \quad (6)$$

from which Theorem 4 follows for the case  $\mu^* = 1$  or  $\beta \geq 1$ . For the case  $\beta < 1$  and  $\mu^* < 1$ , replacing  $\frac{\beta}{\beta+1}$  by  $\frac{\beta}{2}$  leads to a similar version of (5) as

$$\mathbb{E}(\Delta_{I_\ell} \Omega_{I_\ell}) \leq C_1 \sum_{j=1}^{K_n} j^{\frac{2}{\beta}-1} j^{-\frac{1}{\beta}} \frac{\log j}{j} \leq C_2 (\log K_n) K_n^{\frac{1-\beta}{\beta}},$$

which gives the desired convergence rate since  $K_n$  is of order  $n^{\beta/2}$ .